



- 1 (Øksendal 2:2.15) Prove that Brownian motion has the following fractal behavior: If B_t is a standard one-dimensional Brownian motion and $c > 0$, then the process

$$(1) \quad Y_t = \frac{1}{c} B_{c^2 t}$$

is also a standard Brownian motion.

- 2 (Øksendal 2:2.12 and 2:2.15) Let B_t be Brownian motion on \mathbb{R}^d , $t_0 > 0$, and $U \in \mathbb{R}^{d \times d}$ such that $U^T U = I$.

Show that the following processes also are Brownian motions on \mathbb{R}^d :

a) $\tilde{B}_t = B_{t+t_0} - B_{t_0}$.

b) $\tilde{B}_t = U B_t$.

- 3 (Øksendal 2:2.13) Let B_t be a Brownian motion in \mathbb{R}^2 and $\rho > 0$. Compute

$$P(|B_t| < \rho).$$

Hint: B_t is $N(\vec{0}, tI)$, use the probability density function.

- 4 (Øksendal 2:2.8) Let B_t be a Brownian motion in \mathbb{R} .

a) Show that $E(e^{iuB_t}) = e^{-\frac{1}{2}u^2 t}$ and use this to prove that $E(B_t^4) = 3t^2$.

b) Use $E(f(B_t)) = \frac{1}{\sqrt{2\pi t}} \int_{\mathbb{R}} f(x) e^{-\frac{x^2}{2t}} dx$ and integration by parts to show that $E(B_t^4) = 3t^2$.

- 5 Consider the standard Brownian motion in \mathbb{R}^n , B_t . Prove that

$$(2) \quad E(|B_t - B_s|^4) = n(n+2)(t-s)^2.$$

Hint: For this problem you could use that

$$(3) \quad E(X_1^2 + \dots + X_n^2)^2 = E\left(\sum_{i,j=1}^n X_i^2 X_j^2\right),$$

and use the independence and distribution of the increments and the fact that $E(B_i^4(t)) = 3t$ for $i = 1, \dots, n$ (previous exercise).

6 Let B_t be a Brownian motion on \mathbb{R} , and let $0 < t_1 < t_2 < \dots < t_n$.

a) Show that the characteristic function of B_{t_1} and B_{t_2}

$$\Phi_{B_{t_1}, B_{t_1}}(u_1, u_2) = e^{\frac{1}{2}u^T \Sigma u},$$

$$\text{for } u = (u_1, u_2)^T \text{ and } \Sigma = \begin{pmatrix} t_1 & t_1 \\ t_1 & t_2 \end{pmatrix}.$$

Hint: Write $B_{t_2} = B_{t_1} + (B_{t_2} - B_{t_1})$ and use independence of the R.V.'s on the right. What can you say about characteristic functions of independent R.V.'s?

b) Show that B_t is a Gaussian process.

Hint: Show that B_{t_1}, \dots, B_{t_n} is multivariate Gaussian by considering the characteristic function. Obs: You need not compute the covariance matrix explicitly to conclude.

7 Let B_t be a Brownian motion and \mathcal{F}_t the filtration generated by B_t . Show that $B_t^2 - t$ is a martingale w.r.t. \mathcal{F}_t .

Hint: $B_t^p = B_s^p + (B_t^p - B_s^p)$ for $p = 1, 2$, use different properties of the conditional expectation.

8 Let B_t be a Brownian motion in \mathbb{R}^1 and \mathcal{F}_t the filtration generated by B_t .

a) Show that $E((e^{B_t})^p) = e^{\frac{1}{2}p^2 t}$ for $p \in \mathbb{N}$. (e^{B_t} is geometric B.M.)

Hint: Integrate using the probability density function, complete the squares.

b) Show that $e^{B_t - \frac{1}{2}t}$ is a martingale w.r.t. \mathcal{F}_t .

Hint: $B_t = B_s + (B_t - B_s)$, use a) and different properties of the conditional expectation.

9 Let B_t be a Brownian motion on \mathbb{R} with filtration $\{\mathcal{F}_t\}_t$,

$$0 = t_0 < t_1 < t_2 < \dots,$$

$\Delta B_{t_i} = B_{t_i} - B_{t_{i-1}}$ and $h(\omega, t)$ an $\{\mathcal{F}_t\}_t$ -adapted stochastic process such that $E(h^2(\cdot, t)) < \infty$ for each t . Verify that the following process is a $\{\mathcal{F}_{t_i}\}_i$ -martingale (i.e. a discrete martingale):

$$Y_0 = 0, \quad Y_n = \sum_{i=1}^n h(\omega, t_{i-1}) \Delta B_{t_i} \quad \text{for } n \in \mathbb{N}.$$

10 Let B_t be a Brownian motion in \mathbb{R} , $0 = t_0 < t_1 < \dots < t_n = 1$, and \tilde{B}_t the linear interpolant of $\{(t_i, B_{t_i})\}_i$: For every $k = 0, \dots, n$,

$$\tilde{B}_t = B_{t_k} + \frac{t - t_k}{t_{k+1} - t_k} (B_{t_{k+1}} - B_{t_k}) \quad \text{for } t \in [t_k, t_{k+1}].$$

Show that:

a) $\tilde{B}_0 = 0$ a.s. and $E(\tilde{B}_t) = 0$.

b) $|E(\tilde{B}_t \tilde{B}_s) - \min(s, t)| \leq \max_k |t_{k+1} - t_k|$.

c) \tilde{B}_t is a Gaussian process.

Hint c): Use characteristic functions. Similar computation as in Exercise 6.

11 Let X_1, X_2, \dots be i.i.d. $N(0, 1)$ random variables. Define

$$\begin{aligned}\Delta B_1^0 &= X_1, \\ \Delta B_{2^{m-1}}^n &= \Delta B_m^{n-1} + \frac{1}{2^{\frac{n+1}{n}}} X_{i(n,m)}, \\ \Delta B_{2^m}^n &= \Delta B_m^{n-1} - \frac{1}{2^{\frac{n+1}{n}}} X_{i(n,m)},\end{aligned}$$

for $n = 1, 2, 3, \dots$, $m = 1, \dots, 2^{n-1}$ and $i(n, m) = 2^{n-1} + m$.

Prove by induction that for each fixed n ,

$$\Delta B_1^n, \dots, \Delta B_{2^n}^n$$

are independent $N(0, \frac{1}{2^n})$ random variables.

Hint: You may assume that $\Delta B_1^n, \dots, \Delta B_{2^n}^n$ are multivariate Gaussian, so zero correlation shows independence.

[If $X \in \mathbb{R}^d$ is multivariate Gaussian, then AX is also so for any $A \in \mathbb{R}^{d \times d}$]