

Multivariate regression with orthogonal X

$$\begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & x_{11} & x_{21} & \cdots & x_{k1} \\ 1 & x_{12} & x_{22} & \cdots & x_{k2} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_{1n} & x_{2n} & \cdots & x_{kn} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_k \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{bmatrix}$$

X has orthogonal columns means

$$(*) \sum_{i=1}^n x_{ji} x_{li} = 0 \text{ when } j \neq l$$

$$(**) \sum_{i=1}^n x_{li} = 0 \text{ for } l=1, \dots, k \text{ (column sums are 0 because of first column is } \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix})$$

Estimation of β :

$$Q = \sum_{i=1}^n (y_i - b_0 - b_1 x_{1i} - \cdots - b_k x_{ki})^2$$

$$\begin{aligned} \frac{\partial Q}{\partial b_0} &= -2 \sum_{i=1}^n (y_i - b_0 - b_1 x_{1i} - \cdots - b_k x_{ki}) \\ &= -2 \left[\sum_{i=1}^n y_i - n b_0 \right] \text{ by } (**) \end{aligned}$$

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$$\text{so } \frac{\partial Q}{\partial b_0} = 0 \Leftrightarrow \sum_{i=1}^n y_i = n b_0 \Leftrightarrow b_0 = \frac{1}{n} \sum_{i=1}^n y_i = \bar{y}$$

$$\begin{aligned} \frac{\partial Q}{\partial b_1} &= -2 \sum_{i=1}^n (y_i - b_0 - b_1 x_{1i} - \cdots - b_k x_{ki}) x_{1i} \\ &= -2 \left[\sum_{i=1}^n x_{1i} y_i - b_1 \sum_{i=1}^n x_{1i}^2 \right] \text{ by } (*) \text{ and } (**) \end{aligned}$$

$$\text{so } \frac{\partial Q}{\partial b_1} = 0 \Leftrightarrow \sum_{i=1}^n x_{1i} y_i = b_1 \sum_{i=1}^n x_{1i}^2 \Leftrightarrow b_1 = \frac{\sum_{i=1}^n x_{1i} y_i}{\sum_{i=1}^n x_{1i}^2}$$

Similarly

$$b_j = \frac{\sum_{i=1}^n x_{ji} y_i}{\sum_{i=1}^n x_{ji}^2} \text{ for } j=1, 2, \dots, k$$

$$\text{Var}(b_j) = \frac{1}{\left(\sum_{i=1}^n x_{ji}^2 \right)^2} \cdot \sum_{i=1}^n x_{ji}^2 \underbrace{\text{Var}(y_i)}_{\sigma^2} = \frac{\sigma^2}{\sum_{i=1}^n x_{ji}^2}$$

$$\text{Var}(b_j) = \frac{\sigma^2}{\sum_{i=1}^n x_{ji}^2}$$

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Estimation of σ^2 etc.:

$$SST = SSR + SSE$$

$$SSR = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2$$

$$= \sum_{i=1}^n (b_0 + b_1 x_{1i} + \dots + b_k x_{ki} - b_0)^2$$

$$= \sum_{i=1}^n (b_1 x_{1i} + \dots + b_k x_{ki})^2 \quad \leftarrow \text{product terms cancel by (*)}$$

so
$$SSR = b_1^2 \sum_{i=1}^n x_{1i}^2 + \cancel{b_1 b_2} \sum_{i=1}^n x_{2i}^2 + \dots + b_k^2 \sum_{i=1}^n x_{ki}^2$$

by (*)

or
$$SSR = R(\beta_1) + R(\beta_2) + \dots + R(\beta_k)$$

[while in general = $R(\beta_1) + R(\beta_2 | \beta_1) + R(\beta_3 | \beta_1, \beta_2) + \dots$
 $+ R(\beta_k | \beta_1, \beta_2, \dots, \beta_{k-1})$]

Thus
$$SSE = SST - SSR = \sum_{i=1}^n (y_i - \bar{y})^2 - b_1^2 \sum_{i=1}^n x_{1i}^2 - \dots - b_k^2 \sum_{i=1}^n x_{ki}^2$$

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Thus to test

$H_0: \beta_j = 0$ we use

$$z_j = \frac{b_j}{\frac{\sigma}{\sqrt{\sum_{i=1}^n x_{ji}^2}}} = \frac{\sqrt{\sum_{i=1}^n x_{ji}^2} \cdot b_j}{\sigma} \sim N(0,1) \quad \underline{\text{if}} \quad \beta_j = 0$$

so $z_j^2 = \frac{R(\beta_j)}{\sigma^2} \sim \chi_1^2 \quad \underline{\text{if}} \quad \beta_j = 0$

and $\frac{SSR}{\sigma^2} \sim \chi_k^2 \quad \underline{\text{if}} \quad \beta_1 = \dots = \beta_k = 0$ (no regression)

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