

Department of Mathematical Sciences

# Examination paper for TMA4275 Lifetime Analysis

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Examination time (from-to): 09:00-13:00

**Permitted examination support material:** C: Approved Calculator. One yellow sheet (A4 with stamp) with your own formulae and notes.

## Other information:

Tables of the standard normal distribution are enclosed at the end of the exam.

Language: English Number of pages: 5 Number pages enclosed: 3

Checked by:

Date Signature

#### Problem 1 Graft versus host disease

In a study of patients with a plastic anemia (a condition where the bone marrow is not producing enough new blood cells), 64 patients got a bone marrow transplantation and were then randomized to treatment with (i) methotrexate (MTX) and cyclosporin (CSP), or (ii) only MTX. For each patient one measured the time (in days) from randomization until occurrence of the life threatening complication GVHD (= *Graft Versus Host Disease*). We will here consider the 32 patients who got only MTX, of which 15 experienced GVHD.

The following times to GVHD were observed, where  $\ast$  means a censored observation.

9, 11, 12, 20\*, 20, 20, 22, 25\*, 25, 25, 27\*, 28, 28, 30\*, 31, 35, 35, 41\*, 46, 49, 50\*, 53\*, 54\*, 56\*, 58\*, 60\*, 64\*, 65\*, 66\*, 74\*, 75\*, 77\*

a) Use the (edited) MINITAB-output at the end of this subpoint to draw the Kaplan-Meier (KM) curve for the time to GVHD for these patients.

Suppose we are interested in the probability of having no GVHD in the first three weeks (21 days). First show how this probability can be found from the MINITAB-output, and then show how it is calculated from the observed data above. Also, calculate the standard error of this estimate and find a 95% confidence interval for the probability.

(*Hint:* Greenwood's formula is

$$\left(\hat{R}(t)\right)^2 \sum_{T_{(i)} \le t} \frac{d_i}{n_i(n_i - d_i)}.$$
 )

Kaplan-Meier Estimates

	Number	Number	Survival		
Time	at Risk	Failed	Probability		
9	32	1	0,968750		
11	31	1	0,937500		
12	30	1	0,906250		
20	29	2	0,843750		
22	26	1	0,811298		
25	25	2	0,746394		
28	21	2	0,675309		
31	18	1	0,637792		

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35	17	2	0,562758
46	14	1	0,522561
49	13	1	0,482364

Mean(MTTF) 51,3269

**b)** Use the MINITAB-output to estimate the lower quartile and the median in the distribution of time to GVHD. Why cannot the upper quartile be estimated from this analysis?

The MINITAB output also gives an estimate of the mean time to GVHD. Explain briefly how this estimate is obtained. Which property of the expected value E(T) of a general lifetime T is the basis for the estimate?

c) Give a short explanation of the general property of the KM-estimator,  $\hat{R}(t)$ , that  $\hat{R}(t)$  equals 0 at the last failure time if this is the highest reported time, while  $\hat{R}(t)$  has a strictly positive value at the last failure time if the highest reported time is a censoring. In the latter case, the estimate  $\hat{R}(t)$  remains at this value for all larger t.

Which case appears in the present situation? Based on the KM-curve, how would you estimate the probability of beeing free of GVHD for, say, t larger than 90 days?

Discuss, in view of the above, the practical validity of the estimated mean time to GVHD in subpoint b).

One way to interpret the KM-curve in the present situation is that there are two groups of patients in the study. One group will never experience the GVHD, whereas the rest of the patients will get the GVHD at some time.

Let X be the time when a patient gets GVHD. Let  $X = \infty$  for the patients who will not get GVHD, while X is the actual time of GVHD for the rest. The survival function corresponding to X can then be represented by a so-called *cure model* as:

$$P(X > t) = q + (1 - q)R_0(t), \tag{1}$$

defined for all t > 0. Here q ( $0 \le q < 1$ ) is the probability of not getting GVHD ("be cured"), while  $R_0(t)$  is the survival function for the time to GVHD for the patients who get GVHD ("not cured").

d) Explain briefly why the Kaplan-Meier curve in this Problem can be viewed as the estimator of P(X > t) in (1).

How would you estimate q from the KM-estimate?

Let  $t_r$  be the 100*r*-percentile of the distribution with survival function  $R_0(t)$ , defined by

$$R_0(t_r) = 1 - r.$$

Let  $x_p$  be the 100*p*-percentile of the distribution of X given by (1). Show that

$$x_p = t_{\frac{p}{1-q}}$$

for p < 1 - q.

#### Problem 2 Product complaints

The following table gives data on reported complaints for a particular product, for units that were sold on a particular day (day number 1), at ten different locations. Data on complaints were collected for these units for a period of two weeks (days numbered 1-14).

Location	Number sold	Day number when complaint was received	
1	10	3, 4, 9	
2	14	1,2,7	
3	7	2, 9	
4	5	3	
5	12	1, 4, 12	
6	15	2, 2, 7	
7	10	2, 5	
8	13	1, 3, 6, 10	
9	9	2, 3, 7	
10	5	1	
	Total 100	Total number of complaints: 25	
		Sum of day number of complaints: 108	

For a single unit of this product, let W(t) be the expected number of complaints reported until time t (days) after purchase. It is assumed that all units have the same function W(t). Note that a unit may have more than one complaint in the two weeks period.

a) Suppose you have data for Location 1 only. How would you estimate the function W(t) in this case?

What is the estimate of W(14)? How would you interpret this number?

b) Estimate the function W(t) using all the information in the table. What is the name of the estimator that you use?

What is now the estimate of W(14)?

Give a rough sketch of the estimated curve. What can you say about its shape and how would you interpret this shape? Is this the shape that you would expect?

c) In order to check in a formal way whether there is a trend in the times for complaints, one used the pooled Laplace test.

Calculate the test statistic using all the ten locations.

(*Hint:* The test statistic of the pooled Laplace test simplifies when all processes are observed on the same time interval.)

What is the null hypothesis that is tested by the test? What is a natural alternative hypothesis in the current case?

Calculate the *p*-value for the test and express your conclusion based on it.

#### **Problem 3** The log-logistic distribution

The *log-logistic distribution* (sometimes called the *Fisk distribution*) can be parametrized in such a way that the reliability function is given by

$$R(t) = \frac{1}{1 + \left(\frac{t}{\theta}\right)^{\alpha}} \tag{2}$$

for t > 0, where  $\theta > 0, \alpha > 0$ .

a) Check that the function in (2) satisfies the requirements of a reliability function.

Derive an expression for the density f(t) and show that the hazard rate z(t) can be written

$$z(t) = \frac{\alpha t^{\alpha - 1}}{\theta^{\alpha} + t^{\alpha}} \tag{3}$$

for t > 0.

**b)** Explain how the expression for R(t) in (2) can be used to construct a probability plot for checking whether a right-censored set of data can be assumed to come from a log-logistic distribution.

Explain also briefly how estimates of  $\theta$  and  $\alpha$  can be found from the plot.

c) Show that the hazard rate z(t) in (3) is decreasing in t when  $\alpha \leq 1$ .

Then show by differentiation that, for  $\alpha > 1$ , z(t) is first increasing from 0 to a maximum and then decreasing to 0 as t tends to  $\infty$ .

At what value of t occurs the maximum of z(t)?

The behavior of the hazard rate for  $\alpha > 1$  is similar in shape to the hazard rate of the *lognormal* distribution. The log-logistic distribution is often used as a substitute for the lognormal distribution in analyses.

Which advantages can you point to in using the log-logistic distribution instead of the lognormal distribution?

z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
-3.7	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001
-3.6	.0002	.0002	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001
-3.5	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002
-3.3	.0005	.0005	.0005	.0004	.0004	.0004	.0004	.0004	.0004	.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0005	.0005	.0005
-3.1	.0010	.0009	.0009	.0009	.0008	.0008	.0008	.0008	.0007	.0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9	.0019	.0018	.0018	.0017	.0016	.0016	.0015	.0015	.0014	.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	.0033	.0032	.0031	.0030	.0029	.0028	.0027	.0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	.0060	.0059	.0057	.0055	.0054	.0052	.0051	.0049	.0048
-2.4	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107	.0104	.0102	.0099	.0096	.0094	.0091	.0089	.0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110
-2.1	.0179	.0174	.0170	.0166	.0162	.0158	.0154	.0150	.0146	.0143
-2.0	.0228	.0222	.0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
-1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
-1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
-1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379
9	.1841	.1814	.1788	.1762	.1736	.1711	.1685	.1660	.1635	.1611
8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
7	.2420	.2389	.2358	.2327	.2296	.2266	.2236	.2206	.2177	.2148
6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
5	.3085	.3050	.3015	.2981	.2946	.2912	.2877	.2843	.2810	.2776
4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
3	.3821	.3783	.3745	.3707	.3669	.3632	.3594	.3557	.3520	.3483
2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
1	.4602	.4562	.4522	.4483	.4443	.4404	.4364	.4325	.4286	.4247
0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641

## Standard normal fordeling $\Phi(z) = P(Z \le z)$

1

#### $\Phi(z) = P(Z \le z)$ .02 .03 .00 .01 .04 .05 .06 .07 .08 .09 $\hat{z}$ .5000 .5040 .5080 .5120 .5160 .5199 .5239 .5279 .5319 .5359 .0 .5517 .5596 .5636 .5675 .5714 .5753 .5398 .5438 .5478 .5557 .1 .2 .5793 .5832 .5871 .5910 .5948 .5987 .6026 .6064 .6103 .6141 .3 .6255 .6443 .6179 .6217 .6293 .6331 .6368 .6406 .6480 .6517 .4 .6554 .6591 .6628 .6664 .6700 .6736 .6772 .6808 .6844 .6879 .5 .6915 .6950 .6985 .7019 .7054 .7088 .7123 .7157 .7190 .7224 .7257 .7357 .7389 .7422 .7454 .7486 .7517 .6 .7291 .7324 .7549 .7704 .7734 .7764 .7794 .7 .7580 .7611 .7642 .7673 .7823 .7852 .8 .7881 .7910 .7939 .7967 .7995 .8023 .8051 .8078 .8106 .8133 .9 .8159 .8186 .8212 .8238 .8264 .8289 .8315 .8340 .8365 .8389 1.0.8413 .8438 .8461 .8485 .8508 .8531 .8554 .8577 .8599 .8621 .8643 .8665 .8686 .8708 .8729 .8749 .8770 .8790 .8810 .8830 1.1 .8849 .8869 .8888 .8907 .8925 .8944 .8962 .8980 .8997 1.2 .9015 .9082 .9032 .9049 .9066 .9099 .9115 .9131 .9147 .9162 .9177 1.3 1.4 .9192 .9207 .9222 .9236 .9251 .9265 .9279 .9292 .9306 .9319 1.5 .9332 .9345 .9357 .9370 .9382 .9394 .9406 .9418 .9429 .9441 .9452 .9484 .9495 .9505 .9515 .9525 .9535 .9463 .9474 .9545 1.6 .9554 .9564 .9573 .9582 .9591 .9599 .9608 .9616 .9625 1.7 .9633 .9649 .9664 .9678 .9686 1.8 .9641 .9656 .9671 .9693 .9699 .9706 .9719 .9726 .9732 .9744 .9756 .9761 1.9 .9713 .9738 .9750 .9767 .9778 .9788 .9798 2.0 .9772 .9783 .9793 .9803 .9808 .9812 .9817 2.1 .9821 .9826 .9830 .9834 .9838 .9842 .9846 .9850 .9854 .9857 2.2 .9861 .9864 .9868 .9871 .9875 .9878 .9881 .9884 .9887 .9890 2.3 .9898 .9901 .9906 .9909 .9911 .9893 .9896 .9904 .9913 .9916 2.4 .9918 .9920 .9922 .9925 .9927 .9929 .9931 .9932 .9934 .9936 2.5 .9938 .9940 .9941 .9943 .9945 .9946 .9948 .9949 .9951 .9952 2.6 .9953 .9955 .9956 .9957 .9959 .9960 .9961 .9962 .9963 .9964 2.7 .9965 .9966 .9967 .9968 .9969 .9970 .9971 .9972 .9973 .9974 .9977 .9978 .9979 .9979 .9981 2.8 9974 .9975 .9976 9977 .9980 2.9 .9981 .9982 .9982 .9983 .9984 .9984 .9985 .9985 .9986 .9986 3.0 .9987 .9987 .9988 .9988 .9989 .9989 .9989 .9990 .9987 .9990 .9990 .9992 3.1 .9991 .9991 .9991 .9992 .9992 .9992 .9993 .9993 3.2 .9993 .9993 .9994 .9994 .9994 .9994 .9994 .9995 .9995 .9995 3.3 .9995 .9995 .9995 .9996 .9996 .9996 .9996 .9996 .9996 .9997 3.4 .9997 .9997 .9997 .9997 .9997 .9997 .9997 .9997 .9997 .9998 3.5 .9998 .9998 .9998 .9998 .9998 .9998 .9998 .9998 .9998 .9998 3.6 .9998 .9998 .9999 .9999 .9999 .9999 .9999 .9999 .9999 .9999 3.7 .9999 .9999 .9999 .9999 .9999 .9999 .9999 .9999 .9999 .9999

### Standard normalfordeling

2

$P(Z > z_{\alpha}) = \alpha$					
α	$z_{\alpha}$				
.2	0.842				
.15	1.036				
.1	1.282				
.075	1.440				
.05	1.645				
.04	1.751				
.03	1.881				
.025	1.960				
.02	2.054				
.01	2.326				
.005	2.576				
.001	3.090				
.0005	3.291				
.0001	3.719				
.00005	3.891				
.00001	4.265				
.000005	4.417				
.000001	4.753				

## Kritiske verdier i standard normalfordelingen

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