



Problem 1 Bivariate normal distribution

Assume that \mathbf{X} is a bivariate normal random variable with

$$\boldsymbol{\mu} = E\mathbf{X} = \begin{pmatrix} 0 \\ 2 \end{pmatrix} \quad \text{and} \quad \Sigma = \text{Cov } \mathbf{X} = \begin{pmatrix} 3 & 1 \\ 1 & 3 \end{pmatrix}.$$

Let

$$\mathbf{Y} = \begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} = \begin{pmatrix} 1/\sqrt{2} & -1/\sqrt{2} \\ 1/\sqrt{2} & 1/\sqrt{2} \end{pmatrix} \mathbf{X}.$$

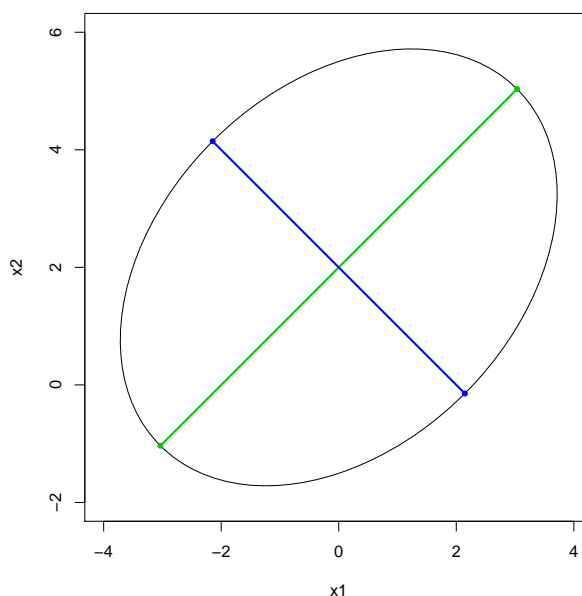
- a) Find the mean vector and covariance matrix of \mathbf{Y} . What is the distribution of \mathbf{Y} ? Are Y_1 and Y_2 independent random variables?

Let f be the pdf of \mathbf{X} . Contours of f are the \mathbf{x} satisfying $f(\mathbf{x}) = a$ for some constant $a > 0$, or equivalently $(\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) = b$ for a corresponding constant $b > 0$. In the figure below, the contour for $b = 4.6$ is shown.

- b) Explain the connections between the covariance matrix Σ , the chosen value of b and features of the ellipse (e.g. principal axes and their half-lengths). Mark these features on the figure (make a drawing or use the printed figure). What is the probability that \mathbf{X} falls within the given ellipse?

The following information might be useful:

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> sigma=matrix(c(3,1,1,3),ncol=2)
> sigma
      [,1] [,2]
[1,]    3    1
[2,]    1    3
> eigen(sigma)
$values
[1] 4 2
$vectors
      [,1]      [,2]
[1,] 0.7071068 -0.7071068
[2,] 0.7071068  0.7071068
> qchisq(0.9,2)
[1] 4.60517
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Problem 2 **Distributional results for \bar{X} and S^2 for a univariate normal sample**

Let X_1, X_2, \dots, X_n be a (univariate) random sample from some population. Define $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ and $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$.

Further, let $\mathbf{1}$ be an n -dimensional vector of 1s. Then $\mathbf{1}\mathbf{1}^T$ is an $n \times n$ matrix of 1s. Let I be an $n \times n$ identity matrix. The matrix

$$C = I - \frac{1}{n} \mathbf{1}\mathbf{1}^T = \begin{pmatrix} 1 - 1/n & -1/n & \cdots & -1/n \\ -1/n & 1 - 1/n & \cdots & -1/n \\ \vdots & \vdots & \ddots & \vdots \\ -1/n & -1/n & \cdots & 1 - 1/n \end{pmatrix}$$

is called the *centring matrix*. Let $\mathbf{X} = (X_1 \ X_2 \ \cdots \ X_n)^T$.

- a) Show that $\bar{X} = \frac{1}{n} \mathbf{1}^T \mathbf{X}$ and that $S^2 = \frac{1}{n-1} \mathbf{X}^T C \mathbf{X}$.

Hint: Start by considering the i th component of $C\mathbf{X}$, and observe that C is symmetric and idempotent.

Now, assume in addition that the random sample is taken from the univariate normal distribution with mean μ and variance σ^2 . In the notation of TMA4267, $\mathbf{X} \sim N(\mu\mathbf{1}, \sigma^2 I)$.

In your first statistics course, you were told that the T -statistic,

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}},$$

follows a t -distribution with $n-1$ degrees of freedom. This result follows from $(\bar{X} - \mu)/(\sigma/\sqrt{n}) \sim N(0, 1)$ and $(n-1)S^2/\sigma^2 \sim \chi_{n-1}^2$, and independence of \bar{X} and S^2 .

With your new skills on the multivariate normal distribution, you can prove independence of \bar{X} and S^2 , and $(n-1)S^2/\sigma^2 \sim \chi_{n-1}^2$.

- b) Show that $\frac{1}{n} \mathbf{1}^T C = \mathbf{0}^T$. What does this imply about $\frac{1}{n} \mathbf{1}^T \mathbf{X}$ and $C\mathbf{X}$? How can you use this to conclude that \bar{X} and S^2 are independent?
- c) Derive the distribution of $(n-1)S^2/\sigma^2$.

Hint: $S^2 = \frac{1}{n-1} \mathbf{X}^T C \mathbf{X}$, where C is symmetric and idempotent. In general, if R is a symmetric and idempotent matrix with rank r , and $\mathbf{Y} \sim N(\mathbf{0}, I)$, then $\mathbf{Y}^T R \mathbf{Y} \sim \chi_r^2$ (which we saw in a lecture and is stated by Fahrmeir, Kneib, Lang and Marx (2013) in Theorem B.8.2 on p. 651). Note, however, that \mathbf{X} is not assumed to be $N(\mathbf{0}, I)$.