

# *INLA: an introduction*

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<sup>1</sup>Joint work with S.Martino (Trondheim) and N.Chopin (Paris)

## Latent Gaussian models

*Stage 1* Observed data  $\mathbf{y} = (y_i)$ ,

$$y_i \mid \mathbf{x}, \boldsymbol{\theta} \sim \pi(y_i \mid x_i, \boldsymbol{\theta})$$

*Stage 2* Latent Gaussian field

$$\mathbf{x} \mid \boldsymbol{\theta} \sim \mathcal{N}(\boldsymbol{\mu}, \mathbf{Q}(\boldsymbol{\theta})^{-1}), \quad \mathbf{A}\mathbf{x} = \mathbf{0}$$

*Stage 3* Priors for the hyperparameters

$$\boldsymbol{\theta} \sim \pi(\boldsymbol{\theta})$$

Unify many of the most used models in statistics

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## Structured additive regression models

### Linear predictor

$$\eta_i = \sum_k \beta_k z_{ki} + \sum_j w_{ji} f_j(z_{ji}) + \epsilon_i$$

- Linear effects of covariates  $\{z_{ki}\}$
- Effects of  $f_j(\cdot)$ 
  - Fixed weights  $\{w_{ji}\}$
  - Commonly:  $f_j(z_{ji}) = f_{j,z_{ji}}$
  - Account for smooth response
    - Temporal or spatially indexed covariates
    - Unstructured terms ("random effects")
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from an (f.ex) exponential family with mean  $\mu_i = g^{-1}(\eta_i)$ .

- *Latent Gaussian model* if

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## *Examples*

- **Dynamic linear models**
- Stochastic volatility
- Generalised linear (mixed) models
- Generalised additive (mixed) models
- Spline smoothing
- Semiparametric regression
- Space-varying (semiparametric) regression models
- Disease mapping
- Log-Gaussian Cox-processes
- Model-based geostatistics (\*)
- Spatio-temporal models

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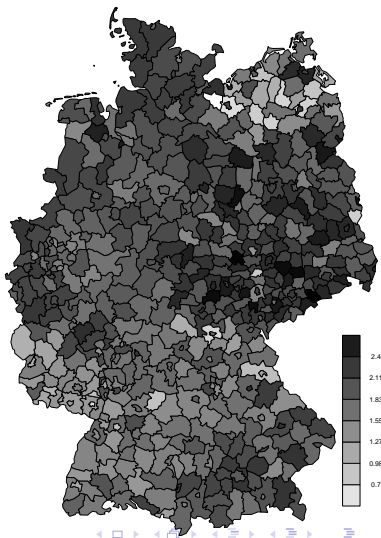
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## Example: Disease mapping (BYM-model)

- Data  $y_i \sim \text{Poisson}(E_i \exp(\eta_i))$
- Log-relative risk  
 $\eta_i = u_i + v_i + \beta^T \mathbf{z}_i$
- Structured component  $\mathbf{u}$
- Unstructured component  $\mathbf{v}$
- Covariates  $\mathbf{z}_i$
- Log-precisions  $\log \kappa_u$  and  $\log \kappa_v$



## *Characteristic features*

- Large dimension of the latent Gaussian field:  $10^2 - 10^5$
- A lot of conditional independence in the latent Gaussian field
- Few hyperparameters  $\theta$ :  $\dim(\theta)$  between 1 and 5
- Non-Gaussian data

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## Main task

- Compute the posterior marginals for the latent field

$$\pi(x_i | \mathbf{y}), \quad i = 1, \dots, n$$

- Compute the posterior marginals for the hyperparameters

$$\pi(\theta_j | \mathbf{y}), \quad j = 1, \dots, \dim(\boldsymbol{\theta})$$

- Today's "standard" approach, is to make use of MCMC
- Main difficulties
  - CPU-time
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## *Integrated nested Laplace approximations (INLA)*

- Utilise of the latent Gaussian field
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- Utilise small  $\dim(\theta)$ 
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- HUGE improvement in both speed and accuracy compared to MCMC alternatives
- Relative error
- Practically “exact” results<sup>2</sup>
- Extensions: Marginal likelihood, DIC, Cross-validation, ...

INLA enable us to treat Bayesian latent Gaussian models properly and bring these models from the research communities to the end-users

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## Main ideas (I)

$$\pi(z) = \frac{\pi(x, z)}{\pi(x|z)} \quad \text{leading to} \quad \tilde{\pi}(z) = \frac{\pi(x, z)}{\tilde{\pi}(x|z)} \Big|_{\text{mode}(z)}$$

- When  $\tilde{\pi}(x|z)$  is the Gaussian-approximation, this is the Laplace-approximation
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Posterior

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## Remarks

1. Expect  $\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$  to be accurate, since
  - $\mathbf{x}|\boldsymbol{\theta}$  is *a priori* Gaussian
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Make use of the conditional independence properties in the latent field

$$x_i \perp x_j \mid \mathbf{x}_{-ij} \iff Q_{ij} = 0$$

where  $\mathbf{Q}$  is the precision matrix (inverse covariance)

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## *Simplified Laplace Approximation*

Expand the Laplace approximation of  $\pi(x_i|\boldsymbol{\theta}, \mathbf{y})$ :

$$\log \tilde{\pi}(x_i|\boldsymbol{\theta}, \mathbf{y}) = -\frac{1}{2}x_i^2 + b_i x_i + \frac{1}{6}d_i x_i^3 + \dots$$

### Remarks

- Correct the Gaussian approximation for error in shift and skewness through  $b_i$  and  $d_i$
- Fit a skew-Normal density

$$2\phi(x)\Phi(ax)$$

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Expand the Laplace approximation of  $\pi(x_i|\boldsymbol{\theta}, \mathbf{y})$ :

$$\log \tilde{\pi}(x_i|\boldsymbol{\theta}, \mathbf{y}) = -\frac{1}{2}x_i^2 + b_i x_i + \frac{1}{6}d_i x_i^3 + \dots$$

Remarks

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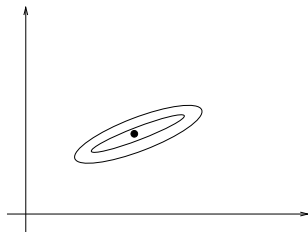
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Step I Explore  $\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$

- Locate the mode
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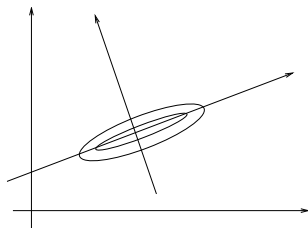




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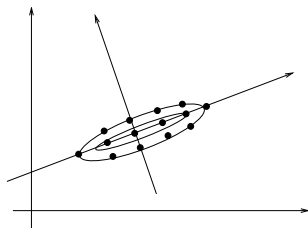
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## *The integrated nested Laplace approximation (INLA) II*

*Step II* For each  $\theta_j$

- For each  $i$ , compute the (simplified) Laplace approximation for  $x_i$

## The integrated nested Laplace approximation (INLA) III

### Step III Sum out $\theta_j$

- For each  $i$ , sum out  $\theta$

$$\tilde{\pi}(x_i | \mathbf{y}) \propto \sum_j \tilde{\pi}(x_i | \mathbf{y}, \theta_j) \times \tilde{\pi}(\theta_j | \mathbf{y})$$

- Build a log-spline corrected Gaussian

$$\mathcal{N}(x_i; \mu_i, \sigma_i^2) \times \exp(\text{spline})$$

to represent  $\tilde{\pi}(x_i | \mathbf{y})$ .

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## *How can we assess the errors in the approximations?*

Important, but asymptotic arguments are difficult:

$$\dim(\mathbf{y}) = \mathcal{O}(n) \quad \text{and} \quad \dim(\mathbf{x}) = \mathcal{O}(n)$$

## *Errors in the approximations of $\pi(x_i|\mathbf{y})$*

Compare a sequence of improved approximations

1. Gaussian approximation
2. Simplified Laplace
3. Laplace

Compute the full Laplace-approximation for  $\pi(x_i|\mathbf{y}, \theta_j)$  only if the Gaussian and the Simplified Laplace approximation disagree.

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## Overall check: Equivalent number of replicates

**Tool 3:** Estimate the “effective” number of parameters

- From the *Deviance Information Criteria*:

$$p_D(\theta) \approx n - \text{trace}(\mathbf{Q}_{\text{prior}}(\theta) \mathbf{Q}_{\text{post.}}(\theta)^{-1})$$

- Compare with the number of observations:

$$\# \text{observations} / p_D(\theta)$$

high ratio is good

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## Examples

- 10:30-11:30: Andrea Riebler: Performance of INLA analysing bivariate meta-regression and age-period-cohort models.
- 12:30-13:30: Birgit Schrödle: Spatio-temporal disease mapping using INLA.
- 13:45-14:45: Virgilio Gómez-Rubio: Approximate Bayesian Inference for Small Area Estimation
- 15:00-16:00: Lea Fortunato: About the *Rapid Inquiry Facility*, and spatial analyses with WinBUGS and INLA.
- 09:00-10:00: Rupali Akerkar: Approximate Bayesian Inference for Survival models.
- 10:15-11:30: Ingelin Steinsland & Anna Marie Holand: Animal Model and INLA.

## *Marginal likelihood*

Marginal likelihood is the normalising constant for  $\pi(\boldsymbol{\theta}|\mathbf{y})$

## *Deviance Information Criteria*

$$D(\mathbf{x}; \boldsymbol{\theta}) = -2 \sum_i \log(y_i | x_i, \boldsymbol{\theta})$$

$$DIC = 2 \times \text{Mean}(D(\mathbf{x}; \boldsymbol{\theta})) - D(\text{Mean}(\mathbf{x}); \boldsymbol{\theta}^*)$$

## Cross-validation

- Based on

$$\pi(x_i | \mathbf{y}_{-i}, \boldsymbol{\theta}) \propto \frac{\pi(x_i | \mathbf{y}, \boldsymbol{\theta})}{\pi(y_i | x_i, \boldsymbol{\theta})}$$

we can compute

$$\pi(y_i | \mathbf{y}_{-i})$$

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## *“Recent” developments in (R-)INLA*

- Improving the code: speedups and improved parallel performance
- Improving the R-interface
- Extending the building-blocks: prior-models and likelihood-models

Lot of things still todo...

- Documentation
- Worked out examples
- Webpage
- +++

## *Improving the code*

The two largest changes the last year

- (summer 2008) Change the way `inla` works in a multi-core environment.  
Especially important for more than “dual-core”
- (xmas 2008) Implement a more efficient rank-one update formula (Thanks Birgit!).  
Especially important for models with many constraints, but gave a huge speedup also for models with a single constraint.

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## *inla* & multi-core

- Gradient and Hessian computations are done in parallel

$$\frac{\partial}{\partial \theta_i} \tilde{\pi}(\boldsymbol{\theta} \mid \mathbf{y}) \quad \text{and} \quad \frac{\partial^2}{\partial \theta_i \partial \theta_j} \tilde{\pi}(\boldsymbol{\theta} \mid \mathbf{y})$$

- All computations of

$$\tilde{\pi}(x_i \mid \boldsymbol{\theta}_j, \mathbf{y})$$

is now done in parallel wrt  $j$ , and not  $i$  as before.

*Multi-core example*

inla-example 7: CANCER-INCIDENCE,  $n = 7138$ ,  $\dim(\theta) = 3$

Number of cores	Time used
1	20.0s
2	13.3s
4	9.1s
6	8.6s
8	8.3s

- Linear algebra (Cholesky/Solve/Inverse):  $\approx 50\%$
- Administration:  $\approx 50\%$

## Recent developments in (R-)INLA

### Improved R-interface

- `inla`-binary is now bundled with the R-package.
- Make use of `model.matrix()` in R:

*formula* ~ ... + *x\*z*

will expand as

*formula* ~ ... + *x* + *z* + *x:z*

- Allow for

*formula* ~ .... + *offset(a)* + *offset(b)*

defining *a+b* to be fixed offset in formula

Larger change for survival-models (more on this tomorrow...)



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## *Building blocks*

```
> names(inla.models())$models
 [1] "iid"           "rw1"           "rw2"
 [4] "crw2"         "seasonal"     "besag"
 [7] "ar1"          "generic"      "2diid"
[10] "2diidwishart" "2diidwishartpart0" "2diidwishartpart1"
[13] "3diidwishartpart0" "3diidwishartpart1" "3diidwishartpart2"
[16] "z"            "rw2d"         "matern2d"
```

## The “z”-model

The z-models is

$$\boldsymbol{\eta} = \dots + \mathbf{Z}\mathbf{z} + \dots$$

where  $\mathbf{Z}$  is a  $n \times k$  matrix and  $\mathbf{z} \sim \mathcal{N}_k(0, \tau\mathbf{I})$ .

## *Likelihood models*

```
> names(inla.models())$lmodels
 [1] "poisson"                "binomial"
 [3] "exponential"           "piecewise.constant"
 [5] "piecewise.linear"      "gaussian"
 [7] "laplace"               "weibull"
 [9] "weibullcure"           "stochvol_t"
[11] "zeroinflated_poisson_0" "zeroinflated_poisson_1"
[13] "zeroinflated_binomial_0" "zeroinflated_binomial_1"
[15] "T"                      "stochvol.nig"
```

## *Zero-inflated models*

The (type 0) likelihood is defined as

$$\text{Prob}(y \mid \dots) = p \times 1_{[y=0]} + (1 - p) \times \text{Poisson}(y \mid y > 0)$$

The (type 1) likelihood is defined as

$$\text{Prob}(y \mid \dots) = p \times 1_{[y=0]} + (1 - p) \times \text{Poisson}(y)$$

[Birgit: Mixture over  $p$  is still on the list... sorry.]

## *The TODO-list*

- Better support for spatial (GMRF) models. More on this tomorrow (Finn.L).
- Alternative spline-models, B-splines? (Thomas.K ?)
- Simultaneous credibility intervals
- Documentation
- Worked out examples
- Webpage
- +++



## Summary (I)

- Latent Gaussian models unifies many models into the same framework: unified approach towards Bayesian inference
  - Use the **same** computer code
  - Near optimal numerical algorithms for the sparse matrices
  - Achieve nice speed-ups in a multi-core environment
  - Practically “exact” results
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