

TMA4250 Spatial Statistics

Assignment 3: Markov Random Fields

IMF/NTNU

March 2017

Introduction

This assignment contains problems related to Markov Random Fields (MRF). The R-package should be used in solving the problems and relevant functions can be found in the R library `spatial`.

Problem 1: Markov RF

This problem is based on observations of seismic data over a domain $\mathcal{D}_s \subset \mathbb{R}^2$. The objective is to identify the underlying lithology (`{sand, shale}`) distribution over \mathcal{D}_s .

The data are collected on a regular (75×75) grid $\mathcal{L}_{\mathcal{D}_s}$, and the seismic data are denoted $Z(\mathbf{s}) : \{Z(\mathbf{s}_i); \mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}\}$. The data are available in the R library MASS in the file `seismic.dat`.

Moreover, observations of the lithology distribution (`{sand, shale}`) in a geologically comparable domain $\mathcal{D}_c \subset \mathbb{R}^2$ are available. The lithology distribution is collected on a regular (66×66) grid $\mathcal{L}_{\mathcal{D}_c}$, with the same spacing as $\mathcal{L}_{\mathcal{D}_s}$, over \mathcal{D}_c . The observations with code 0 for sand and 1 for shale are available in the R library MASS in the file `complit.dat`.

Assume that the underlying lithology surface over \mathcal{D}_s can be represented by $\{Y(\mathbf{s}); \mathbf{s} \in \mathcal{D}_s \subset \mathbb{R}^2\}$ discretized into a lattice $\{Y(\mathbf{s}_i) : \mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}\}$ with $Y(\mathbf{s}_i) \in \{0, 1\}$ representing sand and shale respectively.

The seismic data stored in `seismic.dat` have uncertainty defined by:

$$Z(\mathbf{s}_i)|Y(\mathbf{s}_i) = \begin{cases} 0.02 + \epsilon(\mathbf{s}_i) & \text{if } y(\mathbf{s}_i) = 0 \\ 0.08 + \epsilon(\mathbf{s}_i) & \text{if } y(\mathbf{s}_i) = 1 \end{cases} \quad ; \mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}$$

where $\epsilon(\mathbf{s}_i)$ iid Gauss(0, 0.06²).

a) Plot the two datasets *seismic.dat* and *complit.dat*.

b) Consider a uniform prior model on $Y(\mathbf{s})$, i.e.

$$[Y(\mathbf{s}_i)] = Unif\{0, 1\} \quad \text{for } \mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}.$$

Specify the data model and the process model for the lithology ($\{\text{sand, shale}\}$) at an arbitrary position $\mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}$.

Find the posterior distribution $[Y(\mathbf{s}_i)|Z(\mathbf{s}_i)]$ for $Y(\mathbf{s}_i) \in \{0, 1\}$. Simulate 10 realisations of this distribution on $\mathcal{L}_{\mathcal{D}_s}$ and visualize the realisations.

Derive expressions for the posterior mean $E\{Y(\mathbf{s}_i)|Z(\mathbf{s}_i)\}$ and the posterior variance $\text{Var}\{Y(\mathbf{s}_i)|Z(\mathbf{s}_i)\}$ and visualize the results. Compare these to the 10 realisations you made earlier.

Finally, use Maximum a posteriori estimation (MAP) to estimate the lithology distribution $Y(\mathbf{s}_i)$ over $\mathcal{L}_{\mathcal{D}_s}$. Define the MAP-estimator for $Y(\mathbf{s}_i)$, visualize the results and comment.

c) The Hammersley-Clifford Theorem and Brook's lemma allow us to express a MRF either through one multivariate pdf with respect to cliques (Gibbs formulation) or as a set of univariate conditional local pdf's with respect to neighborhoods (Markov formulation). The Gibbs formulation and the Markov formulation are equivalent.

Now, consider a new prior model for $Y(\mathbf{s})$ with neighborhood system denoted by $\mathcal{N}(\cdot) : \{\mathcal{N}(\mathbf{s}_k); \mathbf{s}_k \in \mathcal{L}_{\mathcal{D}_s}\}$ consisting of the four closest neighbors:

$$[\mathbf{Y}; \beta] = \text{const}_G \times \exp\left\{\beta \sum_{\mathbf{s}_k \in \mathcal{L}_{\mathcal{D}_s}} \sum_{\mathbf{s}_j \in \mathcal{N}(\mathbf{s}_k)} I(y(\mathbf{s}_k) = y(\mathbf{s}_j))\right\}, \quad (1)$$

where $\mathbf{Y} = (Y(\mathbf{s}_1), \dots, Y(\mathbf{s}_n))'$ and n is the number of grid nodes in $\mathcal{L}_{\mathcal{D}_s}$. The function $I(A)$ is an indicator function taking the value 1 if A is true and 0 otherwise. The above prior defines the Gibbs formulation of an Ising model.

Specify the Markov formulation corresponding to the Gibbs formulation in (1), i.e. derive an expression for the discrete distribution $[Y(\mathbf{s}_i)|\mathbf{Y}_{-i}; \beta]$ for each $\mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}$ where $\mathbf{Y}_{-i} = (Y(\mathbf{s}_1), \dots, Y(\mathbf{s}_{i-1}), Y(\mathbf{s}_{i+1}), \dots, Y(\mathbf{s}_n))'$.

Derive expressions for the posterior distributions $[\mathbf{Y}|\mathbf{Z}; \beta]$ and $[Y(\mathbf{s}_i)|\mathbf{Y}_{-i}, \mathbf{Z}; \beta]$ for all $\mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}$. Here, $\mathbf{Z} = (Z(\mathbf{s}_1), \dots, Z(\mathbf{s}_n))'$, a vector containing the seismic observations for all $\mathbf{s}_i \in \mathcal{L}_{\mathcal{D}_s}$.

We are interested in realizations from $[\mathbf{Y}|\mathbf{Z}; \beta]$ and estimates of $E\{\mathbf{Y}|\mathbf{Z}; \beta\}$, $\text{Var}\{\mathbf{Y}|\mathbf{Z}; \beta\}$ and $\text{MAP}\{\mathbf{Y}|\mathbf{Z}; \beta\}$. To obtain this, we need to estimate $\hat{\beta}$. Use

the observations from the geologically comparable domain \mathcal{D}_c to estimate β by a maximum pseudo-likelihood procedure. Denote the estimate $\hat{\beta}$.

Set the parameter $\beta = \hat{\beta}$ and use a Gibbs sampler MCMC-algorithm to simulate from the posterior $[\mathbf{Y}|\mathbf{Z}; \hat{\beta}]$. You can for example use a single-site updating MCMC-algorithm. Use torus boundary conditions to avoid border problems. Document that the algorithm has converged. Include visualizations of 10 independent realizations in your report.

Estimate $E\{\mathbf{Y}|\mathbf{Z}; \beta\}$, $\text{Var}\{\mathbf{Y}|\mathbf{Z}; \beta\}$ and $\text{MAP}\{\mathbf{Y}|\mathbf{Z}; \beta\}$ based on several realizations from $[\mathbf{Y}|\mathbf{Z}; \hat{\beta}]$, and visualize the results.

Comment on the results.

d) Compare the results in **b)** and **c)** and comment.