

Intervaller: Rolle til kritisk verdi X_1, X_2, \dots, X_n u.i.f $N(\mu, \sigma^2)$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i, \quad S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

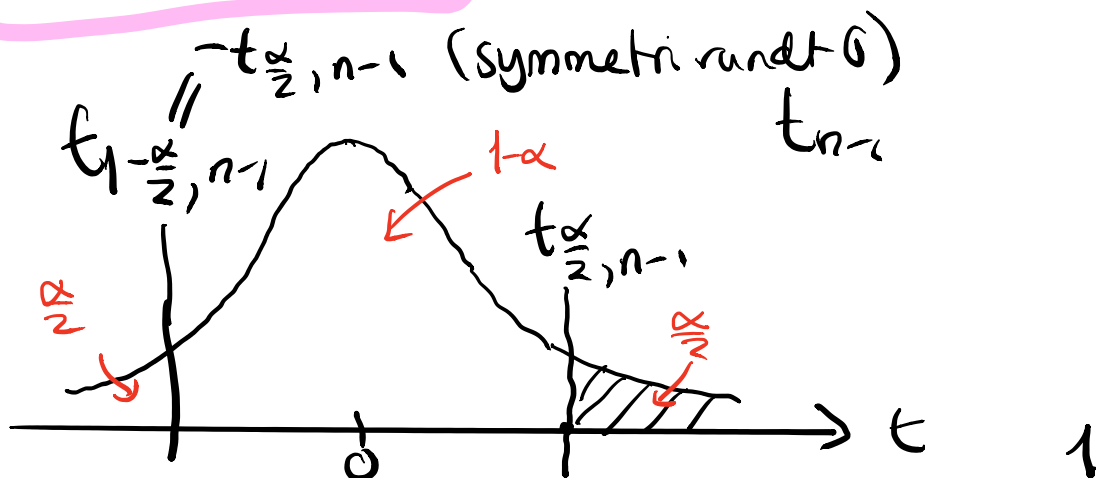
 $(1-\alpha) \cdot 100\%$ KI for μ , når σ ukjent

$$\left[\bar{X} \pm t_{\frac{\alpha}{2}, n-1} \frac{S}{\sqrt{n}} \right]$$

$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0,1)$$

$$\frac{S^2(n-1)}{\sigma^2} = V \sim \chi_{n-1}^2$$

$$\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{n-1}$$



$$P\left(T_{\frac{\alpha}{2}} > t_{\frac{\alpha}{2}, n-1}\right) = \frac{\alpha}{2}$$

$$(1-\alpha) \cdot 100\% = 95\% \Rightarrow \alpha = 0.05$$

$$\frac{\alpha}{2} = 0.025$$

$$Z_{0.025} = 1.96$$

$$t_{\frac{\alpha}{2}, 29} = 2.045$$
$$t_{0.025, 29}$$

$$(1-\alpha) \cdot 100\% = 90\% \Rightarrow \alpha = 0.1$$

$$\frac{\alpha}{2} = 0.05$$

$$Z_{0.05} = 1.645$$

$$t_{\frac{\alpha}{2}, 29} = 1.699$$
$$t_{0.05, 29}$$

$$(1-\alpha) \cdot 100\% = 99\% \Rightarrow \alpha = 0.01$$

$$\frac{\alpha}{2} = 0.005$$

$$Z_{0.005} = 2.576$$

$$t_{\frac{\alpha}{2}, 29} = 2.756$$
$$t_{0.005, 29}$$

NB: $P\left(T_{\frac{\alpha}{2}} > t_{\text{tail}}\right) = \text{tail}$ i tabellene

Prediksjonsintervall [9.6] (PI)

X_1, X_2, \dots, X_n u.i.f $N(\mu, \sigma^2)$
 X_0 uavhengig av X_1, \dots, X_n $\uparrow \sigma^2$ ukjent
 \uparrow ny observasjon $X_0 \sim N(\mu, \sigma^2)$

KI: Intervall vi her stor tillit til at
sann μ ligger. \uparrow 95%

PI: Intervall der vi her stor tillit til at
fremtidig observasjon X_0 ligger.

Utlede PI (σ^2 ukjent):

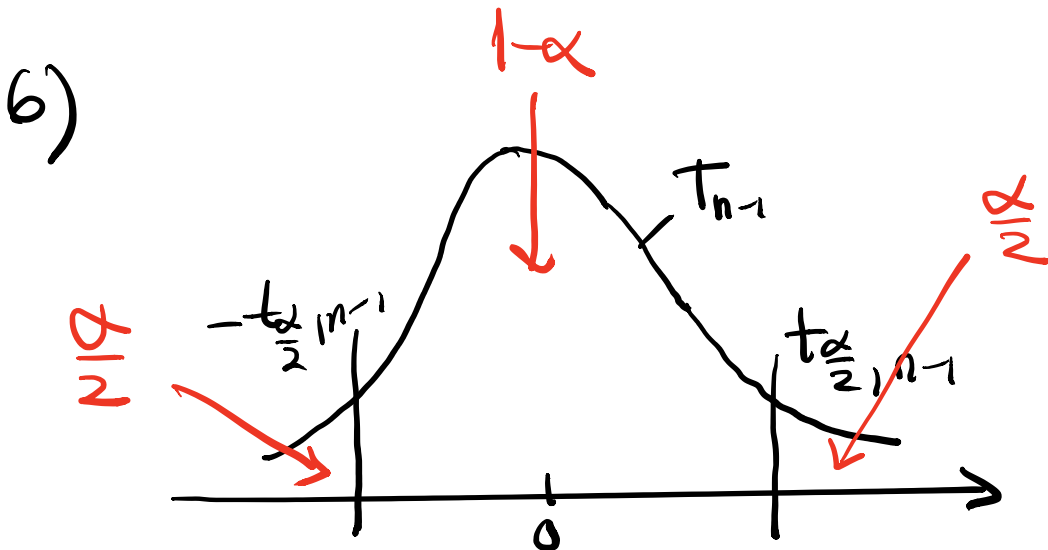
Tricks: ser på fordelingen til $(X_0 - \bar{X})$

- 1) $(X_0 - \bar{X})$ er normal fordelt
- 2) $E(X_0 - \bar{X}) = E(X_0) - E(\bar{X})$
 $= \mu - \mu = 0$

$$\begin{aligned}
 3) \quad \text{Var}(X_0 - \bar{X}) &= \text{Var}(X_0) + \underbrace{(-1)^2}_{\text{varh.}} \text{Var}(\bar{X}) \\
 &= \underbrace{\text{Var}(X_0)}_{\sigma^2} + \underbrace{\text{Var}(\bar{X})}_{\frac{\sigma^2}{n}} = \sigma^2 \left(1 + \frac{1}{n}\right)
 \end{aligned}$$

$$\begin{aligned}
 4) \quad Z &= \frac{X_0 - \bar{X} - E(X_0 - \bar{X})}{\text{SD}(X_0 - \bar{X})} = \frac{X_0 - \bar{X}}{\sigma \sqrt{1 + \frac{1}{n}}} \\
 &\sim N(0, 1)
 \end{aligned}$$

$$5) \quad T = \frac{X_0 - \bar{X}}{\sigma \sqrt{1 + \frac{1}{n}}} \sim t_{n-1}$$



$$P(-t_{\frac{\alpha}{2}, n-1} < T < t_{\frac{\alpha}{2}, n-1}) = 1-\alpha$$

$$7) \quad \frac{\mu_0 - \bar{X}}{\sqrt{1+\frac{1}{n}} s}$$

Løse ut så μ_0 i niaten

- $(\sqrt{1+\frac{1}{n}} s)$

- $+ \bar{X}$

$$P(\underbrace{\bar{X} - t_{\frac{\alpha}{2}, n-1} \sqrt{1+\frac{1}{n}} s}_{X_L} < \mu_0 < \underbrace{\bar{X} + t_{\frac{\alpha}{2}, n-1} \sqrt{1+\frac{1}{n}} s}_{X_U}) = 1-\alpha$$

Vi setter inn numeriske verdier $(\bar{X}, s, n, t_{\frac{\alpha}{2}, n-1})$

og får et $[X_L, X_U]$ $(1-\alpha) 100\%$

PI for en fremtidig μ_0 .

EKS: Høyde kvinner 95% PI

$$\left[\bar{x} \pm t_{\frac{\alpha}{2}, n-1} \sqrt{1 + \frac{1}{n}} s \right]$$

↑
169.5

$t_{0.025, 148}$
1.976

$\sqrt{1 + \frac{1}{149}}$

6.6

$$= [156.5, 182.5]$$

Husk 95% KI [168.9, 170.6]

⇒ Et PI er mye bredere enn et KI

$$PI: \sqrt{1 + \frac{1}{n}} \gg \sqrt{\frac{1}{n}} : KI$$

To utvalg [9.8, 8.4]

Notasjon:

\bar{X}_{1i} (els: bensinforbruk biltype 1) $i=1, \dots, n_1$ u.i.f $\bar{X}_{1i} \sim N(\mu_1, \sigma_1^2 = 2)$

\bar{X}_{2j} (— " — biltype 2) $j=1, \dots, n_2$ u.i.f $\bar{X}_{2j} \sim N(\mu_2, \sigma_2^2 = 3)$

\bar{X}_{1i} og \bar{X}_{2j} uavh. avhengende

n_1 n_2 10 10

Oppdrag: Estimator & KI for $\mu_1 - \mu_2$

1) Estimator for $\mu_1 - \mu_2$:

$$\bar{\bar{X}}_1 - \bar{\bar{X}}_2$$

\uparrow \uparrow

$$\frac{1}{n_1} \sum_i \bar{X}_{1i} \quad \frac{1}{n_2} \sum_j \bar{X}_{2j}$$

2) Egenskaper:

$\bar{X}_1 - \bar{X}_2$ er normal fordelt
fordi differansen av to
uavh. N-fordelte SV

$$\begin{aligned} E(\bar{X}_1 - \bar{X}_2) &= E(\bar{X}_1) - E(\bar{X}_2) \\ &= \mu_1 - \mu_2 \end{aligned}$$

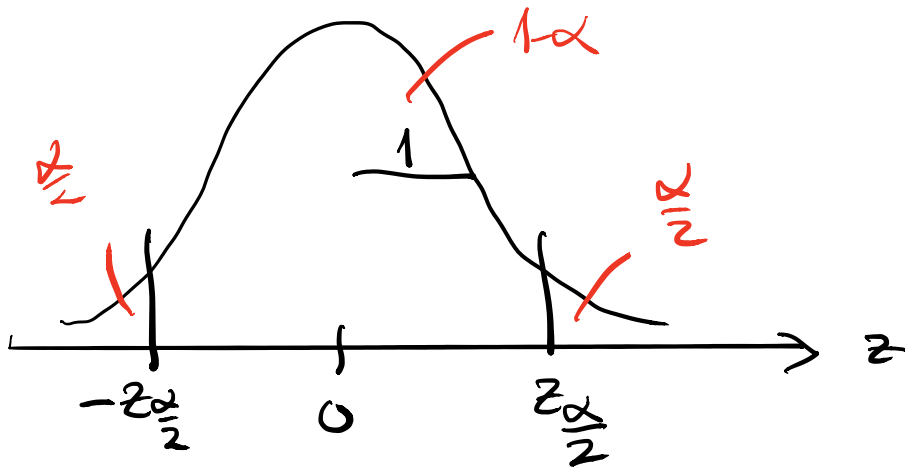
$$\text{Var}(\bar{X}_1 - \bar{X}_2) = \text{Var}(\bar{X}_1) + (-1)^2 \text{Var}(\bar{X}_2)$$

↑
uavhengige

$$= \underbrace{\text{Var}(\bar{X}_1)}_{\frac{\sigma_1^2}{n_1}} + \underbrace{\text{Var}(\bar{X}_2)}_{\frac{\sigma_2^2}{n_2}}$$

$$Z = \frac{\bar{X}_1 - \bar{X}_2 - E(\bar{X}_1 - \bar{X}_2)}{\text{SD}(\bar{X}_1 - \bar{X}_2)}$$

$$= \frac{\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$



$$P(-z_{\frac{\alpha}{2}} < \underset{\uparrow}{Z} < z_{\frac{\alpha}{2}}) = 1 - \alpha$$

$$\frac{\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

og løs ut så $(\mu_1 - \mu_2)$ i midten

$$(1-\alpha) 100\% \text{ KI: } \left[(\bar{X}_1 - \bar{X}_2) \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \right]$$

Hva hvis σ_1 og σ_2 er ukjent?

$$S_1^2 = \frac{1}{n_1 - 1} \sum_{i=1}^{n_1} (X_{1i} - \bar{X}_1)^2$$

$$S_2^2 = \frac{1}{n_2 - 1} \sum_{j=1}^{n_2} (X_{2j} - \bar{X}_2)^2$$

$$\frac{\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}} \approx t_{\nu}$$

funksjon av
 S_1, S_2, n_1, n_2

$$KI: \left[\bar{X}_1 - \bar{X}_2 \pm t_{\frac{\alpha}{2}, \nu} \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}} \right]$$

ν blir oppgitt på eksempl.

(Innlevering 3: 1c morsvin
→ der må du regne ut ν selv!)