

Department of Mathematical Sciences

Examination paper for TMA4275 Lifetime Analysis

Academic contact during examination: Ioannis Vardaxis

Phone: 95 36 00 26

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Examination time (from-to): 09:00-13:00

Permitted examination support material: C: Approved Calculator. One yellow sheet (A4 with stamp) with your own formulae and notes.

Other information:

Tables of the χ^2 -distribution with 1 degree of freedom, and of the standard normal distribution are enclosed at the end of the exam.

Language: English Number of pages: 5 Number pages enclosed: 4

Checked by:

Problem 1 Diesel generator fan

Let T be the time to failure for a particular type of diesel generator fans. In a sample of 40 fans one has obtained the following data. Here the Y_i are observed times (in hours), while the δ_i are censoring statuses, where $\delta_i = 1$ if Y_i is a failure time, and $\delta_i = 0$ if it is a right censored time.

i	Y_i	δ_i	i	Y_i	δ_i
1	450	1	21	4850	0
2	1150	1	22	4850	0
3	1600	0	23	5000	0
4	1660	1	24	5000	0
5	1850	0	25	6100	0
6	2030	0	26	6100	0
7	2070	1	27	6100	1
8	2200	0	28	6700	0
9	3000	0	29	7450	0
10	3000	0	30	7800	0
11	3000	0	31	7800	0
12	3100	1	32	8100	0
13	3200	0	33	8100	0
14	3450	1	34	8200	0
15	3750	0	35	8500	0
16	4150	0	36	8500	0
17	4150	0	37	8750	0
18	4300	0	38	8750	1
19	4300	0	39	8750	0
20	4300	0	40	9400	0

a) Let z(t) and Z(t) be, respectively, the hazard function and the cumulative hazard function for T.

Describe briefly how a graph of Z(t) can give us information about the shape of the hazard function z(t).

Calculate and plot the Nelson-Aalen estimator $\hat{Z}_{NA}(t)$ for Z(t) based on the given data.

Does the plot give some indications about the shape of z(t)?

Use the plot to calculate an approximate (constant) value for the hazard rate z(t) for the first 3000 hours. How can you use this value to calculate a rough estimate of the expected time to failure (MTTF) for the fan?

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b) What is the Nelson-Aalen estimate of Z(1800)?

Find the standard error for this estimate. Use it to calculate the standard 95% confidence interval for Z(1800).

(*Hint:* The following formula from the lectures can be used:

$$\widehat{Var(\hat{Z}_{NA}(t))} = \sum_{T_{(i)} \le t} \frac{d_i}{n_i^2} \quad).$$

One also wants an estimate and a 95% confidence interval for the proportion of fans which survive the first 1800 hours, i.e. R(1800), where R(t) is the reliability function for T. Show how you can use the results for Z(1800)calculated above to find both the estimate and the confidence interval for R(1800).

c) Let $\mathcal{T}(t)$ be the Total Time on Test at time t for the given sample.

The table below shows the value of $\mathcal{T}(t)$ at the 8 failure times.

Explain briefly what the value of $\mathcal{T}(t)$ means, and show in detail how the first three values of $\mathcal{T}(t)$ in the table are calculated.

t	450	1150	1660	2070	3100	3450	6100	8750
$\mathcal{T}(t)$	18000	45300	64620	79120	111910	121460	174010	200860

Draw the TTT-plot based on the table. Does the plot give any indications about the shape of z(t)?

Also calculate the test statistic for Barlow-Proschan's test.

Does it lead to rejection of a null hypothesis that z(t) is constant, versus the alternative that it has a monotonic trend? Use significance level 5% for this test.

d) Assume in this subproblem that T is exponentially distributed with unknown hazard rate λ .

The sum of all the 40 observations Y_i is 201 510. Use this information to find the maximum likelihood estimate for λ . Compare it to the rough estimate of z(t) that you found in subproblem a).

Also, find an estimate, as well as an approximate 95% confidence interval, for Z(1800) under the model assumption in this subproblem.

Compare both the estimate and the confidence interval with the corresponding ones obtained in subproblem b). Give a comment.

Problem 2 Repairable system

We consider a repairable system that operates from time t = 0. At failures, the system is repaired and put back into operation immediately.

a) Explain briefly what it means that a repair is, respectively, *minimal* and *perfect*.

Which of these two types of repair is modeled by NHPP? Give a short explanation why an NHPP is an appropriate model in this case.

What kind of model is used for the other type of repair? Give a short explanation.

Let N(t) denote the number of failures in the time interval (0, t]. In the following we assume that $\{N(t) : t > 0\}$ is a non-homogeneous Poisson process (NHPP) with intensity function (ROCOF) $w(t; \theta)$ and cumulative intensity function (CROCOF) $W(t; \theta) = \int_0^t w(u; \theta) du$. Here θ is a parameter, possibly a vector.

In order to do statistical inference about θ , one observes the system in the time interval $[0, \tau]$, for some given $\tau > 0$. The data consist of number of failures D_i in time intervals $(h_{i-1}, h_i]$, for i = 1, 2, ..., r, where

$$h_0 = 0 < h_1 < h_2 < \dots < h_r = \tau.$$

b) Explain why we can conclude that D_1, \ldots, D_r are independent and Poissondistributed, with

$$E(D_i) = W(h_i; \theta) - W(h_{i-1}; \theta) \text{ for } i = 1, 2 \dots, r.$$

Use this to show that the likelihood function for θ , when the observations of D_1, \ldots, D_r are d_1, \ldots, d_r , can be written

$$L(\theta) = \left\{ \prod_{i=1}^{r} \frac{[W(h_i; \theta) - W(h_{i-1}; \theta)]^{d_i}}{d_i!} \right\} e^{-W(\tau; \theta)}.$$
 (1)

For a machine one has recorded number of failures per week for r = 10 weeks. Let the weeks be numbered 1, 2, ..., 10, and let d_i be the number of failures observed in week number i, i = 1, ..., 10. Page 4 of 5

It is assumed that the failures occur as an NHPP with cumulative intensity function

$$W(t; \lambda, \beta) = \lambda t^{\beta}$$
 for $t > 0$,

where the time unit is *weeks*, and $\lambda > 0, \beta > 0$ are unknown parameters.

The observed data are given by the following table:

Week no. i	1	2	3	4	5	6	7	8	9	10
Number of failures d_i	0	3	4	5	3	1	3	5	7	4

c) Use the formula (1) to show that the log-likelihood of the given data can be written as

$$\ell(\lambda,\beta) = (\ln \lambda) \sum_{i=1}^{10} d_i + \sum_{i=1}^{10} d_i \ln \left(i^\beta - (i-1)^\beta \right) - \sum_{i=1}^{10} \ln(d_i!) - \lambda \cdot 10^\beta.$$

(*Hint*: Let $h_i = i$ for i = 0, 1, ..., 10).

Show that if β is known, then the maximum likelihood estimator for λ is

$$\hat{\lambda}(\beta) = \frac{\sum_{i=1}^{10} d_i}{10^{\beta}}.$$

Compute the value of $\hat{\lambda}(\beta)$ with the given data if it is known that $\beta = 1.3$. With these values of β and λ , what are the probabilities of no failures in, respectively, the first week and in the 10th week?

d) Calculate the profile log likelihood function $\tilde{\ell}(\beta)$ of β .

A graph of $\ell(\beta)$ with the given data is given on the next page.

You may use that the maximum value of $\tilde{\ell}(\beta)$ is -18.60.

Use the graph (in a rough manner)

- 1. to find the maximum likelihood estimate of β ,
- 2. to find an approximate 95% confidence interval of β ,
- 3. to test, with significance level 5%, the null hypothesis of no time trend in the occurrence of failures of the machine.

Find also the maximum likelihood estimate of λ .

(In your exam paper you may illustrate all this by a rough sketch of the graph).



beta

x	$P(X \le x)$	x	$P(X \le x)$
0.01	0.079656	2.7	0.899652
0.02	0.112463	2.8	0.905736
0.03	0.137510	2.9	0.911420
0.04	0.158519	3.0	0.916735
0.05	0.176937	3.1	0.921708
0.06	0.193504	3.2	0.926362
0.07	0.208663	3.3	0.930720
0.08	0.222703	3.4	0.934804
0.09	0.235823	3.5	0.938631
0.10	0.248170	3.6	0.942220
0.20	0.345279	3.7	0.945588
0.30	0.416118	3.8	0.948747
0.40	0.472911	3.9	0.951714
0.50	0.520500	4.0	0.954500
0.60	0.561422	4.1	0.957117
0.70	0.597216	4.2	0.959576
0.80	0.628907	4.3	0.961888
0.90	0.657218	4.4	0.964061
1.00	0.682689	4.5	0.966105
1.10	0.705734	4.6	0.968028
1.20	0.726678	4.7	0.969837
1.30	0.745787	4.8	0.971540
1.40	0.763276	4.9	0.973143
1.50	0.779329	5.0	0.974653
1.60	0.794097	5.1	0.976074
1.70	0.807712	5.2	0.977413
1.80	0.820288	5.3	0.978675
1.90	0.831922	5.4	0.979863
2.00	0.842701	5.5	0.980984
2.10	0.852701	5.6	0.982040
2.20	0.861989	5.7	0.983035
2.30	0.870626	5.8	0.983974
2.40	0.878665	5.9	0.984859
2.50	0.886154	6.0	0.985694
2.60	0.893136	6.1	0.986482

Table of cumulative probabilities of the $\chi^2\text{-distribution}$ with 1 degree of freedom.

	$\Phi(z) = P(Z \le z)$									
z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
-3.7	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001
-3.6	.0002	.0002	.0001	.0001	.0001	.0001	.0001	.0001	.0001	.0001
-3.5	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002	.0002
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002
-3.3	.0005	.0005	.0005	.0004	.0004	.0004	.0004	.0004	.0004	.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0005	.0005	.0005
-3.1	.0010	.0009	.0009	.0009	.0008	.0008	.0008	.0008	.0007	.0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9	.0019	.0018	.0018	.0017	.0016	.0016	.0015	.0015	.0014	.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	.0033	.0032	.0031	.0030	.0029	.0028	.0027	.0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	.0060	.0059	.0057	.0055	.0054	.0052	.0051	.0049	.0048
-2.4	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107	.0104	.0102	.0099	.0096	.0094	.0091	.0089	.0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110
-2.1	.0179	.0174	.0170	.0166	.0162	.0158	.0154	.0150	.0146	.0143
-2.0	.0228	.0222	.0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
-1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
-1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
-1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379
9	.1841	.1814	.1788	.1762	.1736	.1711	.1685	.1660	.1635	.1611
8	.2119	.2090	.2061	.2033	.2005	.1977	.1949	.1922	.1894	.1867
7	.2420	.2389	.2358	.2327	.2296	.2266	.2236	.2206	.2177	.2148
6	.2743	.2709	.2676	.2643	.2611	.2578	.2546	.2514	.2483	.2451
5	.3085	.3050	.3015	.2981	.2946	.2912	.2877	.2843	.2810	.2776
4	.3446	.3409	.3372	.3336	.3300	.3264	.3228	.3192	.3156	.3121
3	.3821	.3783	.3745	.3707	.3669	.3632	.3594	.3557	.3520	.3483
2	.4207	.4168	.4129	.4090	.4052	.4013	.3974	.3936	.3897	.3859
1	.4602	.4562	.4522	.4483	.4443	.4404	.4364	.4325	.4286	.4247
0	.5000	.4960	.4920	.4880	.4840	.4801	.4761	.4721	.4681	.4641

Standard normalfordeling

1

	$\Phi(z) = P(Z \le z)$									
z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
.1	.5398	.5438	.5478	.5517	.5557	.5596	.5636	.5675	.5714	.5753
.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
.3	.6179	.6217	.6255	.6293	.6331	.6368	.6406	.6443	.6480	.6517
.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
.5	.6915	.6950	.6985	.7019	.7054	.7088	.7123	.7157	.7190	.7224
.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	.7549
.7	.7580	.7611	.7642	.7673	.7704	.7734	.7764	.7794	.7823	.7852
.8	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
.9	.8159	.8186	.8212	.8238	.8264	.8289	.8315	.8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	.8621
1.1	.8643	.8665	.8686	.8708	.8729	.8749	.8770	.8790	.8810	.8830
1.2	.8849	.8869	.8888	.8907	.8925	.8944	.8962	.8980	.8997	.9015
1.3	.9032	.9049	.9066	.9082	.9099	.9115	.9131	.9147	.9162	.9177
1.4	.9192	.9207	.9222	.9236	.9251	.9265	.9279	.9292	.9306	.9319
1.5	.9332	.9345	.9357	.9370	.9382	.9394	.9406	.9418	.9429	.9441
1.6	.9452	.9463	.9474	.9484	.9495	.9505	.9515	.9525	.9535	.9545
1.7	.9554	.9564	.9573	.9582	.9591	.9599	.9608	.9616	.9625	.9633
1.8	.9641	.9649	.9656	.9664	.9671	.9678	.9686	.9693	.9699	.9706
1.9	.9713	.9719	.9726	.9732	.9738	.9744	.9750	.9756	.9761	.9767
2.0	.9772	.9778	.9783	.9788	.9793	.9798	.9803	.9808	.9812	.9817
2.1	.9821	.9826	.9830	.9834	.9838	.9842	.9846	.9850	.9854	.9857
2.2	.9861	.9864	.9868	.9871	.9875	.9878	.9881	.9884	.9887	.9890
2.3	.9893	.9896	.9898	.9901	.9904	.9906	.9909	.9911	.9913	.9916
2.4	.9918	.9920	.9922	.9925	.9927	.9929	.9931	.9932	.9934	.9936
2.5	.9938	.9940	.9941	.9943	.9945	.9946	.9948	.9949	.9951	.9952
2.6	.9953	.9955	.9956	.9957	.9959	.9960	.9961	.9962	.9963	.9964
2.7	.9965	.9966	.9967	.9968	.9969	.9970	.9971	.9972	.9973	.9974
2.8	.9974	.9975	.9976	.9977	.9977	.9978	.9979	.9979	.9980	.9981
2.9	.9981	.9982	.9982	.9983	.9984	.9984	.9985	.9985	.9986	.9986
3.0	.9987	.9987	.9987	.9988	.9988	.9989	.9989	.9989	.9990	.9990
3.1	.9990	.9991	.9991	.9991	.9992	.9992	.9992	.9992	.9993	.9993
3.2	.9993	.9993	.9994	.9994	.9994	.9994	.9994	.9995	.9995	.9995
3.3	.9995	.9995	.9995	.9996	.9996	.9996	.9996	.9996	.9996	.9997
3.4	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9997	.9998
3.5	.9998	.9998	.9998	.9998	.9998	.9998	.9998	.9998	.9998	.9998
3.6	.9998	.9998	.9999	.9999	.9999	.9999	.9999	.9999	.9999	.9999
3.7	.9999	.9999	.9999	.9999	.9999	.9999	.9999	.9999	.9999	.9999

Standard normalfordeling

2

P(Z > z	$(\alpha) = \alpha$
α	z_{α}
.2	0.842
.15	1.036
.1	1.282
.075	1.440
.05	1.645
.04	1.751
.03	1.881
.025	1.960
.02	2.054
.01	2.326
.005	2.576
.001	3.090
.0005	3.291
.0001	3.719
.00005	3.891
.00001	4.265
.000005	4.417
.000001	4.753

Kritiske verdier i standard normalfordelingen

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