TMA4265 Stochastic Processes Week 36 – Solutions

Problem 2: Changes in stock prices

The model in the problem can be written as

$$Z = X_0 + \sum_{i=1}^N X_i,$$

where X_0 is always included and one or more X_i , for $i \ge 1$, may be included. The desired variance can be calculated via the law of total variance,

$$Var[Z] = E[Var[Z|N]] + Var[E[Z|N]]$$
$$= E[\sigma^2 + N\sigma^2] + Var[0 + N \cdot 0]$$
$$= \sigma^2 + \nu\sigma^2 = (1 + \nu)\sigma^2.$$

Problem 3: Joint distribution

1.

$$p_X(x) = \sum_y p(x, y) = \sum_y \exp(-2\lambda) \frac{\lambda^{x+y}}{x! y!}$$
$$= \exp(-\lambda) \frac{\lambda^x}{x!} \underbrace{\sum_y \exp(-\lambda) \frac{\lambda^y}{y!}}_{1}$$
$$= \exp(-\lambda) \frac{\lambda^x}{x!}$$

Hence, X is Poisson distributed with parameter λ , e.g. $X \sim \mathcal{P}(\lambda)$. Analogously, $Y \sim \mathcal{P}(\lambda)$. We find that X and Y are independent, since $p(x, y) = p_X(x)p_Y(y)$ is fulfilled. Hence,

$$\operatorname{Cov}(X,Y) = E(XY) - E(X)E(Y) \stackrel{X \text{ and } Y \text{ indep.}}{=} E(X)E(Y) - E(X)E(Y) = 0,$$

 \Rightarrow the covariance of X and Y is zero.

2. We know from the lecture that $X + Y \sim P(2\lambda)$. Hence

$$P(X|Z = X + Y) = \frac{P(X = x, Z = x + Y)}{P(Z = z)}$$

$$= \frac{P(X = x, Z = x + Y)}{P(Z = z)}$$

$$= \frac{P(X = x, Y = z - x)}{P(Z = z)}$$

$$= \frac{P(X = x)P(Y = z - x)}{P(Z = z)}$$

$$= \frac{\exp(-\lambda)\frac{\lambda^{x}}{x!}\exp(-\lambda)\frac{\lambda^{z-x}}{(z-x)!}}{\exp(-2\lambda)\frac{(2\lambda)^{z}}{z!}}$$

$$= \frac{\lambda^{z}}{\frac{x!(x-z)!}{\frac{(2\lambda)^{z}}{z!}}}$$

$$= \frac{z!}{x!(x-z)!} \left(\frac{1}{2}\right)^{z}$$

$$= \binom{z}{x} \left(\frac{1}{2}\right)^{z}$$

$$= \binom{z}{x} \left(\frac{1}{2}\right)^{x} \left(\frac{1}{2}\right)^{z-x}$$

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Hence, X|X+Y is binomially distributed with size z and probability 0.5, e.g. $X|X+Y \sim \mathcal{B}(z, 0.5)$.

$$Cov(X + Y, X - Y) = E((X + Y)(X - Y)) - E(X + Y)E(X - Y)$$

= $E(X^2 - Y^2) - (E(X) + E(Y)) \cdot \underbrace{(E(X) - E(Y))}_{0}$
= $E(X^2) - E(Y^2)$
= 0

Problem 4: Expectation

1. This exercise is a special case of Example 3.15 in the book. Let N_2 be the number of necessary rolls until two consecutive sixes appear, and let M_2 denote its mean. We condition on N_1 the number of trials needed for one six. Hence

$$M_2 = E(N_2) = E(E(N_2|N_1)),$$

where

3.

$$E(N_2|N_1) = \underbrace{p(N_1+1)}_{\text{case 1}} + \underbrace{(1-p)(N_1+1+E(N_2))}_{\text{case 2}}$$
$$= p \cdot N_1 + p + N_1 + 1 - p \cdot N_1 - p + (1-p)E(N_2)$$
$$= N_1 + 1 + (1-p)E(N_2)$$

It takes N_1 rolls to get one six, then either the next roll is a six (with probability $p = \frac{1}{6}$) as well, and we are done (case 1), or it is not a six (with probability $1 - p = \frac{5}{6}$) and we must begin anew (case 2). For case 2, it is important to have in mind that we have already needed $N_1 + 1$ rolls to get that far.

Taking expectations of both sides of the preceding yields

$$M_2 = M_1 + 1 + (1 - p)M_2$$

or

$$M_2 = \frac{M_1 + 1}{p}.$$

Since N_1 , the time of the first six, is geometric with parameter p we see that

$$M_1 = \frac{1}{p} = \frac{1}{\frac{1}{6}} = 6,$$

and thus

$$M_2 = \frac{6+1}{\frac{1}{6}} = 42.$$

The expected of rolls we need until the firs pair of consecutive sixes appears is 42.