Chapter 13

• One-way ANOVA. Assumptions: k samples of sizes $n_1, n_2, ..., n_k$ (from k populations) are independent and normally distributed with means $\mu_1, \mu_2, ..., \mu_k$ and common variance σ^2 .

Hypothesis

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

$$H_1: \mu_i \neq \mu_j \text{ for at least one pair } (i, j)$$

• Model. Data

$$Y_{ij} = \mu_i + \epsilon_{ij}, \ i = 1, ..., k, \ j = 1, ..., n_i, \ \epsilon_{ij} \sim N(0, \sigma^2).$$

An alternative form

$$Y_{ij} = \mu + \alpha_i + \epsilon_{ij}, \ \mu = \frac{1}{k} \sum_{i=1}^k \mu_i$$

 $(\alpha_i \text{ is called the effect of the } i\text{-th treatment})$

$$H_0: \alpha_1 = \alpha_2 = \dots = \alpha_k = 0$$

 $H_1: \alpha_i \neq 0$ for at least one *i*

• Means and sums of squares:

$$\bar{Y}_{i.} = \frac{1}{n_i} \sum_{j=1}^{n_i} Y_{ij}, \ \bar{Y}_{..} = \frac{1}{N} \sum_{i=1}^k n_i \bar{Y}_{i.} \ \left(N = \sum_{i=1}^k n_i\right),$$
$$SST = \sum_{i=1}^k \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{..})^2$$

(total sum of squares),

$$SSA = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (\bar{Y}_{i.} - \bar{Y}_{..})^2$$

(treatment sum of squares),

$$SSE = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i.})^2$$

(error sum of squares).

- Properties:
- 1) SST = SSA + SSE,
- 2) SSA and SSE are independent,
- 3) $SSE/\sigma^2 \sim \chi^2_{N-k}$, 4) $SSA/\sigma^2 \sim \chi^2_{k-1}$ under H_0 .
- Test statistic

$$F = \frac{SSA/(k-1)}{SSE/(N-k)}.$$

Under H_0 F has the Fisher distribution with k-1 and N-k degrees of freedom. Test: if $F \ge f_{\alpha,k-1,N-k}$, then H_0 is rejected.

• Bartlett's test (testing equality of the variances). Sample variances:

$$S_i^2 = \frac{1}{n_i - 1} \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i.})^2$$

Pooled variance

$$S^{2} = \frac{1}{N-k} \sum_{i=1}^{k} (n_{i} - 1)S_{i}^{2}.$$

Test statistic

$$B = \frac{\left[(S_1^2)^{n_1 - 1} (S_2^2)^{n_2 - 1} \cdots (S_k^2)^{n_k - 1} \right]^{1/(N-k)}}{S^2}$$

If $n_1 = n_2 = \dots = n_k = n$, then H_0 is rejected if

$$B \le b_k(\alpha; n)$$

where $b_k(\alpha; n)$ is taken from the table of the Bartlett distribution. If n_i are different, then the critical value is approximative (see the book).

• A contrast

$$w = \sum_{i=1}^{k} c_i \mu_i$$

where $\sum_{i=1}^{k} c_i = 0$. Many hypotheses about μ -s can be formulated as

$$H_0: w = 0 \quad H_1: w \neq 0.$$

Test: if

$$\frac{(\sum_{i=1}^{k} c_i \bar{Y}_i)^2}{S_p^2 \sum_{i=1}^{k} (c_i^2/n_i)} \ge f_{\alpha, 1, N-k},$$

then H_0 is rejected.