

Plan for the lecture

- ★ Brief summary of the Nelson-Aalen estimator
- ★ Do two groups have different hazard rates?
 - formulated as a hypothesis test
- ★ Using a transformation to find an alternative confidence interval

Multiplicative intensity model

- ★ Multiplicative intensity model

$$\lambda(t) = \alpha(t)Y(t)$$

- $Y(t)$: predictable process
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- ★ Frequent situation leading to a multiplicative intensity model
 - n individuals
 - each individual has the same hazard rate $\alpha(t)$
 - may have truncation and/or censoring
 - $Y(t)$: number of individuals at risk just before time t

Nelson–Aalen estimator

★ $N(t)$: counting process with $\lambda(t) = \alpha(t)Y(t)$

★ Notation:

– $J(t) = \mathbb{I}(Y(t) > 0)$

– $A(t) = \int_0^t \alpha(t) dt$

– $A^*(t) = \int_0^t J(t)\alpha(t) dt$

★ Nelson–Aalen estimator for $A(t)$:

$$\hat{A}(t) = \int_0^t \frac{J(s)}{Y(s)} dN(s) = \sum_{j: T_j \leq t} \frac{1}{Y(T_j)}$$

★ Recall: $E[\hat{A}(t) - A^*(t)] = 0$

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★ Using the optional variation process we found

$$\text{Var}[\hat{A}(t) - A^*(t)] = E \left[\int_0^t \frac{J(s)}{Y^2(s)} dN(s) \right]$$

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$$\text{Var}[\hat{A}(t) - A^*(t)] = E \left[\int_0^t \frac{J(s)}{Y^2(s)} dN(s) \right]$$

★ So an unbiased estimator for $\text{Var}[\hat{A}(t) - A^*(t)]$ is

$$\hat{\sigma}^2(t) = \int_0^t \frac{J(s)}{Y^2(s)} dN(s) = \sum_{j: T_j \leq t} \frac{1}{Y^2(T_j)}$$

Nelson–Aalen with tied observations

- ★ Recall

$$\hat{A}(t) = \int_0^t \frac{J(s)}{Y(s)} dN(s), \hat{\sigma}^2(t) = \int_0^t \frac{J(s)}{Y(s)^2} dN(s)$$

- ★ Why tied observations:

- (i) events happens in continuous time, but we observe ties because of rounding
- (ii) events happens in discrete time

- ★ Denote event times by $T_1 < T_2 < \dots$ and multiplicities $d_1, d_2, \dots$

- ★ Assuming (i) we get

$$\hat{A}(t) = \sum_{j: T_j \leq t} \left[\sum_{\ell=0}^{d_j-1} \frac{1}{Y(T_j) - \ell} \right], \hat{\sigma}^2(t) = \sum_{j: T_j \leq t} \left[\sum_{\ell=0}^{d_j-1} \frac{1}{(Y(T_j) - \ell)^2} \right]$$

- ★ Assuming (ii) we get (we haven't discussed the reason for $\hat{\sigma}^2(t)$)

$$\hat{A}(t) = \sum_{j: T_j \leq t} \frac{d_j}{Y(T_j)}, \hat{\sigma}^2(t) = \sum_{j: T_j \leq t} \frac{(Y(T_j) - d_j)d_j}{Y(T_j)^3}$$

Large sample properties for $\hat{A}(t)$

★ Assume:

- n individuals
- each individual has the same hazard rate $\alpha(t)$
- may have truncation and/or censoring
- $Y(t)$ is number of individuals at risk just before time t

★ Multiplicative intensity process: $\lambda(t) = \alpha(t)Y(t)$

★ Assume also

$$\frac{Y(t)}{n} \rightarrow y(t) > 0 \text{ when } n \rightarrow \infty$$

★ Then Rebolledo's theorem gives that

$$\sqrt{n}(\hat{A}(t) - A^*(t))$$

converges to a Gaussian martingale $U(t)$ with

$$\langle U \rangle(t) = \int_0^t \frac{\alpha(s)}{y(s)} ds$$

★ Thus, for large n : $\hat{A}(t) \approx N(A(t), \sigma^2(t))$

Stochastic integral

★ Stochastic integral

$$\mathbb{I}(t) = \int_0^t H(s) dM(s) = \lim_{n \rightarrow \infty} \sum_{k=1}^n H_k (M_k - M_{k-1})$$

★ Consequences of the definitions

- $\mathbb{I}(t)$ is a mean zero martingale
- $\langle \int H dM \rangle = \int H^2 d\langle M \rangle$
- $[\int H dM] = \int H^2 d[M]$
- $\langle \int H_1 dM_1, \int H_2 dM_2 \rangle = \int H_1 H_2 d\langle M_1, M_2 \rangle$
- $[\int H_1 dM_1, \int H_2 dM_2] = \int H_1 H_2 d[M_1, M_2]$

Counting processes

- ★ $N(t)$: A counting process, $N(t)$ adapted to $\{\mathcal{F}_t\}$
 - right continuous
 - jumps of size one
 - constant between jumps

- ★ Recall informal definition of intensity process

$$\lambda(t)dt = P(dN(t) = 1 | \mathcal{F}_{t-}) = E[dN(t) | \mathcal{F}_{t-}]$$

- ★ Precise definition of intensity process by Doob-Meyer

$$N(t) = \Lambda(t) + M(t)$$

- ★ When $\Lambda(t)$ is absolutely continuous

$$\Lambda(t) = \int_0^t \lambda(s)ds$$

- ★ Variation processes

$$\langle M \rangle(t) = \Lambda(t) = \int_0^t \lambda(s)ds \quad \text{and} \quad [M](t) = N(t)$$