

LØSNINGSFORSLAG

SIF5064 ROMLIG STATISTIK

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OPPGAVE 1 KONTINUERLIGE FEJL

$$\{R(x); x \in \mathbb{R}^n\} \rightarrow \begin{aligned} E\{R(x)\} &= \mu \\ \text{Var}\{R(x) - R(x+h)\} &= 2 \cdot \gamma(h) \end{aligned}$$

a) Annen ordens stasjonært

$$\Downarrow \\ E\{R(x)\} = \mu \quad \text{uavh. av } x$$

$$\text{Var}\{R(x)\} = \sigma^2 < \infty$$

$$\text{Cov}\{R(x), R(x+h)\} = G(h) \quad \text{uavh. av } x$$

herav

↑ pos. def. funksjon

$$\gamma(h) = \frac{1}{2} \text{Var}\{R(x) - R(x+h)\}$$

$$= \frac{1}{2} [\text{Var}\{R(x)\} + \text{Var}\{R(x+h)\} - 2 \text{Cov}\{R(x), R(x+h)\}]$$

$$= \sigma^2 - G(h)$$

b) IRF-1

$$[R(x) - R(x+h)] \quad \text{annen ordens stasjonært} \\ \neq x, \neq h$$

$$\Downarrow \\ E\{R(x)\} = \mu \quad \text{uavh. av } x$$

$$\text{Var}\{R(x) - R(x+h)\} = 2 \cdot \gamma(h) < \infty$$

$[-\gamma(h)]$ må være betinget pos. def. funksjon:

$$-\sum_{i=1}^n \sum_{j=1}^n \alpha_i \alpha_j \gamma(x_i - x_j) \geq 0$$

for $\forall \alpha = (\alpha_1, \dots, \alpha_n)$ slik at $\sum_{i=1}^n \alpha_i = 0$
 \forall konfigurasjoner (x_1, \dots, x_n)
 $\forall n$

I prediksjon ønsker en $\hat{\alpha}$ predikter $Z(x_0)$ basert på observasjonene $\{Z(x_1), \dots, Z(x_n)\}$. Lineær, forventningsrett minste kvadraters estimator er:

$$\hat{Z}(x_0) = \sum_{i=1}^n \hat{\alpha}_i Z(x_i)$$

hvor

$$\hat{\alpha} = \underset{\alpha}{\operatorname{argmin}} \left\{ \operatorname{Var} \left\{ Z(x_0) - \sum_{i=1}^n \alpha_i Z(x_i) \right\} \right\}$$

$$E \left\{ Z(x_0) - \sum_{i=1}^n \alpha_i Z(x_i) \right\} = 0$$

herav

$$\operatorname{Prediksjonsvarians}_n \operatorname{Var} \left\{ \sum_{i=0}^n \beta_i Z(x_i) \right\} \rightarrow \begin{cases} \beta_0 = 1 \\ \beta_i = -\alpha_i \quad i=1, 2, \dots, n \end{cases}$$

$$\text{hvor } \sum_{i=0}^n \beta_i = 0.$$

Prediksjonsvariansen omhandler keen lineær-kombinasjoner hvor $\sum_{i=0}^n \beta_i = 0$, dvs kontraster.

9) IRF-1 med $\gamma(h) = \tau/|h|^\nu$ $\tau > 0, 0 \leq \nu < 2$

'Affine similar'

↓

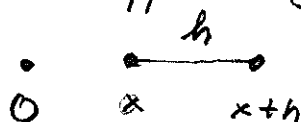
$$\{R(x); x \in \mathbb{R}^n\} \rightarrow \begin{aligned} E\{R(x)\} &= \mu \\ \text{Var}\{R(x) - R(x+h)\} &= 2 \cdot \gamma(h) \end{aligned}$$

$$\{R(\omega_3 x); x \in \mathbb{R}^n\} \rightarrow \begin{aligned} E\{R(x)\} &= \mu \\ \text{Var}\{R(x) - R(x+h)\} &= 2 \cdot \omega_3 \cdot \gamma(h) \end{aligned}$$

↑
avh. av $h!$

dvs reskalering av referanseområdet endrer kun skala, men ikke variogramstruktur i feltet.

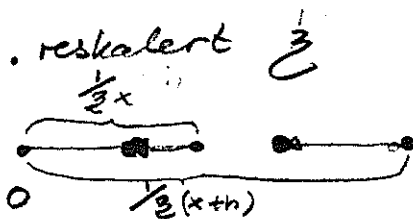
Betrakt
• opprinnelig



$$\text{Var}\{R(x) - R(x+h)\}$$

$$= 2 \cdot \gamma(h)$$

$$= 2 \cdot \tau/|h|^\nu$$



$$\text{Var}\{R(\frac{1}{3}x) - R(\frac{1}{3}(x+h))\}$$

$$= 2 \cdot \gamma(\frac{1}{3}h)$$

$$= 2 \cdot \tau/|\frac{1}{3}h|^\nu$$

$$= 2 \tau \cdot \underbrace{\frac{1}{3^\nu}}_{\omega_3} |h|^\nu$$

OPPGAVE 2 HENDERSES FELT

Stasjonært punktfelt:

$$\lambda = \frac{E\{N(\mathcal{D})\}}{|\mathcal{D}|}$$

$$K(r) = E\{N(rB_0)\} / \lambda$$

$$\{x_1, \dots, x_n \in \mathcal{B}\}$$

a) Poisson punktfelt.

$$\hat{\lambda} = \frac{N}{|\mathcal{B}|} \rightarrow E\{\hat{\lambda}\} = \frac{E\{N\}}{|\mathcal{B}|} = \frac{\lambda |\mathcal{B}|}{|\mathcal{B}|} = \lambda - \text{forv. rett}$$

$$\text{Var}\{\hat{\lambda}\} = \frac{\text{Var}\{N\}}{|\mathcal{B}|^2} = \frac{\lambda |\mathcal{B}|}{|\mathcal{B}|^2} = \frac{\lambda}{|\mathcal{B}|}$$

$$\hat{K}(r) = \pi r^2 \rightarrow K(r) \xrightarrow{\text{pr. def}} \frac{\lambda \pi r^2}{\lambda} = \pi r^2$$

Forventningsrett / Varians 0!!

b) Vilkarlig punktfelt.

$$\hat{\lambda} = \frac{N}{|\mathcal{B}|} \rightarrow E\{\hat{\lambda}\} = \frac{E\{N\}}{|\mathcal{B}|} = \frac{\lambda |\mathcal{B}|}{|\mathcal{B}|} = \lambda - \text{forv. rett}$$

$$\text{Var}\{\hat{\lambda}\} = \frac{\text{Var}\{N\}}{|\mathcal{B}|^2}$$

$$= \frac{1}{|\mathcal{B}|^2} \sum_i i^2 \text{Prob}\{N(\mathcal{B})=i\} - \lambda^2$$

$$\hat{K}(r) = \frac{1}{\hat{\lambda}^2} \cdot \frac{1}{n} \sum_{i=1}^n \left[N(b(x_i, r)) - 1 \right] - \text{Forv. skjev} \sim \frac{1}{n}$$

Rel. høy varians
- rand. probl.


bedre (dvs. mindre varians) estimator:

$$\hat{K}(r) = \frac{1}{\hat{\lambda}^2} \cdot \sum_{\substack{i=1 \\ i \neq j}}^n \sum_{j=1}^n \frac{I(|x_i - x_j| < r)}{|B_{x_i} \cap B_{x_j}|} \leftarrow I(A) = \begin{cases} 1 & \text{hvis } A \text{ sann} \\ 0 & \text{ellers} \end{cases}$$

B_{x_i} er B sentrert i x_i

OPPGAVE 3 MOSAIKK FELT

$$L: \{L_x; x \in \mathcal{L}_D\} ; L_x \in \{-1, 1\}$$

Stationært Ising-felt -  N .

a) Markov formulering:

$$\text{Prob}\{L_x = l_x | L_y = l_y; y \in \mathcal{L}_D; y \neq x\}$$

$$= \text{konst} \times \exp\left\{ \beta_H \sum_{j \in N_x} l_x l_j \right\}$$

Naboskap

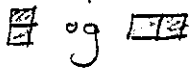


$$= \text{Prob}\{L_x = l_x | L_y = l_y; y \in N_x\}$$

Gibbs formulering:

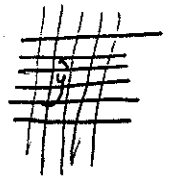
$$\text{Prob}\{L = l\} = \text{konst} \times \exp\left\{ \beta_G \sum_{i,j} l_i l_j \right\}$$

$\beta_H = \beta_G$

'cliques'
+


b) Ta utgangspunkt i M-H-algoritmen.

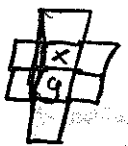
$$q_{ij} = \text{Prob}\{L_x = l_x^j, L_y = l_y^j | L_z = l_z^i; z \in \mathcal{L}_D^{-xy}\}$$



$$= \frac{\text{Prob}\{L_x = l_x^j, L_y = l_y^j, (L_z = l_z^i; z \in \mathcal{L}_D^{-xy})\}}{\sum_{x,y} \text{Prob}\{L_x = l_x^j, L_y = l_y^j, (L_z = l_z^i; z \in \mathcal{L}_D^{-xy})\}}$$

$$= \frac{\exp\left\{ \beta \sum_{x \sim k} l_x l_k + \sum_{\substack{y \sim k \\ y \neq x}} l_y l_k + \sum_{\substack{k \sim l \\ k, l \in N_{xy}}} l_k l_l \right\}}{\sum_{x,y} \exp\left\{ \dots \right\}}$$

$$= \frac{\sum_{x,y} \exp\left\{ \beta \sum_{x \sim k} l_x l_k + \sum_{\substack{y \sim k \\ y \neq x}} l_y l_k \right\}}{\sum_{xy} \exp\left\{ \beta \sum_{x \sim k} l_x l_k + \sum_{\substack{y \sim k \\ y \neq x}} l_y l_k \right\}}$$



$$= \frac{\sum_{xy} \exp\left\{ \beta \sum_{x \sim k} l_x l_k + \sum_{\substack{y \sim k \\ y \neq x}} l_y l_k \right\}}{\sum_{xy} \exp\left\{ \dots \right\}}$$

kan regnes ut ved $(1, -1) (-1, 1) (1, 1)$ gitt $l_z^i; z \in \mathcal{L}_D^{-xy}$.

Akseptanssynligheten er:

$$\alpha_{ij} = \min \left\{ 1, \frac{\text{Prob}\{L=l^j\}}{\text{Prob}\{L=l^i\}} \cdot \frac{q_{ji}}{q_{ij}} \right\}$$

MERK: $\{l_z = l_z^j = l_z^i; z \in \mathcal{L}_D^{-xy}\}$

$$\alpha_{ij} = \min \left\{ 1, \frac{\exp\{A_{xy}^{ij}\}}{\sum_{xy} \exp\{A_{xy}^{ij}\}} \cdot \left[\frac{\exp\{A_{xy}^{ji}\}}{\sum_{xy} \exp\{A_{xy}^{ji}\}} \right]^{-1} \cdot \frac{q_{ji}}{q_{ij}} \right\}$$

\uparrow
 $\{l_z; z \in \mathcal{L}_D^{-xy}\}$

$$= \min \left\{ 1, q_{ij} \cdot [q_{ji}]^{-1} \cdot \frac{q_{ji}}{q_{ij}} \right\} = \underline{\underline{1}}$$

Algoritme:

Initiate

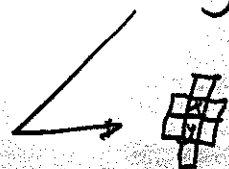
$$L = l^0$$

Itererer $i=1, \dots$

$$L^i = (l_x, l_y, (l_z; z \in \mathcal{L}_D^{-xy}))$$

$$\leadsto \text{Prob}\{L_x=l_x, L_y=l_y \mid L_z=l_z^{i-1}; z \in \mathcal{L}_D^{-xy}\}$$

$$= \text{Prob}\{L_x=l_x, L_y=l_y \mid L_z=l_z^{i-1}; z \in \mathcal{N}_{xy}\}$$



se foran