

# TMA4275 LIFETIME ANALYSIS

Slides 4: Gumbel distribution. Log-location-scale families

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# EXTREME VALUE DISTRIBUTIONS

Let  $T_1, T_2, \dots, T_n$  be lifetimes of  $n$  components, with ordered values denoted by  $T_{(1)} < T_{(2)} < \dots < T_{(n)}$ . Thus  $T_{(1)}$  is the minimum and corresponds to the lifetime of a series system.

For large  $n$ ,  $T_{(1)}$  is approximately Weibull-distributed. *This motivates the widespread use of the Weibull-distribution!*

If the  $T_i$  are no longer lifetimes, but have support in  $(-\infty, \infty)$ , then the limiting distribution of a properly normalized version of  $T_{(1)}$  equals the distribution of a random variable  $Y$  with cdf

$$F_Y(y) = 1 - e^{-e^{\frac{y-\mu}{\sigma}}}, \quad -\infty < y < \infty$$

This is the so called “Distribution of smallest extreme”, or “Extreme value distribution of type I”, or (which we will call it) the **Gumbel-distribution**;  $Y \sim \text{Gumbel}(\mu, \sigma)$ .

We write  $Y \sim \text{Gumbel}(\mu, \sigma)$

Show that

$$F_Y(y) = 1 - e^{-e^{\frac{y-\mu}{\sigma}}}, \quad -\infty < y < \infty$$

satisfies the requirements for a cdf, i.e.

- Increasing in  $y$
- $\lim_{y \rightarrow -\infty} F_Y(y) = 0$
- $\lim_{y \rightarrow \infty} F_Y(y) = 1$

# WHY ARE WE INTERESTED IN THE GUMBEL DISTRIBUTION?

If  $T$  is Weibull-distributed,  $T \sim \text{Weib}(\alpha, \theta)$ , then  $Y = \ln T$  is Gumbel-distributed,  $Y \sim \text{Gumbel}(\mu, \sigma)$ , with  $\mu = \ln \theta$ ,  $\sigma = 1/\alpha$ .

*Proof:* Note first that  $T = e^Y$  and  $R(t) = P(T > t) = e^{-\left(\frac{t}{\theta}\right)^\alpha}$ . Then:

$$\begin{aligned}P(Y > y) &= P(e^Y > e^y) = P(T > e^y) = R(e^y) \\&= e^{-\left(\frac{e^y}{\theta}\right)^\alpha} = e^{-\left(\frac{e^y}{e^{\ln \theta}}\right)^\alpha} \\&= e^{-(e^{y-\ln \theta})^\alpha} = e^{-e^{\left(\frac{y-\ln \theta}{1/\alpha}\right)}}\end{aligned}$$

Thus,  $F_Y(y) = 1 - P(Y > y) = 1 - e^{-e^{\left(\frac{y-\ln \theta}{1/\alpha}\right)}}$ , which shows that  $Y \sim \text{Gumbel}(\ln \theta, 1/\alpha)$ .

*We shall see later why this is a useful and interesting result (and not just a curiosity...)*

Let  $Y \sim \text{Gumbel}(\mu, \sigma)$  and recall the cdf

$$F_Y(y) = P(Y \leq y) = 1 - e^{-e^{\frac{y-\mu}{\sigma}}} \quad \text{for } -\infty < y < \infty$$

The cdf of  $\text{Gumbel}(0,1)$ , called *the standard Gumbel distribution*, is

$$G(w) = 1 - e^{-e^w} \quad \text{for } -\infty < w < \infty$$

Suppose  $W \sim \text{Gumbel}(0,1)$  and let

$$Y = \mu + \sigma W$$

Then

$$F_Y(y) = P(Y \leq y) = P(\mu + \sigma W \leq y) = P(W \leq \frac{y - \mu}{\sigma}) = G\left(\frac{y - \mu}{\sigma}\right)$$

so  $Y \sim \text{Gumbel}(\mu, \sigma)$ .

Thus we can represent  $Y \sim \text{Gumbel}(\mu, \sigma)$  as

$$Y = \mu + \sigma W$$

where  $W \sim \text{Gumbel}(0, 1)$ . Further,

$$F_Y(y) = P(Y \leq y) = G\left(\frac{y - \mu}{\sigma}\right)$$

where  $G(\cdot)$  is the cdf of  $\text{Gumbel}(0, 1)$ .

*This defines the cdf of the  $\text{Gumbel}(\mu, \sigma)$  in terms of the cdf of the standard Gumbel, in the same way as the cdf of  $Y \sim N(\mu, \sigma)$  can be expressed by the cdf of the standard normal.*

Recall once more that if  $W \sim \text{Gumbel}(0, 1)$ , then  $W$  has the cdf

$$G(w) = 1 - e^{-e^w}$$

The pdf of  $W$  is hence

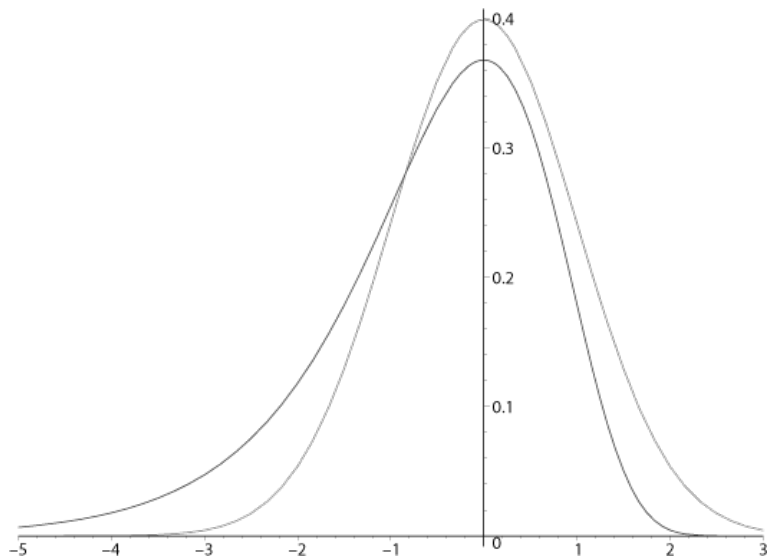
$$g(w) = G'(w) = -e^{-e^w} (-e^w) = e^w e^{-e^w}$$

We also have

$$E(W) = \int_{-\infty}^{\infty} w e^w e^{-e^w} dw = -\gamma,$$

where  $\gamma = -0.5772$  is *Euler's constant*.

# STANDARD GUMBEL AND NORMAL DISTRIBUTIONS





We have seen:

$$T \sim \text{lognorm}(\mu, \sigma) \iff Y = \ln T \sim N(\mu, \sigma)$$

$$T \sim \text{Weib}(\alpha, \theta) \iff Y = \ln T \sim \text{Gumbel}(\mu, \sigma), \text{ with } \mu = \ln \theta, \sigma = 1/\alpha.$$

- Both distributions thus define **log-location-scale families**, which are characterized by the fact that  $Y = \ln T$  has a cdf which can be expressed as

$$F_Y(y) = P(Y \leq y) = \Psi\left(\frac{y - \mu}{\sigma}\right)$$

where  $\Psi(\cdot)$  is the cdf of some “standardized distribution” on  $(-\infty, \infty)$ .

- Equivalently, log-location-scale families are characterized by representations

$$\ln T = \mu + \sigma U$$

where  $U$  has cdf  $\Psi(\cdot)$  as described above.

In the representation

$$\ln T = \mu + \sigma U,$$

- $U$  has a “standard” distribution with support  $(-\infty, +\infty)$ , (e.g.  $N(0, 1)$ , Gumbel(0, 1))
- $\mu \in (-\infty, +\infty)$  is called the *location parameter*
- $\sigma > 0$  is called the *scale parameter*

A random variable  $Y$  has the **logistic distribution** with location parameter  $\mu$  and scale parameter  $\sigma$ , written  $Y \sim \text{logistic}(\mu, \sigma)$ , if

$$F_Y(y) = P(Y \leq y) = H\left(\frac{y - \mu}{\sigma}\right) \quad \text{for } -\infty < y < \infty$$

where

$$H(v) = P(V \leq v) = \frac{e^v}{1 + e^v} \quad \text{for } -\infty < v < \infty$$

is the cdf of the standard logistic distribution,  $\text{logistic}(0, 1)$ .

A lifetime  $T$  has the **log-logistic** distribution with location parameter  $\mu$  and scale parameter  $\sigma$  if  $Y = \ln T \sim \text{logistic}(\mu, \sigma)$ . In this case we have the representation

$$\ln T = \mu + \sigma V$$

where  $V \sim \text{logistic}(0, 1)$ .

Recall that if  $V \sim \text{logistic}(0, 1)$ , then the cdf of  $V$  is  $H(v) = P(V \leq v) = \frac{e^v}{1+e^v}$  for  $-\infty < v < \infty$ .

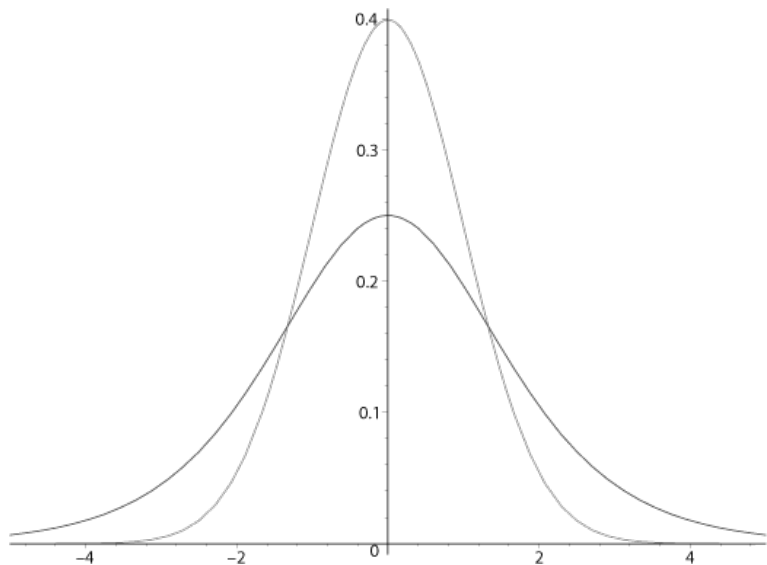
Hence the pdf of  $V$  is

$$h(v) = H'(v) = \frac{e^v}{(1 + e^v)^2} \quad (\text{do the differentiation!})$$

Like the standard normal, this density is symmetric around the  $y$ -axis (which is not the case for the standard Gumbel).

*Check this by showing that  $h(-v) = h(v)$  for all  $v$ .*

# STANDARD LOGISTIC AND STANDARD NORMAL



By assumption,  $Y = \ln T$  has a cdf which can be expressed as

$$F_Y(y) = P(Y \leq y) = \Psi \left( \frac{y - \mu}{\sigma} \right) \quad \text{for } -\infty < y < \infty$$

where  $\Psi(\cdot)$  is the cdf of a standard distribution. Let further  $\psi(u) = \Psi'(u)$ .

Then

$$R(t) = P(T > t) = P(\ln T > \ln t) = 1 - \Psi \left( \frac{\ln t - \mu}{\sigma} \right)$$

$$f(t) = -R'(t) = \psi \left( \frac{\ln t - \mu}{\sigma} \right) \cdot \frac{1}{t\sigma}$$

$$z(t) = \frac{f(t)}{R(t)} = \frac{\psi \left( \frac{\ln t - \mu}{\sigma} \right) / (t\sigma)}{1 - \Psi \left( \frac{\ln t - \mu}{\sigma} \right)}$$

*(as already obtained for the lognormal distribution).*

- *Extreme value distributions*
  - Weibull-distribution
  - Gumbel distribtuion,  $\text{Gumbel}(\mu, \sigma)$
- $T \sim \text{Weibull}(\alpha, \theta) \Rightarrow \ln T \sim \text{Gumbel}(\ln \theta, 1/\alpha)$
- Thus:
  - $\text{Gumbel}(\mu, \sigma)$  is a location-scale family
  - $\text{Weibull}(\alpha, \theta)$  is a log-location-scale family based on Gumbel-distribution
- General definition and properties of log-location-scale families
- Another example: The *logistic* and *log-logistic* distributions