

SOLUTION / TMA 4250 SPATIAL STATISTICS

Exam May 22, 2006

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PROBLEM 1 Continuous Fields

consider GRF $\{R(x); x \in \mathbb{R}^1\}$:

$$E\{R(x)\} = 0$$

$$\text{Cov}\{R(x'), R(x'')\} = \begin{cases} 1 & \text{for } x' = x'' \\ 0 & \text{else} \end{cases}$$

a) Kriging in $R(0)$ given $R(-1) = r_1$ and $R(1) = r_1$.

Kriging is Best Linear Unbiased Predictor under Squared Error Loss;

$$R^*(0) = \alpha_0 + \alpha_1 R(-1) + \alpha_2 R(1)$$

unbiased:

$$\begin{aligned} E\{R^*(0)\} &= \alpha_0 + \alpha_1 E\{R(-1)\} + \alpha_2 E\{R(1)\} \\ &= \alpha_0 [= 0] \end{aligned}$$

hence

$$\alpha_0 = 0$$

Prediction variance

$$\begin{aligned} \text{Var}\{R^*(0) - R(0)\} &= \alpha_1^2 \overset{=1}{\text{Var}\{R(-1)\}} + 2\alpha_1\alpha_2 \overset{=0}{\text{Cov}\{R(-1), R(1)\}} + \alpha_2^2 \overset{=1}{\text{Var}\{R(1)\}} \\ &\quad - 2\alpha_1 \overset{=0}{\text{Cov}\{R(-1), R(0)\}} - 2\alpha_2 \overset{=0}{\text{Cov}\{R(1), R(0)\}} + \overset{=1}{\text{Var}\{R(0)\}} \\ &= \alpha_1^2 + \alpha_2^2 + 1 \end{aligned}$$

Min. Variance for $\alpha_1 = \alpha_2 = 0$, i.e.

$$R^*(0) = 0$$

$$\text{Var}\{R^*(0) - R(0)\} = 1$$

which is not surprising since $R(0)$, $R(-1)$ and $R(1)$ are mutually independent.

$$S_d(x) = \int_{x-d/2}^{x+d/2} R(y) dy$$

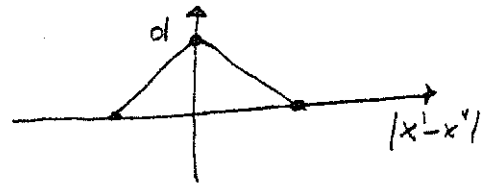
b) $S_d(x)$ is Gaussian since it is a linear combination of $R(x)$ which is Gaussian.

$$E\{S_d(x)\} = \int_{x-d/2}^{x+d/2} E\{R(y)\} dy = 0$$

$$\text{Cov}\{S_d(x'), S_d(x'')\} = \int_{x'-d/2}^{x'+d/2} \int_{x''-d/2}^{x''+d/2} \text{Cov}\{R(y), R(z)\} dy dz$$

$$= \left[\text{overlap of intervals } [x'-d/2, x'+d/2] \text{ and } [x''-d/2, x''+d/2] \right]$$

$$= \begin{cases} d - |x' - x''| & \text{for } |x' - x''| < d \\ 0 & \text{else} \end{cases}$$



c) Kriging:

$$S_2^*(0) = \alpha_0 + \alpha_1 S_1(0) + \alpha_2 S_3(0)$$

$$E\{S_2^*(0)\} = 0 \Rightarrow \alpha_0 = 0$$

$$\text{Var}\{S_2^*(0) - S_2(0)\}$$

$$= \alpha_1^2 \text{Var}\{S_1(0)\} + 2\alpha_1\alpha_2 \text{Cov}\{S_1(0), S_3(0)\} + \alpha_2^2 \text{Var}\{S_3(0)\} \\ - 2\alpha_1 \text{Cov}\{S_1(0), S_2(0)\} - 2\alpha_2 \text{Cov}\{S_3(0), S_2(0)\} \\ + \text{Var}\{S_2(0)\}$$

$$\left[\text{NOTE: } \text{Cov}\{S_{d_1}(0), S_{d_2}(0)\} = \iint_{-d/2}^{d/2} \text{Cov}\{R(y), R(z)\} dy dz = \min\{d_1, d_2\} \right]$$

$$= \alpha_1^2 + 2\alpha_1\alpha_2 + \alpha_2^2 \cdot 3 - 2\alpha_1 - 2\alpha_2 \cdot 3 + 2$$

$$\frac{d}{d\alpha_i} = 0 \Rightarrow \alpha_1 = \frac{1}{2}; \alpha_2 = \frac{1}{2}$$

$$S_2^*(0) = \frac{1}{2} S_1(0) + \frac{1}{2} S_3(0)$$

$$\text{Var}\{S_2^*(0) - S_2(0)\} = \underline{\underline{\frac{1}{2}}}$$

PROBLEM 2 Event Fields

Poisson point RF in \mathbb{R}^2

Arbitrary location x_0 , $R_{(i)}$ - distance to i th closest point

a) Definition of Poisson point RF:

A, B arbitrary domains in \mathbb{R}^2 with $A \cap B = \emptyset$

$N(A)$ - number of points in A

$N(B)$ - number of points in B

Poisson point RF $\Leftrightarrow [N(A), N(B)]$ independent

Poisson point RF on $D \subset \mathbb{R}^2$ with $|D| < \infty$

Poisson point RF $\Leftrightarrow \{X_1, \dots, X_N; D \subset \mathbb{R}^2\} \Leftrightarrow$

$N = N(D) \sim \text{Poi}(\lambda |D|) = -$ Poisson total per $\lambda |D|$

$[X_1, \dots, X_N | N]$ - iid $\text{Uni}[D]$ - Uniform over D

Given:

$$f(r_{(1)}, \dots, r_{(k)}) = \begin{cases} [2\lambda\pi]^k \prod_{i=1}^k r_{(i)} \exp\{-\lambda\pi r_{(k)}^2\} & ; 0 < r_{(1)} < \dots < r_{(k)} \\ 0 & \text{else} \end{cases}$$

b) Use induction:

Assume correct for $(k-1)$:

$$f(r_{(1)}, \dots, r_{(k-1)}) = \begin{cases} [2\lambda\pi]^{k-1} \prod_{i=1}^{k-1} r_{(i)} \exp\{-\lambda\pi r_{(k-1)}^2\} & ; 0 < r_{(1)} < \dots < r_{(k-1)} \\ 0 & \text{else} \end{cases}$$

then

$$f(r_{(1)}, \dots, r_{(k)}) = f(r_{(k)} | r_{(1)}, \dots, r_{(k-1)}) \cdot f(r_{(1)}, \dots, r_{(k-1)}) \quad (1)$$

with

$$F(r_{(k)} < r | r_{(1)}, \dots, r_{(k-1)}) = F(r_{(k)} < r | r_{(k-1)}) = 1 - F(r_{(k)} > r | r_{(k-1)})$$

$$= 1 - \text{Prob}\{N(\text{circle}) = 0\} = 1 - \exp\{-\lambda\pi(r_{(k)}^2 - r_{(k-1)}^2)\}$$

$$\Rightarrow f(r_{(k)} | r_{(1)}, \dots, r_{(k-1)}) = 2\lambda\pi r_{(k)} \exp\{-\lambda\pi(r_{(k)}^2 - r_{(k-1)}^2)\} \quad r > r_{(k-1)}$$

$$\text{From (1): } f(r_{(1)}, \dots, r_{(k)}) = [2\lambda\pi]^k \prod_{i=1}^k r_{(i)} \exp\{-\lambda\pi r_{(k)}^2\} \quad r_{(k)} > r_{(k-1)}$$

Since $f(r_{(1)}) = 2\lambda\pi r_{(1)} \exp\{-\lambda\pi r_{(1)}^2\}$ - is correct,

$f(r_{(1)}, \dots, r_{(k)})$ is correct by induction

c) Given $r_{(1)}$ and $r_{(3)}$, the likelihood is:

$$\begin{aligned} f(r_{(1)}, r_{(3)}) &= \int_{r_{(1)}}^{r_{(3)}} f(r_{(1)}, r_{(2)}, r_{(3)}) dr_{(2)} \\ &= (2\lambda\pi)^3 r_{(1)} r_{(3)} \exp\{-\lambda\pi r_{(3)}^2\} \int_{r_{(1)}}^{r_{(3)}} r dr \end{aligned}$$

$$= (2\lambda\pi)^3 r_{(1)} r_{(3)} \frac{1}{2} (r_{(3)}^2 - r_{(1)}^2) \exp\{-\lambda\pi r_{(3)}^2\}$$

MLE:

$$\hat{\lambda} = \operatorname{argmax}_{\lambda} \{ f(r_{(1)}, r_{(3)}; \lambda) \} = \operatorname{argmax}_{\lambda} \{ \ln f(r_{(1)}, r_{(3)}; \lambda) \}$$

$$\frac{d \cdot}{d\lambda} = 0 \Rightarrow 3 \frac{1}{\lambda} - \pi r_{(3)}^2 = 0 \Rightarrow$$

$$\hat{\lambda} = \frac{3}{\pi r_{(3)}^2} \quad - \text{which is not unreasonable!}$$

PROBLEM 3 Mosaic Fields

MRF with \mathcal{D}_s :



a)
$$\text{Prob}\{k=l\} = \text{const} \times \exp\left\{-\sum_{c \in \mathcal{C}} z_c(l)\right\}$$

\mathcal{C} -clique system corresponding to \mathcal{D}
cliques related to \mathcal{D}_s :

