

Maximum likelihood estimation of regression parameters

$$\boldsymbol{\varepsilon} : N(\boldsymbol{0}, \sigma^2 \mathbf{I}) \Rightarrow \hat{\boldsymbol{\beta}}_{ML} = \hat{\boldsymbol{\beta}}_{LS} \text{ and } \hat{\boldsymbol{\beta}}_{ML} \sim N\left(\boldsymbol{\beta}, \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}\right)$$

Partitioning of variation

$$\sum_{i=1}^n (y_i - \bar{y})^2 = \sum_{i=1}^n (y_i - \hat{y}_i)^2 + \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 \text{ or}$$

$$SS_T = SS_E + SS_R \text{ or}$$

$$\mathbf{y}^T \left(\mathbf{I} - \frac{1}{n} \mathbf{1} \mathbf{1}^T \right) \mathbf{y} = \mathbf{y}^T (\mathbf{I} - \mathbf{H}) \mathbf{y} + \mathbf{y}^T \left(\mathbf{H} - \frac{1}{n} \mathbf{1} \mathbf{1}^T \right) \mathbf{y}$$

where \mathbf{H} , $\frac{1}{n} \mathbf{1} \mathbf{1}^T$, $\mathbf{I} - \mathbf{H}$, $\mathbf{I} - \frac{1}{n} \mathbf{1} \mathbf{1}^T$, $\mathbf{H} - \frac{1}{n} \mathbf{1} \mathbf{1}^T$ are idempotent.

Specific results in linear regression

$$\text{R1 } E[\boldsymbol{\varepsilon}] = \mathbf{0}, Cov[\boldsymbol{\varepsilon}] = \sigma^2 \mathbf{I} \Rightarrow E[\hat{\sigma}_{LS}^2] = E\left[\sum_{i=1}^n \frac{(Y_i - \hat{Y}_i)^2}{n - (k+1)} \right] = \sigma^2$$

$$\text{R2 } \boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}) \Rightarrow \frac{\mathbf{Y}^T (\mathbf{I} - \mathbf{H}) \mathbf{Y}}{\sigma^2} \sim \chi^2(n - k - 1)$$

$$\text{R3a) } \left. \begin{array}{l} \boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}) \\ \beta_1 = \beta_2 = \dots = \beta_k = 0 \end{array} \right\} \Rightarrow \frac{\mathbf{Y}^T \left(\mathbf{I} - \frac{1}{n} \mathbf{1} \mathbf{1}^T \right) \mathbf{Y}}{\sigma^2} \sim \chi^2(n - 1)$$

$$\text{b)} \quad \left. \begin{array}{l} \boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}) \\ \beta_1 = \beta_2 = \dots = \beta_k = 0 \end{array} \right\} \Rightarrow \frac{\mathbf{Y}^T \left(\mathbf{H} - \frac{1}{n} \mathbf{1} \mathbf{1}^T \right) \mathbf{Y}}{\sigma^2} \sim \chi^2(k)$$

R4 $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}) \Rightarrow \hat{\boldsymbol{\beta}}$ and $\hat{\sigma}^2$ are independent

Fisher distribution

$$X \sim \chi^2(\nu_1), Y \sim \chi^2(\nu_2) \Rightarrow \frac{X/\nu_1}{Y/\nu_2} \sim F_{\nu_1, \nu_2}$$