

Conjugate priors

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- ★ A distribution $f(x|\theta)$ belongs to the one-parameter exponential family if

$$f(x|\theta) = a(x)e^{\phi(\theta)t(x)+b(\theta)}$$

for some functions $a(x)$, $\phi(\theta)$, $t(x)$ and $b(\theta)$.

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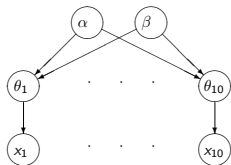
- ★ Today:
 - conjugate prior for a k -parameter exponential family
 - conditional conjugacy

Recall: Bayesian example

- ★ A simple example (from George et al., 1993)
 - Analysis of 10 power plant pumps
 - x_i, t_i : number of failures for pump i and length of operation time on that pump (in 1000 hours)
 - Modelling:
 - + $x_i | \theta_i \sim \text{Poisson}(\theta_i; t_i)$
 - + conjugate prior for θ_i : $\theta_i | \alpha, \beta \sim \text{Gamma}(\alpha, \beta)$
 - + hyper-prior distribution on α and β

$$\alpha \sim \text{Exp}(1.0) , \beta \sim \text{Gamma}(0.1, 1.0)$$

- graphical model:

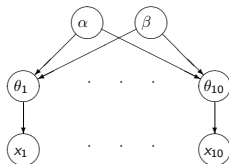


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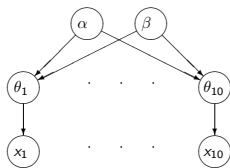


- observed: x_1, \dots, x_n
- posterior distribution of interest:

$$f(\alpha, \beta, \theta_1, \dots, \theta_{10} | x_1, \dots, x_{10})$$

Posterior distribution of interest

- ★ Graphical model:



- ★ Posterior distribution of interest:

$$\begin{aligned} f(\alpha, \beta, \theta_1, \dots, \theta_{10} | x_1, \dots, x_{10}) &= \frac{f(\alpha, \beta, \theta_1, \dots, \theta_{10}, x_1, \dots, x_{10})}{f(x_1, \dots, x_{10})} \\ &\propto f(\alpha, \beta, \theta_1, \dots, \theta_{10}, x_1, \dots, x_{10}) \\ &\propto f(\alpha) f(\beta) \prod_{i=1}^{10} f(\theta_i | \alpha, \beta) \prod_{i=1}^{10} f(x_i | \theta_i) \\ &\propto e^{-\alpha} \cdot \beta^{-0.9} e^{-\beta} \cdot \left[\prod_{i=1}^{10} \frac{\beta^\alpha}{\Gamma(\alpha)} (\theta_i)^{\alpha-1} e^{-\beta \theta_i} \right] \cdot \left[\prod_{i=1}^{10} (\theta_i t_i)^{x_i} e^{-\theta_i t_i} \right] \end{aligned}$$

- ★ How to do MCMC from $f(\alpha, \beta, \theta_1, \dots, \theta_{10} | x_1, \dots, x_{10})$?

Single site updates

- ★ Posterior distribution: $f(\alpha, \beta, \theta_1, \dots, \theta_{10} | x_1, \dots, x_{10})$

$$\propto e^{-\alpha} \cdot \beta^{-0.9} e^{-\beta} \cdot \left[\prod_{i=1}^{10} \frac{\beta^\alpha}{\Gamma(\alpha)} (\theta_i)^{\alpha-1} e^{-\beta\theta_i} \right] \cdot \left[\prod_{i=1}^{10} (\theta_i t_i)^{x_i} e^{-\theta_i t_i} \right]$$

- ★ Full conditional for θ_k :

$$f(\theta_k | \dots) \propto$$

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$$\begin{aligned} f(\theta_k | \dots) &\propto \theta_k^{\alpha-1} e^{-\beta\theta_k} \theta_k^{x_k} e^{-\theta_k t_k} \\ &\propto \theta_k^{\alpha+x_k-1} e^{-(\beta+t_k)\theta_k} \end{aligned}$$

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- ★ Thus:

- $\theta_k | \dots \sim \text{Gamma}(\alpha + x_k, \beta + t_k)$
- we can use a Gibbs update for θ_k

Single site updates

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$$f(\beta | \dots) \propto$$

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$$\begin{aligned} f(\beta | \dots) &\propto \beta^{-0.9} e^{-\beta} \prod_{i=1}^{10} [\beta^\alpha e^{-\beta\theta_i}] \\ &\propto \beta^{-0.9+10\alpha} e^{-(1+\sum_{i=1}^{10} \theta_i)\beta} \end{aligned}$$

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- ★ Thus:

- $\beta | \dots \sim \text{Gamma} \left(0.1 + 10\alpha, 1 + \sum_{i=1}^{10} \theta_i \right)$
- we can use a Gibbs update for β

Single site updates

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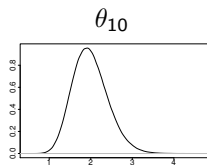
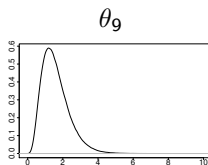
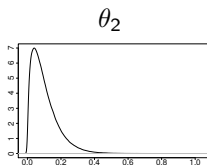
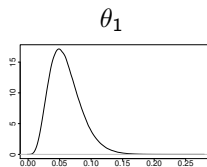
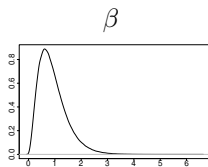
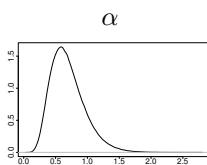
- $\alpha | \dots$ is not a familiar distribution, neither easy to sample from
- we can use a random walk proposal for α
- get an algorithm with one tuning parameter

Simulation output — after burn-in

★ Data:

Pump	1	2	3	4	5	6	7	8	9	10
t_i	94.3	15.7	62.9	126	5.24	31.4	1.05	1.05	2.1	10.5
x_i	5	1	5	14	3	19	1	1	4	22

★ Posterior density plots:



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– Posterior mean for θ_i compared to x_i/t_i

parameter	posterior mean	x_i/t_i
θ_1	0.0598	0.0530
θ_2	0.1017	0.0636
θ_3	0.0892	0.0795
θ_4	0.1157	0.1111
θ_5	0.6011	0.5725
θ_6	0.6095	0.6051
θ_7	0.8910	0.9524
θ_8	0.8928	0.9524
θ_9	1.5867	1.9047
θ_{10}	1.9901	2.0952

