



Contact person:
Håvard Rue 73593533/92600021

Exam in TMA4267 Linear Statistical Models
Tuesday June 2, 2009
Time: 09:00–13:00

Permitted assisting material: **None**

You may answer in English or Norwegian.
Du kan besvare enten på engelsk eller norsk.

Notation: $\mathbf{y} \sim \mathcal{N}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, denotes that $\mathbf{y} = (y_1, y_2, \dots, y_p)^T$ is normal distributed with mean $\boldsymbol{\mu}$ and covariance matrix $\boldsymbol{\Sigma}$. \mathbf{I} is the diagonal matrix.

Problem 1 Let

$$\mathbf{y} \sim \mathcal{N}_2 \left(\begin{pmatrix} 1 \\ 1 \end{pmatrix}, \begin{pmatrix} 9 & 2 \\ 2 & 9 \end{pmatrix} \right)$$

- a) What is the distribution of $z_1 = y_1 + y_2$?
 What is the distribution of $z_2 = z_1 - 2y_2$?
 What is $\text{Cov}(z_1, z_2)$?

Problem 2 Assume the data \mathbf{y} follows the model

$$\mathbf{y} = \mathbf{X}_1\boldsymbol{\beta}_1 + \mathbf{X}_2\boldsymbol{\beta}_2 + \boldsymbol{\epsilon}$$

where $\boldsymbol{\epsilon} \sim \mathcal{N}_n(\mathbf{0}, \sigma^2 \mathbf{I})$, but you estimate the following model using ordinary least squares,

$$\mathbf{y} = \mathbf{X}_1\boldsymbol{\beta}_1^* + \boldsymbol{\epsilon}^*.$$

Here, $\mathbf{X}_1 \neq \mathbf{X}_2$, and both \mathbf{X}_1 and \mathbf{X}_2 have full rank.

- a) Show that $\hat{\boldsymbol{\beta}}_1^*$ has the following properties

$$\mathbb{E}(\hat{\boldsymbol{\beta}}_1^*) = \boldsymbol{\beta}_1 + (\mathbf{X}_1^T \mathbf{X}_1)^{-1} \mathbf{X}_1^T \mathbf{X}_2 \boldsymbol{\beta}_2$$

and

$$\text{Cov}(\hat{\boldsymbol{\beta}}_1^*) = \sigma^2 (\mathbf{X}_1^T \mathbf{X}_1)^{-1}.$$

- b) Under what condition (of \mathbf{X}_1 and \mathbf{X}_2) is $\hat{\boldsymbol{\beta}}_1^*$ unbiased?
 What is the geometrical interpretation of this condition?

Problem 3 The Cook's distance is defined in the book as

$$D_i = \frac{(\hat{\mathbf{y}}_{(i)} - \hat{\mathbf{y}})^T (\hat{\mathbf{y}}_{(i)} - \hat{\mathbf{y}})}{(k+1)s^2}$$

- a) Explain each term in the Cook's distance.
 Explain the use of the Cook's distance.

Problem 4 Let

$$y_1, y_2, \dots, y_n \stackrel{\text{iid}}{\sim} \mathcal{N}_1(\mu, \sigma^2).$$

Define

$$\mathbf{j} = \begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix} \quad \text{and} \quad \mathbf{J} = \mathbf{j}\mathbf{j}^T = \begin{pmatrix} 1 & 1 & \dots & 1 \\ 1 & 1 & \dots & 1 \\ \vdots & \vdots & \ddots & \vdots \\ 1 & 1 & \dots & 1 \end{pmatrix}$$

$$\text{and } \bar{y} = \frac{1}{n} \sum_{i=1}^n y_i.$$

a) Show that

$$\sum_{i=1}^n y_i^2 = \sum_{i=1}^n (y_i - \bar{y})^2 + n\bar{y}^2$$

can be expressed as

$$\mathbf{y}^T \mathbf{I} \mathbf{y} = \mathbf{y}^T \left(\mathbf{I} - \frac{1}{n} \mathbf{J} \right) \mathbf{y} + \mathbf{y}^T \left(\frac{1}{n} \mathbf{J} \right) \mathbf{y}.$$

b) Show that

1. $\mathbf{I} - \frac{1}{n} \mathbf{J}$ and $\frac{1}{n} \mathbf{J}$ are idempotent.
2. $(\mathbf{I} - \frac{1}{n} \mathbf{J}) (\frac{1}{n} \mathbf{J}) = \mathbf{0}$

Problem 4 continuous on the next page...

Here is a transcript of two main results from the book:

Theorem 5.5. Let \mathbf{y} be distributed as $\mathcal{N}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, let \mathbf{A} be a symmetric matrix of constants of rank r , and let $\lambda = \frac{1}{2}\boldsymbol{\mu}^T \mathbf{A} \boldsymbol{\mu}$. Then $\mathbf{y}^T \mathbf{A} \mathbf{y}$ is $\chi^2(r, \lambda)$, if and only if $\mathbf{A} \boldsymbol{\Sigma}$ is idempotent.

Theorem 5.6a. Suppose \mathbf{B} is a $k \times p$ matrix of constants, \mathbf{A} is a $p \times p$ symmetric matrix of constants, and \mathbf{y} is distributed as $\mathcal{N}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$. Then $\mathbf{B}\mathbf{y}$ and $\mathbf{y}^T \mathbf{A} \mathbf{y}$ are independent if and only if $\mathbf{B}\boldsymbol{\Sigma}\mathbf{A} = \mathbf{0}$.

c) Use Theorem 5.5 and Theorem 5.6a to show that for $\boldsymbol{\mu} = \mathbf{0}$, then

1.

$$\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{\sigma^2} \sim \chi_{n-1}^2$$

2.

$$n \frac{\bar{y}^2}{\sigma^2} \sim \chi_1^2$$

3.

$$\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{\sigma^2} \quad \text{and} \quad n \frac{\bar{y}^2}{\sigma^2} \quad \text{are independent.}$$

d) Use the results in c) to construct a F -test (Fisher-test) for

$$H_0 : \boldsymbol{\mu} = \mathbf{0}, \quad H_1 : \boldsymbol{\mu} \neq \mathbf{0}$$

with significance level α .

e) Show that the test in d) is equivalent to a (classical) student- t test for testing the same hypothesis.