

TMA 4250 ROMIG STATISTIKK
 VÅREN 2008
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LØSNINGSFORSLAG

Oppgave 1 Kontinuerlig stokastisk felt

$$\{R(x); x \in D \subset \mathbb{R}^2\}$$

$$E\{R(x)\} = \mu_R - ukjent$$

$$\text{Cov}\{R(x'), R(x'')\} = \sigma_R^2 \rho_R(x' - x'') - kjent$$

a) $\rho_R(x' - x'')$ - pos. def. korrelasjonsfunksjon

Krav:

$$\rho(0) = 1$$

$$\sum_{i=1}^n \sum_{j=1}^n \alpha_i \alpha_j \rho_R(x_i - x_j) > 0$$

$$\forall (\alpha_1, \dots, \alpha_n)$$

$$\forall n$$

b) Observasjoner: $R(x_1) = r_1$ og $R(x_2) = r_2$

Estimator: $\hat{\mu}_R = \alpha_0 + \alpha_1 R(x_1) + \alpha_2 R(x_2)$

Først. retthet

$$E\{\hat{\mu}_R - \mu_R\} = 0$$

$$\alpha_0 + \alpha_1 E\{R(x_1)\} + \alpha_2 E\{R(x_2)\} = \mu_R$$

$$\alpha_0 + (\alpha_1 + \alpha_2) \mu_R = \mu_R$$

$$\alpha_0 = 0$$

$$\alpha_1 + \alpha_2 = 1$$

Estimeringsvarians:

$$\text{Var}\{\hat{\mu}_R - \mu_R\} = \text{Var}\{\alpha_1 R(x_1) + \alpha_2 R(x_2) - \mu_R\}$$

$$= \alpha_1^2 \sigma_R^2 + \alpha_2^2 \sigma_R^2 + 2\alpha_1 \alpha_2 \sigma_R^2 \rho(x_1 - x_2)$$

Kriterium:

$$\hat{\alpha} = \underset{\alpha}{\operatorname{argmin}} \text{Var}\{\hat{\mu}_R - \mu_R\}$$

$$\alpha_1 + \alpha_2 = 1$$

herav

$$\frac{d}{d\alpha_1} \hat{\alpha} = \underset{\alpha_1}{\operatorname{argmin}} \left\{ \alpha_1^2 \sigma_R^2 + (1-\alpha_1)^2 \sigma_R^2 + 2\alpha_1(1-\alpha_1) \sigma_R^2 \rho(x_1 - x_2) \right\}$$

$$\frac{d}{d\alpha_1} = 0 \Rightarrow 2\alpha_1 \sigma_R^2 + 2(1-\alpha_1)(-1)\sigma_R^2 + 2(\sigma_R^2) \rho(x_1 - x_2) [(1-\alpha_1) + \alpha_1(-1)] = 0$$

$$[1-2\alpha_1] [\rho(x_1 - x_2) - 1] = 0 \Rightarrow \underline{\underline{\alpha_1 = \frac{1}{2}}}$$

$$\underline{\underline{\alpha_2 = 1 - \alpha_1 = \frac{1}{2}}}$$

Kan også sees ved symmetribetrakninger!!

$$\hat{\mu}_R = \frac{1}{2} [R(x_1) + R(x_2)]$$

$$\text{Var}\{\hat{\mu}_R - \mu_R\} = \frac{1}{2} \sigma_R^2 [1 + \rho(x_1 - x_2)]$$

c)

Observasjoner: $R(x_1) = r_1$ og $\Delta_{23} = [R(x_2) - R(x_3)] = \sigma_{23}$

Prediktor: $\hat{R}(x_0) = \alpha_0 + \alpha_1 R(x_1) + \alpha_2 \Delta_{23}$

Før rettet

$$E\{\hat{R}(x_0) - R(x_0)\} = 0$$

$$\Rightarrow \alpha_0 + \alpha_1 \mu_R + \alpha_2 \cdot 0 = \mu_R$$

$$\alpha_0 = 0$$

$$\alpha_1 = 1$$

α_2 - vilkårlig

Pred. varians

$$\begin{aligned} \text{Var}\{\hat{R}(x_0) - R(x_0)\} &= \text{Var}\{R(x_1) + \alpha_2 [R(x_2) - R(x_3)] - R(x_0)\} \\ &= \sigma_R^2 + \alpha_2^2 \left[2\sigma_R^2 - 2\alpha_2 \sigma_R^2 [\rho_R(x_2-x_3) + \sigma_R^2] + \sigma_R^2 \right. \\ &\quad \left. + 2\alpha_2 \sigma_R^2 [\rho_R(x_1-x_2) - \rho_R(x_1-x_3)] - 2\sigma_R^2 [\rho_R(x_1-x_0) - 2\alpha_2 \sigma_R^2 [\rho_R(x_0-x_2) - \rho_R(x_0-x_3)]] \right] \end{aligned}$$

Kriterium:

$$\begin{aligned} \hat{\alpha}_2 &= \underset{\alpha_2}{\operatorname{argmin}} \text{Var}\{\hat{R}(x_0) - R(x_0)\} \\ \frac{d}{d\alpha_2} &= 0 \Rightarrow 2\alpha_2 [2\sigma_R^2 (1 - \rho_R(x_2-x_3))] \\ &\quad + 2\sigma_R^2 [\rho_R(x_1-x_2) - \rho_R(x_1-x_3)] = 0 \\ \Rightarrow \hat{\alpha}_2 &= -2\sigma_R^2 [\rho_R(x_0-x_2) - \rho_R(x_0-x_3)] = 0 \\ \Rightarrow \hat{\alpha}_2 &= \frac{[\rho_R(x_0-x_2) - \rho_R(x_0-x_3)] - [\rho_R(x_1-x_2) - \rho_R(x_1-x_3)]}{2[1 - \rho_R(x_2-x_3)]} \end{aligned}$$

$$\hat{R}(x_0) = R(x_1) + \hat{\alpha}_2 \Delta_{23}$$

$$\text{Var}\{\hat{R}(x_0) - R(x_0)\} = \dots \quad \text{set min } \hat{\alpha}_2!$$

Oppgave 2 Hendelses stokastiske felt

$$\{\mathbb{X}_i; i=1, \dots, N; \mathcal{D} \subset \mathbb{R}^2\}$$

$$A, B \subset \mathcal{D}$$

$N(t)$ = * punkter i t

a) Homogent Poisson felt med int λ - kjent

$$N(t) - \text{Poi}(n; \lambda|t|) = \frac{[\lambda|t|]^n}{n!} e^{-\lambda|t|}$$

herav

$$\mathbb{E}\{N(t)\} = \lambda|t|$$

$$\text{Var}\{N(t)\} = \lambda|t|$$

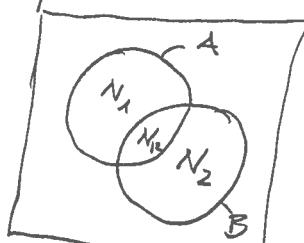
b) Homogent Poisson felt med int λ - kjent

$$\text{Cov}\{N(t), N(B)\}$$

$$= \text{Cov}\{(N_1 + N_{12}), (N_2 + N_{12})\}$$

$$= \text{Cov}\{N_1, N_2\} + \text{Cov}\{N_1, N_{12}\} + \text{Cov}\{N_{12}, N_2\} + \text{Cov}\{N_{12}, N_{12}\}$$

$$= 0 \quad = 0 \quad = 0 \quad = \text{Var}\{N_{12}\}$$



$$= \text{Var}\{N(A \cap B)\}$$

$$= \lambda|A \cap B|$$

Kommentarer:

- hvis $A = B$ $\text{Cov}\{N(t), N(t)\} = \text{Var}\{N(t)\} = \lambda|t|$ ok!

- hvis $A \cap B = \emptyset$ $\text{Cov}\{N(t), N(B)\} = \lambda|\emptyset| = 0$ ikke ok!
fordi $[N(t), N(B)]$ uavh.

c) Cox felt med stok. int $\lambda \Rightarrow f(\lambda) \begin{cases} E\{\lambda\} = \mu_\lambda \\ \text{Var}\{\lambda\} = \sigma_\lambda^2 \end{cases}$

Husk $E_x\{\lambda\} = E_y\{E_x\{\lambda|Y\}\}$

$$\begin{aligned} \text{Cov}\{N_A, N_B\} &= \text{Cov}\{N_A, N_B\} \\ &= E\{N_A \cdot N_B\} - E\{N_A\} E\{N_B\} \end{aligned}$$

videre:

$$\begin{aligned} E_n\{N_A\} &= E_\lambda\{E_n\{N_A|\lambda\}\} = E_\lambda\{\lambda|A|\} = \mu_\lambda|A| \\ E_n\{N_B\} &= \mu_\lambda|B| \end{aligned}$$

$$\begin{aligned} E\{N_A \cdot N_B\} &= E_\lambda\{E_n\{N_A N_B|\lambda\}\} \\ &= E_\lambda\left\{ \text{Cov}\{N_A, N_B|\lambda\} + E\{N_A|\lambda\} E\{N_B|\lambda\} \right\} \\ &= E_\lambda\left\{ \lambda|A \cap B| + \lambda^2|A||B| \right\} \\ &= \mu_\lambda|A \cap B| + |A||B| E_\lambda\{\lambda^2\} \\ &= \mu_\lambda|A \cap B| + |A||B| [\sigma_\lambda^2 + \mu_\lambda^2] \end{aligned}$$

setter inn:

$$\begin{aligned} \text{Cov}\{N_A, N_B\} &= \\ &= \mu_\lambda|A \cap B| + |A||B| [\sigma_\lambda^2 + \mu_\lambda^2] - \mu_\lambda|A| \cdot \mu_\lambda|B| \\ &= \underline{\underline{\mu_\lambda|A \cap B| + \sigma_\lambda^2|A||B|}} \end{aligned}$$

Kommentar:

- $\sigma_\lambda^2 = 0 \Leftrightarrow \lambda = \mu_\lambda \quad \text{Cov}\{N_A, N_B\} = \lambda|A \cap B| \quad \text{ok}$

- $\sigma_\lambda^2 \neq 0$ aldrig varit mellom $[N_A, N_B]$ - koplet via int λ !

Oppgave 3 Mosaikk stokastisk felt

$$\mathcal{L} : \{l_s ; s \in \mathcal{L}_0\} \quad l_s \in \{-1, 1\}$$

Naboskapsystem $\mathcal{J} : \{\mathcal{J}(s) ; s \in \mathcal{L}_0\}$

'Clique'-systeme $\mathcal{G} : \{c_1, \dots, c_n\}$

a) Gibbsformulering

$$\text{Prob}\{l = l\} = \text{const} \times \exp\left\{-\sum_{c \in \mathcal{G}} V_c(l)\right\}$$

med

$V_c(l)$ - funksjon av $l_s \in c$ bare

Markovformulering

$$\text{Prob}\{l_s = l_s | l_r = l_r ; r \neq s, r \in \mathcal{L}_0\}$$

$$= \text{Prob}\{l_s = l_s | l_r = l_r ; r \in \mathcal{J}(s)\} ; \forall s \in \mathcal{L}_0$$

Merk!

Gibbsformuleringen spesifiserer simultanfordelingen for hele \mathcal{L} .

Markovformuleringen består av univariate betingete fordelinger for hver node $s \in \mathcal{L}_0$

Hammersley-Clifford Teoremet impliserer at formuleringene er ekvivalente for tilhørende naboskaps/clique-systemet.

b) 'dlique' systemet består av alle:

$$c: \quad \dots \quad |$$

tilhørende naboskap:

$$\mathcal{S}(s): \quad \dots \underset{|}{\overset{i \in s}{x}} \dots$$

Gitt observasjonene $\mathcal{L} = \mathcal{L}^o : \{l_j^o ; j \in \mathcal{D}\}$

Full-likelihood:

$$\begin{aligned} FL(\beta) &= \text{Prob}\{ \mathcal{L} = \mathcal{L}^o ; \beta \} = \\ &= \text{const}(\beta) \times \exp \left\{ \beta \sum_{\substack{u, v \in \mathcal{D} \\ \langle u, v \rangle}} l_u^o l_v^o \right\} \end{aligned}$$

Maksimum full-likelihood estimator:

$$\hat{\beta} = \operatorname{argmax}_{\beta} \left\{ \log \overbrace{FL(\beta)}^{\text{log const}(\beta) + \beta \sum l_u^o l_v^o} \right\}$$

- det krever et uttrykk for $\text{const}(\beta)$ som er svært krevende å bestemme.

Pseudo-likelihood:

$$PL(\beta) = \prod_{s \in \mathcal{D}} \text{Prob}\{ l_s = l_s^o \mid l_r = l_r^o ; r \in \mathcal{S}(s); \beta \}$$

Maksimum pseudo-likelihood estimator:

$$\hat{\beta} = \operatorname{argmax}_{\beta} \left\{ \log \overbrace{PL(\beta)}^{\sum \log \text{Prob}\{ \cdot \}} \right\}$$

- konstanten i den betingete fordelingen er avh av β , men kan beregnes lett som sum av to ledsl. Selve optimeringen må gjøres numerisk