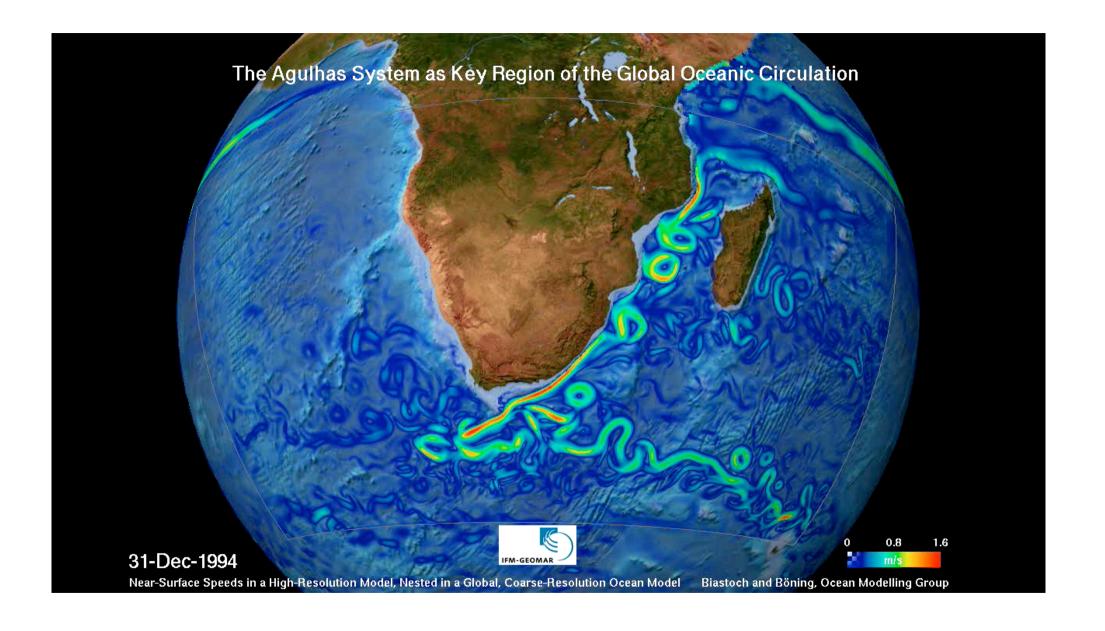
# Particle Filters for high-dimensional problems

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## How is DA used today in geosciences?

Present-day data-assimilation systems are based on linearizations and state covariances are essential.

#### 4DVar, Representer method (PSAS):

- Gaussian pdf's for the state, solves only for posterior mode, needs error covariance of initial state (B matrix)

#### (Ensemble) Kalman filter:

- assumes Gaussian pdf's for the state, approximates posterior mean and covariance, doesn't minimize anything in nonlinear systems, needs inflation and localisation

Combinations of these: hybrid methods

## Nonlinear filtering: Particle filter

$$p(x|y) = \frac{p(y|x)p(x)}{\int p(y|x)p(x) dx}$$

Use ensemble 
$$p(x) = \sum_{i=1}^{N} \frac{1}{N} \delta(x - x_i)$$

$$p(x|y) = \sum_{i=1}^{N} w_i \delta(x - x_i)$$

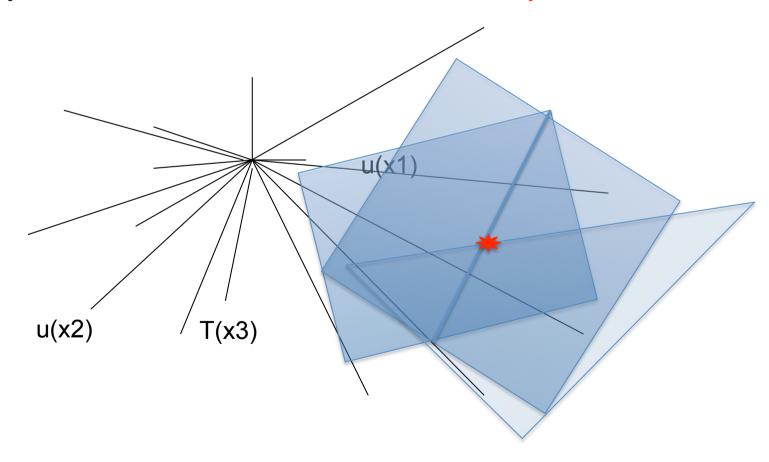
with

$$w_i = \frac{p(y|x_i)}{\sum_j p(y|x_j)}$$

the weights.

## Why are particle filters degenerate I

Probability space in large-dimensional systems is 'empty': the curse of dimensionality



## Why are Particle Filters degenerate II

 The volume of a hypersphere of radius r in an M dimensional space is

$$V \propto \frac{r^M}{\Gamma(M/2 - 1)}$$

• Taking for the radius  $r \approx 3\sigma_y$  we find, using Stirling:

$$V \propto \left[\frac{9\sigma_y}{M/2}\right]^{M/2}$$

So very small indeed.

## Why are Particle Filters degenerate III

For the optimal proposal density we find, for Gaussian process model and Gaussian observation errors:

$$w_i \propto p(y^n | x_i^{n-1})$$

$$\propto \exp\left[-\frac{1}{2}(y^n - Hf(x_i^{n-1}))(HQH^T + R)^{-1}\right]$$

$$\times (y^n - Hf(x_i^{n-1})).$$

Ignoring covariances we find:

$$var[-\log(w_i)] \propto \frac{M}{2} \left( \frac{V_x}{V_\beta + V_y} \right)^2 \left( 1 + 2 \left( \frac{V_y + V_\beta}{V_x} \right) \right)$$

## Why are Particle Filters degenerate?

- 'Number of particles needed grows exponentially with dimension of the state vector (Bickel et al, 2007).'
- A slightly different view: degeneracy due to number of independent observations.
- This is related to the extremely narrow likelihood, a tiny move of a particle gives a completely different weight.

#### The statistics

• The Stochastic PDE: 
$$x^n = f(x^{n-1}) + \beta^{n-1}$$

Observations:

$$y^n$$

• Relation between the two: 
$$y^n = H(x^n) + \epsilon^n$$

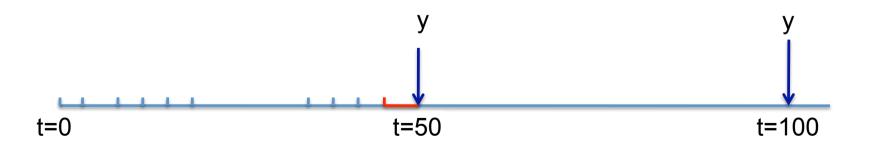
Assume:  $\beta \sim N(0,Q)$ 

$$\epsilon \sim N(0,R)$$

H is linear

## The Equivalent-Weights Particle Filter

- Use simple proposal at each time step, e.g. relaxation to observations.
- Use different proposal at final time step to ensure that weights are very similar.



## Transition density for $x^n$

Stochastic model

$$x^{n} = f(x^{n-1}) + \beta^{n-1}$$

With

$$\beta^{n-1} \sim N(0, Q)$$

Hence transition density

$$p(x^n|x^{n-1}) = N(f(x^{n-1}), Q)$$

### Bayes Theorem and the proposal density

Bayes Theorem can be written as:

$$p(x^{n}|y^{n}) = \frac{p(y^{n}|x^{n})p(x^{n})}{p(y)}$$

$$= \frac{p(y^{n}|x^{n})}{p(y)} \int p(x^{n}|x^{n-1})p(x^{n-1}) dx^{n-1}$$

Multiply and divide this expression by a proposal transition density *q*:

$$p(x^n|y^n) = \frac{p(y^n|x^n)}{p(y)} \int \frac{p(x^n|x^{n-1})}{q(x^x|x^{n-1}, y^n)} q(x^n|x^{n-1}, y^n) p(x^{n-1}) dx^{n-1}$$

## Proposal transition density

For each particle at time n-1 draw a sample from the proposal transition density q, to find:

$$p(x^{n}|y^{n}) = \frac{1}{N} \sum_{i=1}^{N} \frac{p(y^{n}|x_{i}^{n})}{p(y)} \frac{p(x_{i}^{n}|x_{i}^{n-1})}{q(x_{i}^{n}|x_{i}^{n-1}, y^{n})} \delta(x^{n} - x_{i}^{n})$$

Which can be rewritten as:

$$p(x^n|y^n) = \sum_{i=1}^N w_i \delta(x^n - x_i^n)$$

with weights

$$w_{i} = \frac{p(y^{n}|x_{i}^{n})}{p(y^{n})} \frac{p(x_{i}^{n}|x_{i}^{n-1})}{q(x_{i}^{n}|x_{i}^{n-1}, y^{n})}$$

Likelihood weight

Proposal weight

## Proposal density between observations

We can explore the fact that the model needs several O(100) time steps between observations, e.g. by using a relaxation term in the proposal:

$$q(x^{n}|x_{i}^{n-1}, y^{m}) = N\left(f(x_{i}^{n-1}) + S\left(y^{m} - H(x_{i}^{n-1})\right), Q\right)$$

Corresponding to an evolution equation for each particle

$$x_i^n = f(x_i^{n-1}) + \hat{\beta}_i^n + S(y^n - H(x_i^{n-1}))$$

## Proposal density between observations

- One could also use the 'optimal proposal density' between observations.
- This can be implemented as a minimization method for each particle, and is also known as the Implicit Particle Filter.
- This is related to a method called 4DVar in meteorology and oceanography, which explores only the mode of the joint-in-time pdf.

## Proposal density at observation time: the essence of the Equivalent-Weights Particle Filter

The proposal density depends on the maximum weight a particle can achieve using a deterministic time step. It is defined as:

$$q(x^{n}|x_{i}^{n-1}, y^{n}) = \begin{cases} q_{1}(x^{n}|x_{i}^{n-1}, y^{n}) & \text{if } w_{i}^{max} > w^{target} \\ q_{2}(x^{n}|x_{i}^{n-1}, y^{n}) & \text{if } w_{i}^{max} < w^{target} \end{cases}$$

The target weight is set by the user, as e.g. the weight that 80% of the particles can achieve.

## The maximum weights

1. We know:

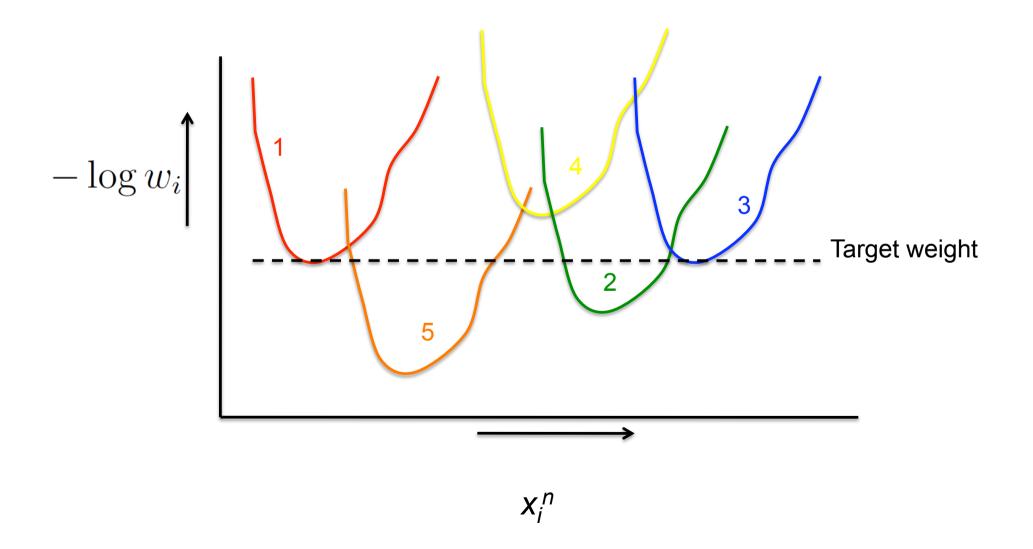
$$w_i = \frac{p(y^n | x_i^n)}{p(y^n)} \frac{p(x_i^n | x_i^{n-1})}{q(x_i^n | x_i^{n-1}, y^n)}$$

2. Write down expression for each weight ignoring proposal:

$$w_i \propto w_i^{rest} \exp \left[ -\frac{1}{2} \left( x_i^n - f(x_i^{n-1}) \right)^T Q^{-1} \left( x_i^n - f(x_i^{n-1}) \right) - \frac{1}{2} (y^n - H(x_i^n))^T R^{-1} (y^n - H(x_i^n)) \right]$$

3. When H is linear this is a quadratic function in  $x_i^n$  for each particle. Otherwise linearize.

## The target weight



#### The Equivalent-Weights Particle Filter

The proposal density is chosen as:

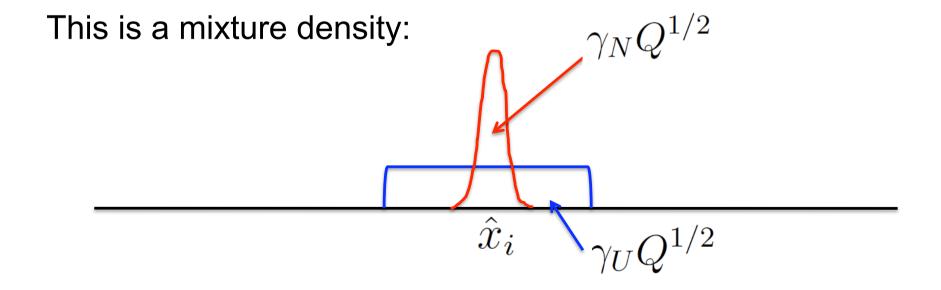
$$q(x^{n}|x_{i}^{n-1}, y^{n}) = \begin{cases} q_{1}(x^{n}|x_{i}^{n-1}, y^{n}) & \text{if } w_{i}^{max} > w^{target} \\ q_{2}(x^{n}|x_{i}^{n-1}, y^{n}) & \text{if } w_{i}^{max} < w^{target} \end{cases}$$

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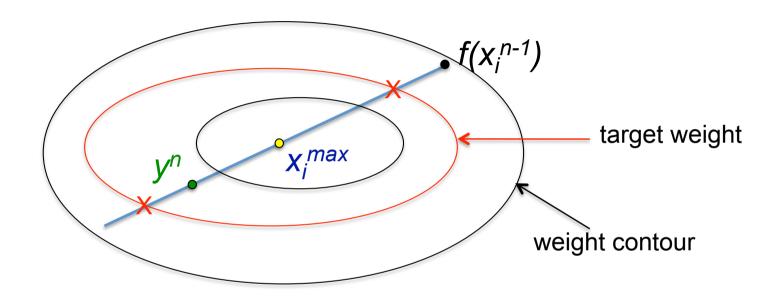
## The two proposal densities

For particles that can reach the target weight we use:

$$q_1(x^n|x_i^{n-1}, y^m) = (1 - \epsilon)U\left(\hat{x}_i - \gamma_U Q^{1/2}\mathbf{1}, \hat{x}_i + \gamma_U Q^{1/2}\mathbf{1}\right) + \epsilon N\left(\hat{x}_i, \gamma_N^2 Q\right)$$



#### The deterministic move



Determine  $\alpha$  at crossing of line with target weight contour in:

$$\hat{x}_i = f(x_i^{n-1}) + \alpha_i K\left(y^n - H(f(x_i^{n-1}))\right)$$

with

$$K = QH^T(HQH^T + R)^{-1}$$

## The stochastic part of the proposal

A draw from the uniform density gives:

$$w_i = \frac{|Q|^{1/2} (2\gamma_U)^k}{1 - \epsilon} w_i^{rest} p(x_i^n | x_i^{n-1}) p(y^n | x_i^n)$$

A draw from the Gaussian density gives:

$$w_{i} = \frac{w_{i}^{rest} p(x_{i}^{n} | x_{i}^{n-1}) p(y^{n} | x_{i}^{n})}{\frac{\epsilon}{(2\pi)^{k/2} |\gamma_{N}^{2} Q|^{1/2}} \exp(-\frac{1}{2} \gamma_{U} d\beta_{i}^{n} (\gamma_{U}^{2} Q)^{-1} \gamma_{U} d\beta_{i}^{n})}$$

The ratio between the two is (ignoring the exp part):

$$\frac{(2\pi)^{k/2}|\gamma_N^2 Q|^{1/2}}{\epsilon} \frac{(1-\epsilon)}{|Q|^{1/2}(2\gamma_U)^k}$$

which can be made equal to one when:

$$\gamma_N = \frac{2^{k/2} \epsilon}{\pi^{k/2} (1 - \epsilon)} \gamma_U^k$$

## **Equivalent-Weights Particle Filter**

- Use relaxation-term proposal up to last time step
- Calculate  $w_i^{max}$  and target weight (e.g. 80%)
- Calculate deterministic moves for high-weight particles:

$$\hat{x}_i = f(x_i^{n-1}) + \alpha_i K\left(y^n - H(f(x_i^{n-1}))\right)$$

Determine stochastic move

$$p(\hat{\beta}_i^{n-1}) \propto (1-a)U[-b,b] + aN(0,\hat{Q})$$

Calculate new weights and resample 'lost' particles

# How essential are Gaussian assumptions?

- Allows for analytical expressions.
- But no real need.
- $w_i^{max}$  calculations do not have to be very accurate.
- Same for w<sup>target.</sup>
- Deterministic move has to be very accurate, good iterative schemes should be used.

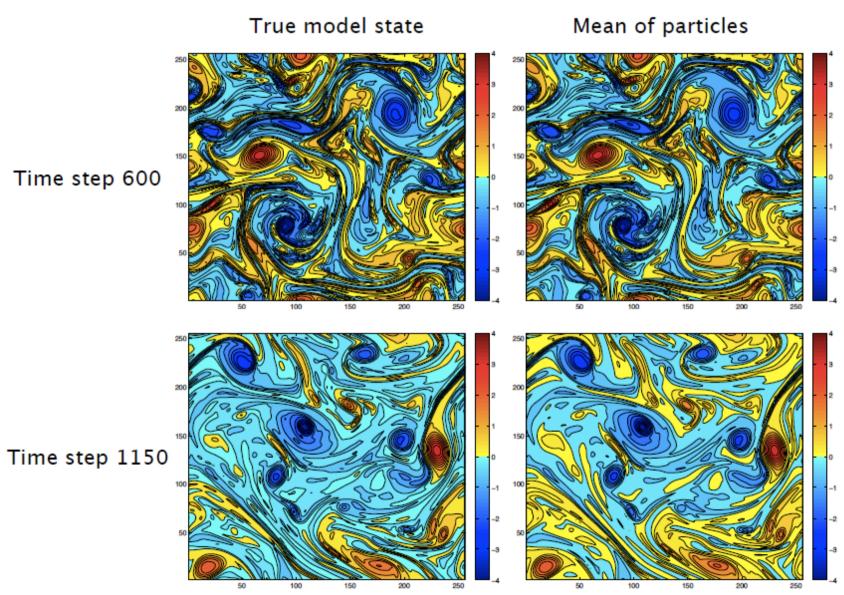
#### Application: the barotropic vorticity equation

Stochastic barotropic vorticity equation:

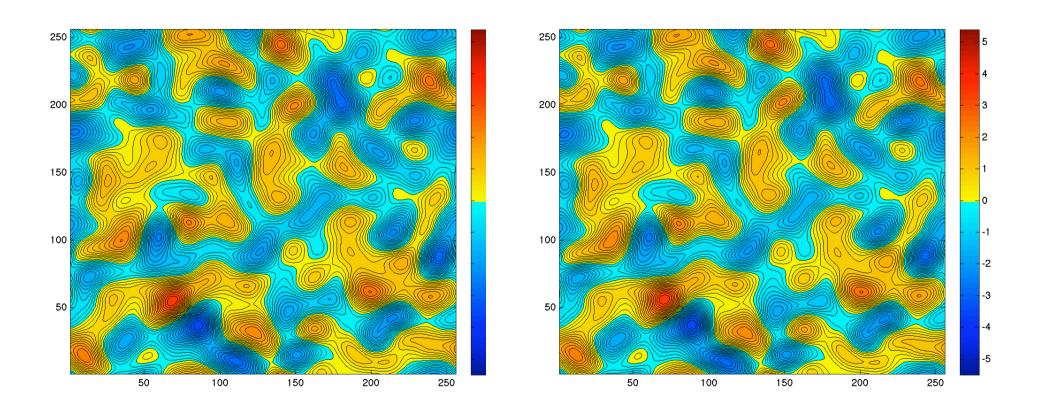
$$\frac{\partial q}{\partial t} + u \cdot \nabla q = F$$

- 256 by 256 grid 65,536 variables
- Double periodic boundary conditions
- Semi-Langrangian time stepping scheme
- Twin experiments
- Observations every 50 time steps decorrelation time of 42
- 32 particles

## Fully observed system



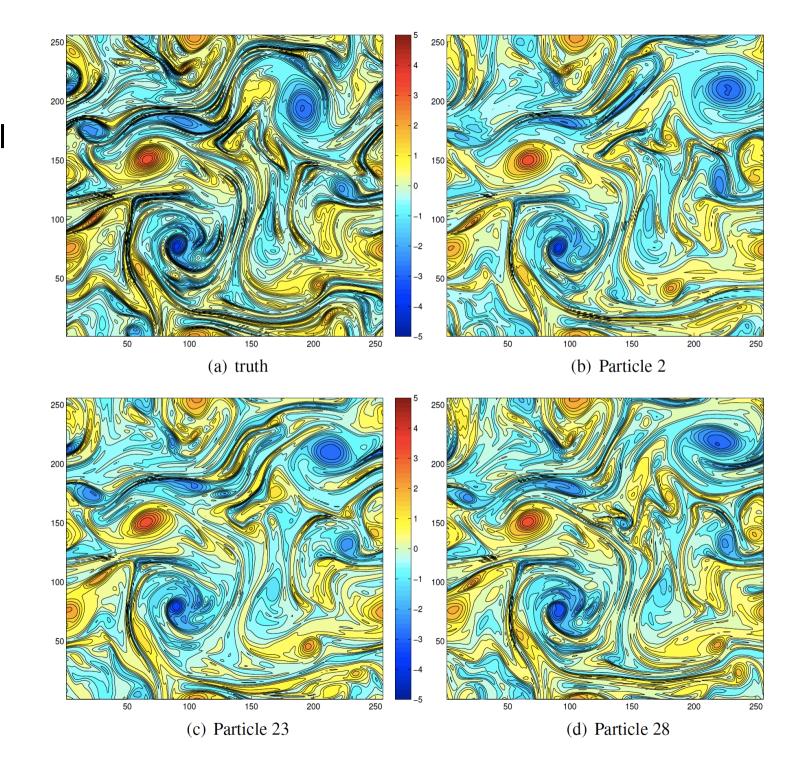
### 1/4 Observations over half of state



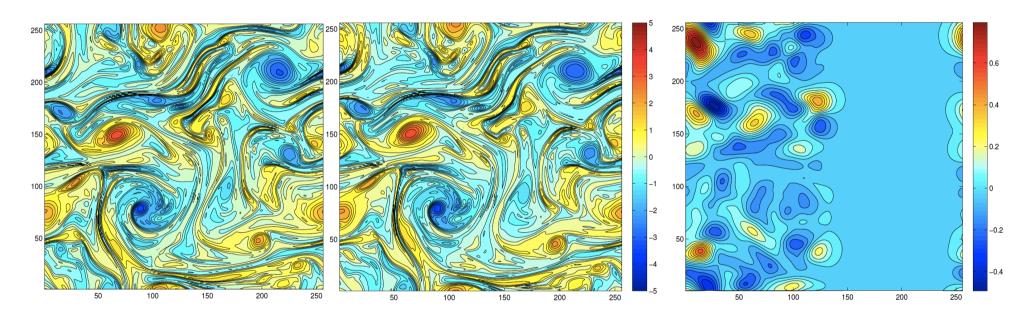
Truth

Mean of particle filter ensemble

Individual particles are not smooth.



## The update of the unobserved part

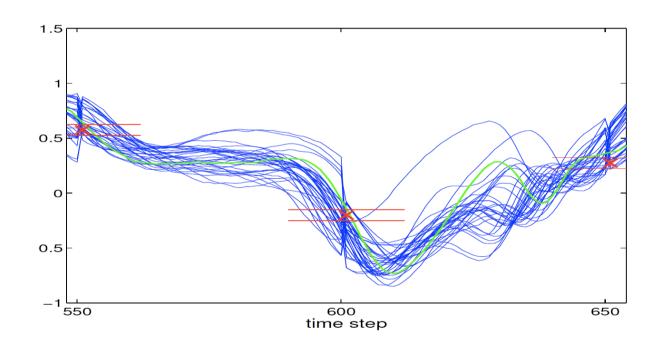


Particle 23 before update

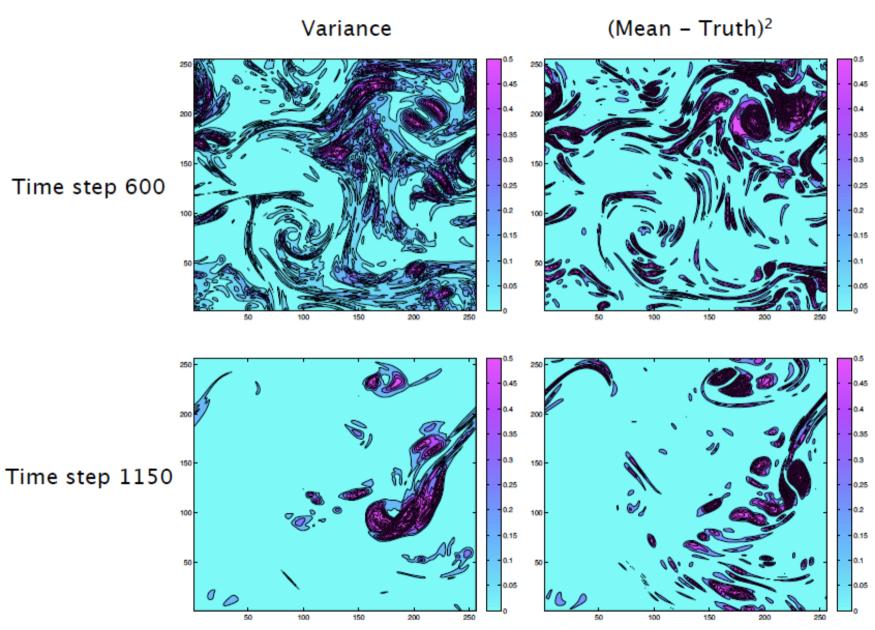
Particle 23 after update

Difference

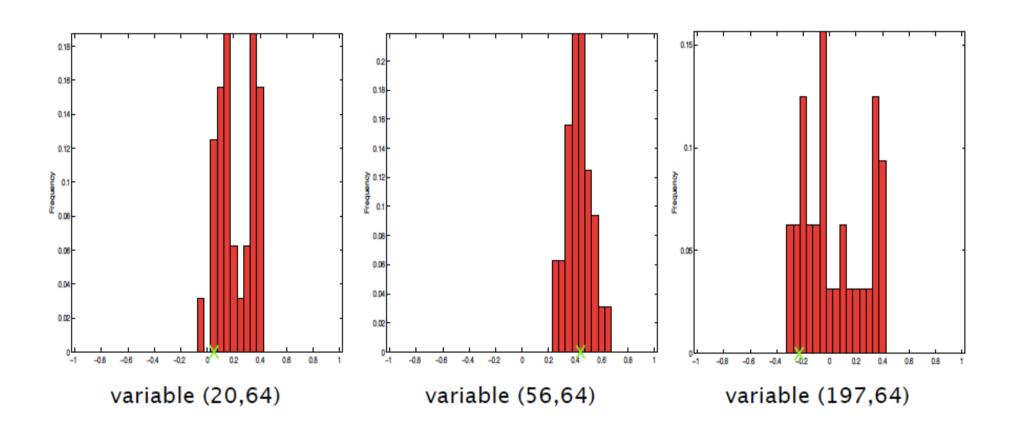
## Time evolution at a specific grid point



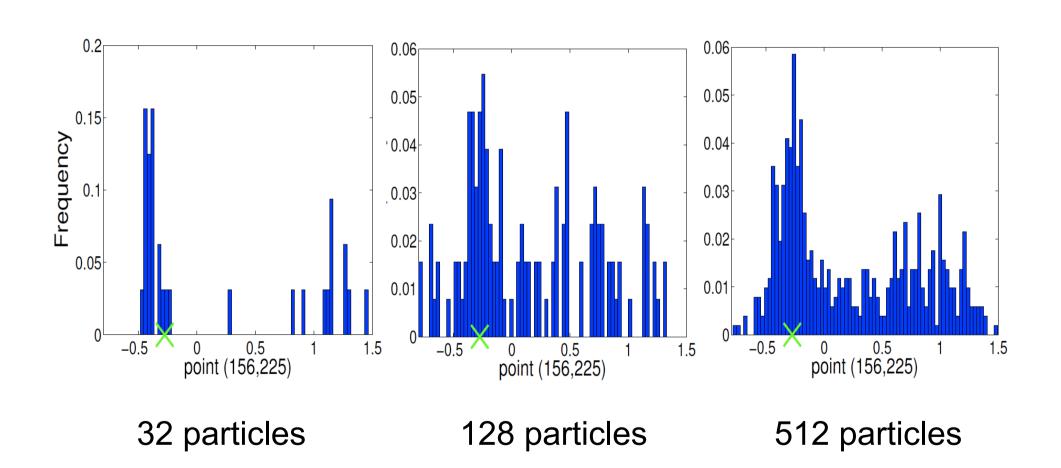
### ¼ observations over half of state



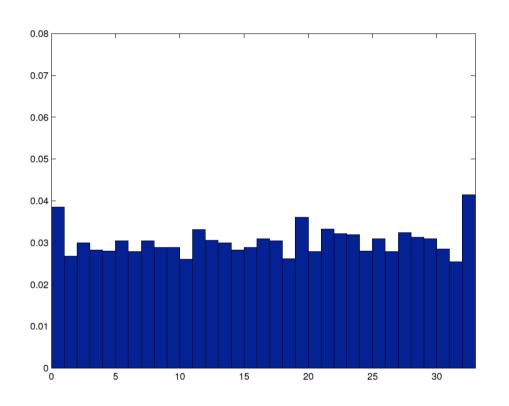
### Marginal posterior probability densities

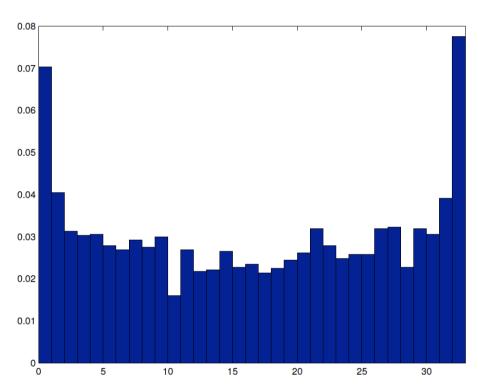


## Convergence of the pdf's



## Rank histograms

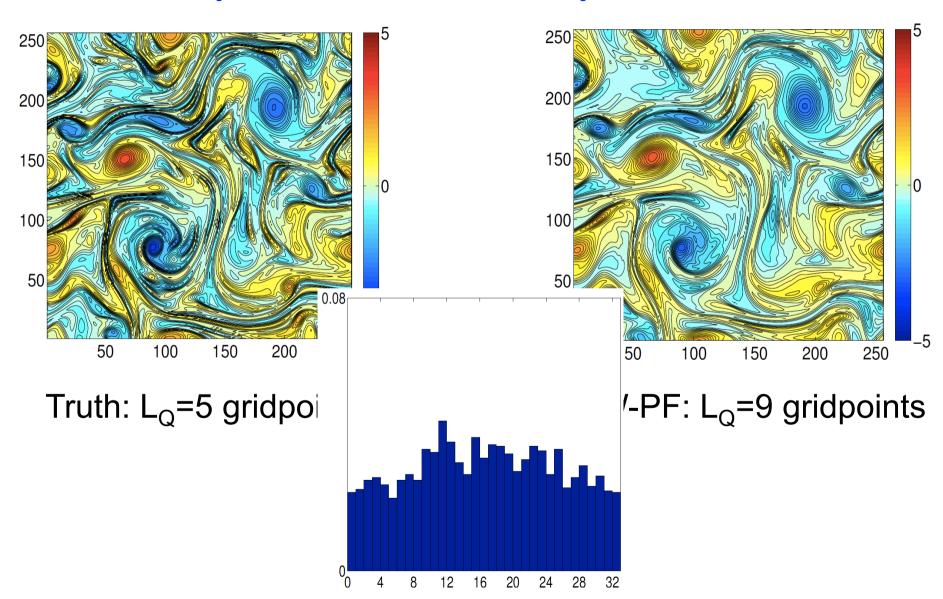




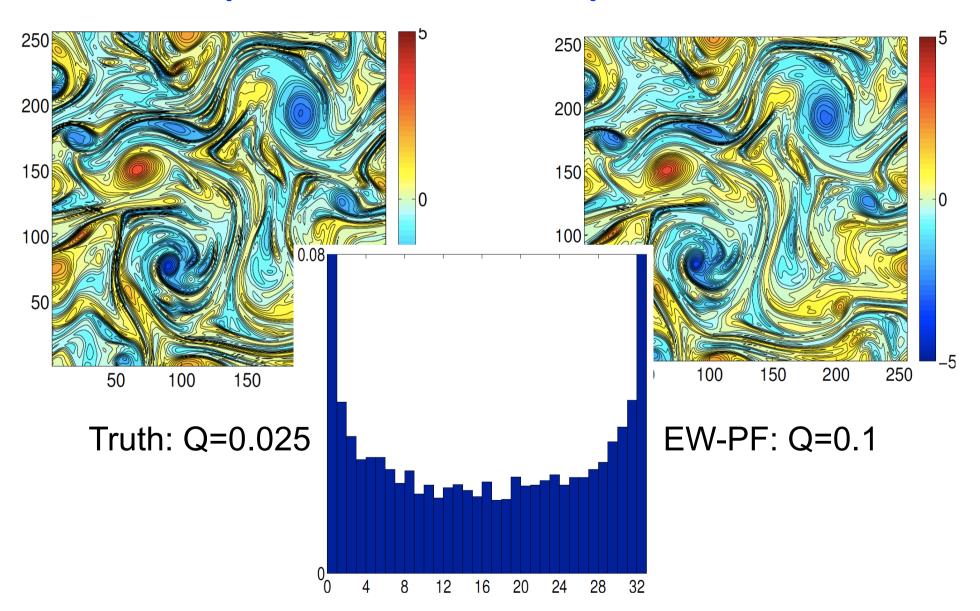
Full state observed

1/4 of half state observed

## Miss-specification of process noise



## Miss-specification of process noise



#### Conclusions

- Particle filters do not need state covariances.
- Degeneracy is related to number of observations, not to size of the state space.
- Proposal density allows enormous freedom
- Equivalent-weights scheme solves dimensionality problem?
- Other efficient schemes are being derived.
- Present work: numerical weather prediction, climate forecasting

### We need more people!

- The Data Assimilation group at reading consists of 30 scientists
- We still have room in the

Data Assimilation and Inverse Methods in Geosciences MSc programme

#### References

All references can be found on my website http://www.met.reading.ac.uk/~xv901096/research/publications.html

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