

FORELESNING 10

Våren 2004

1. april

# TMA4275 LEVETIDSANALYSE

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	C1	C2	C3	C4	C5-T	C6	C7	C8	C9	C10	C11
	Temp	ArrTemp	Plant	FailureT	Censor	Design	NewTemp	ArrNewT	NewPlant		
1	170	26,1865		1	343 F		80	80	32,8600	1	
2	170	26,1865		1	869 F		100	80	32,8600	2	
3	170	26,1865		1	244 C			100	31,0988	1	
4	170	26,1865		1	716 F			100	31,0988	2	
5	170	26,1865		1	531 F						
6	170	26,1865		1	738 F						
7	170	26,1865		1	461 F						
8	170	26,1865		1	221 F						
9	170	26,1865		1	665 F						
10	170	26,1865		1	384 C						
11	170	26,1865		2	394 C						
12	170	26,1865		2	369 F						
13	170	26,1865		2	366 F						
14	170	26,1865		2	507 F						
15	170	26,1865		2	461 F						
16	170	26,1865		2	431 F						
17	170	26,1865		2	479 F						
18	170	26,1865		2	106 F						
19	170	26,1865		2	545 F						
20	170	26,1865		2	536 F						
21	150	27,4242		1	2134 C						
22	150	27,4242		1	2746 F						
23	150	27,4242		1	2859 F						
24	150	27,4242		1	1826 C						

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## Example of Accelerated Life Testing

 [main topic](#) [interpreting results](#) [session command](#) [see also](#)

Suppose you want to investigate the deterioration of an insulation used for electric motors. The motors normally run between 80 and 100° C. To save time and money, you decide to use accelerated life testing.

First you gather failure times for the insulation at abnormally high temperatures – 110, 130, 150, and 170° C – to speed up the deterioration. With failure time information at these temperatures, you can then extrapolate to 80 and 100° C. It is known that an Arrhenius relationship exists between temperature and failure time. To see how well the model fits, you will draw a probability plot based on the standardized residuals.

- 1 Open the worksheet INSULATE.MTW.
- 2 Choose **Stat > Reliability/Survival > Accelerated Life Testing**.
- 3 In **Variables/Start variables**, enter **FailureT**. In **Accelerating variable**, enter **Temp**.
- 4 From **Relationship**, choose **Arrhenius**.
- 5 Click **Censor**. In **Use censoring columns**, enter **Censor**, then click **OK**.
- 6 Click **Graphs**. In **Enter design value to include on plot**, enter **80**. Click **OK**.
- 7 Click **Estimate**. In **Enter new predictor values**, enter **Design**, then click **OK** in each dialog box.

*Session window output*

### Regression with Life Data: FailureT versus Temp

```
Response Variable: FailureT

Censoring Information          Count
Uncensored value                66
Right censored value             14
Censoring value: Censor = C

Estimation Method: Maximum Likelihood
Distribution: Weibull
Transformation on accelerating variable: Arrhenius
```

## Regression Table

Predictor	Coef	Error	Standard		95.0% Normal CI	
			Z	P	Lower	Upper
Intercept	-15.1874	0.9862	-15.40	0.000	-17.1203	-13.2546
Temp	0.83072	0.03504	23.71	0.000	0.76204	0.89940
Shape	2.8246	0.2570			2.3633	3.3760

Log-Likelihood = -564.693

## Anderson-Darling (adjusted) Goodness-of-Fit

At each accelerating level

Level	Fitted Model
110	*
130	*
150	*
170	*

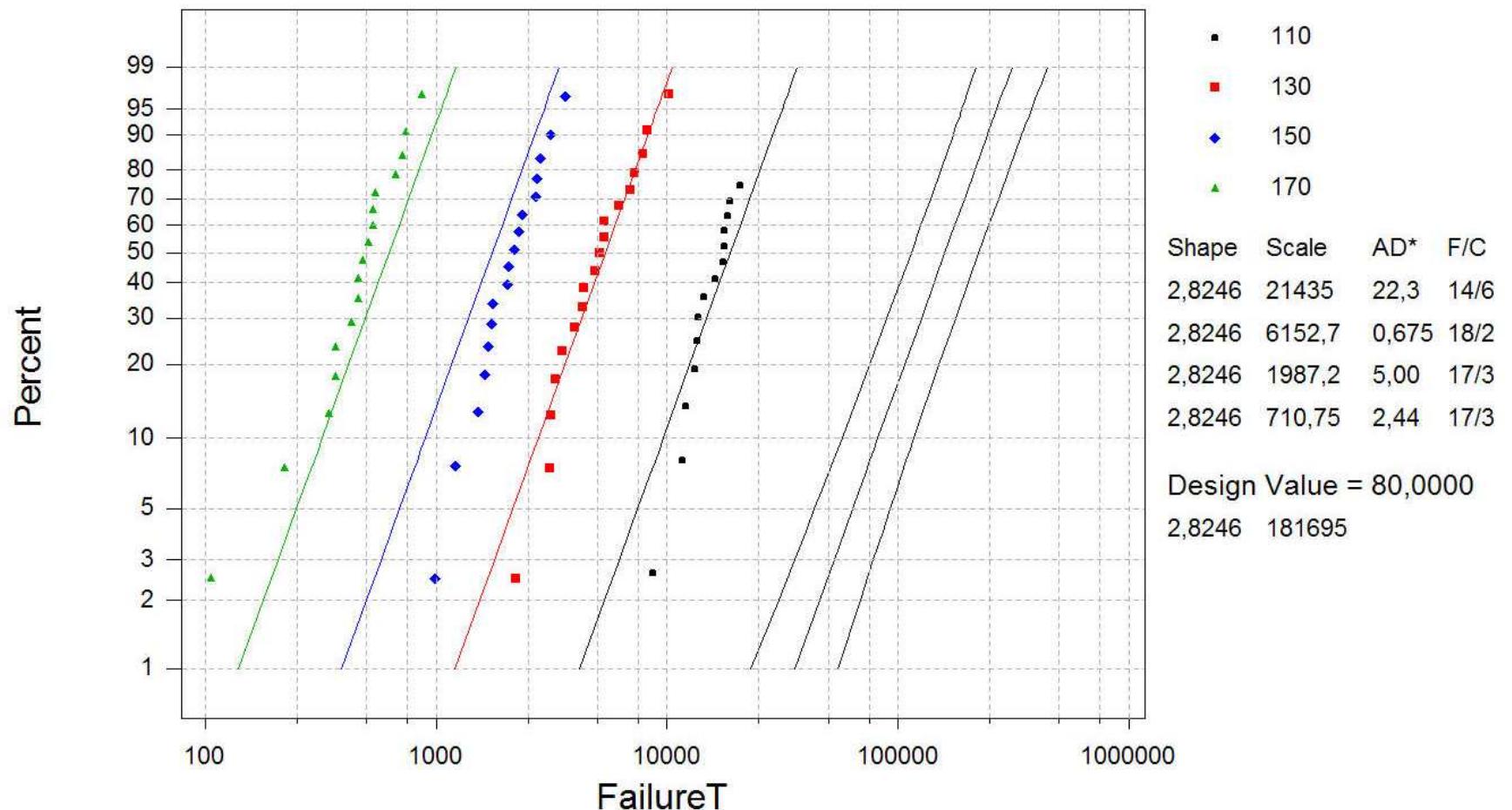
## Table of Percentiles

Percent	Temp	Percentile	Standard		95.0% Normal CI	
			Error	Lower	Upper	
50	80.0000	159584.5	27446.85	113918.2	223557.0	
50	100.0000	36948.57	4216.511	29543.36	46209.94	

## Probability Plot (Fitted Arrhenius) for FailureT

Weibull Distribution - ML Estimates - 95,0% CI

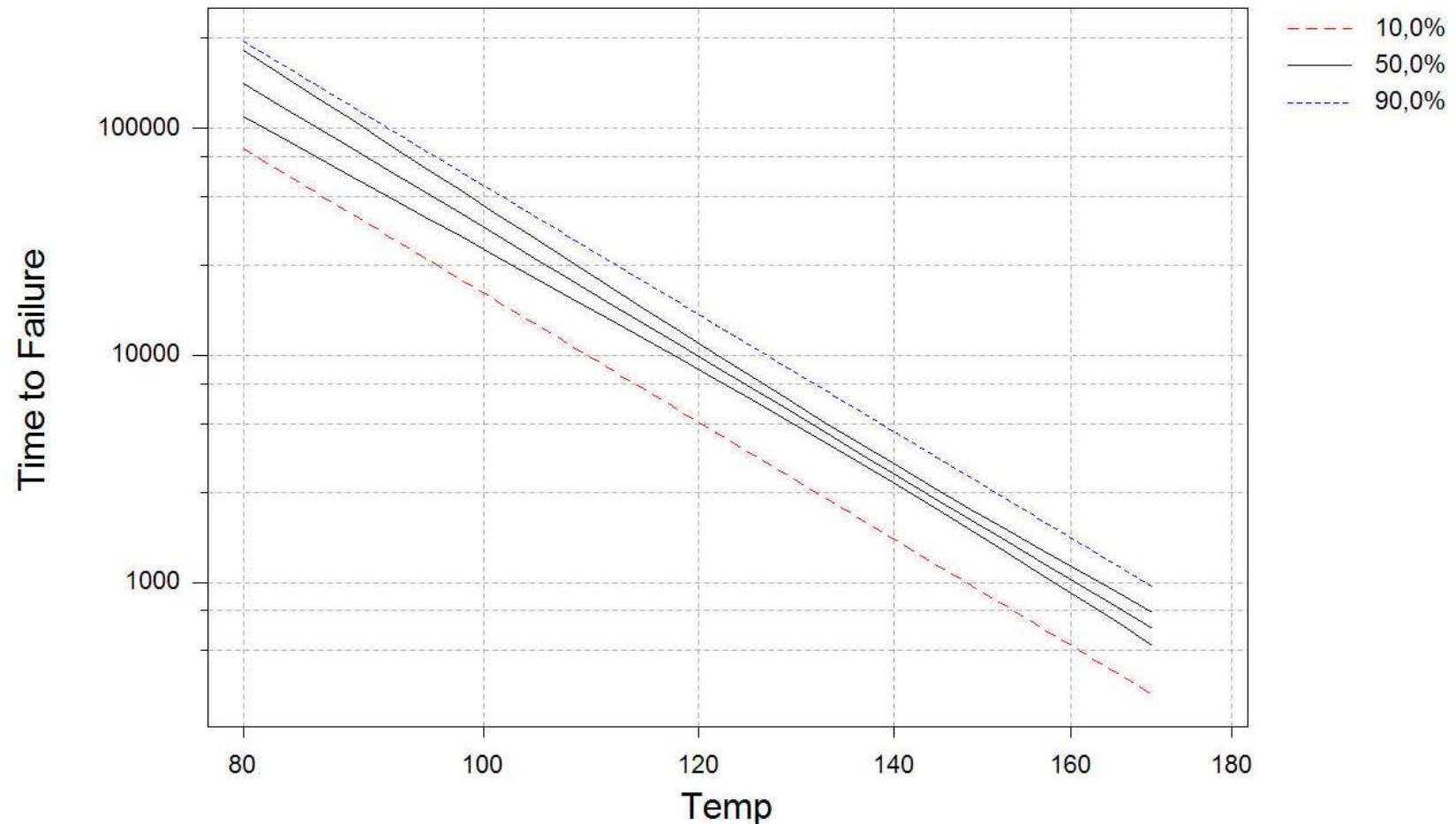
Censoring Column in Censor



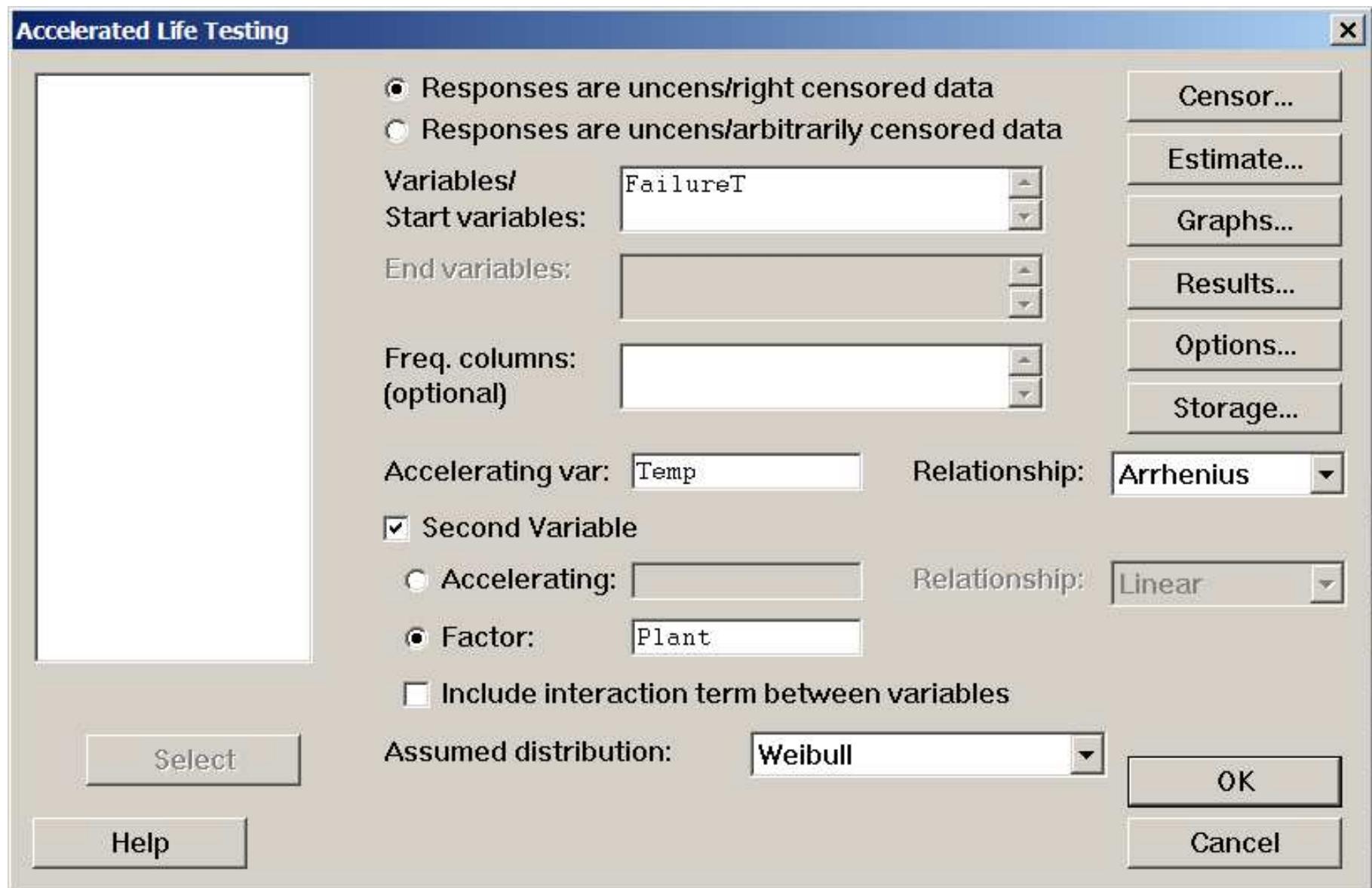
## Relation Plot (Fitted Arrhenius) for FailureT

Weibull Distribution - ML Estimates - 95,0% CI

Censoring Column in Censor



## MED FAKTOR "PLANT" I TILLEGG:



## EKSAMEN MAI 2003, OPPGAVE 3:

En komponent antas under normalstress å ha overlevelsfunksjon (reliability function)  $R_0(t)$  og hasardfunksjon (sviktintensitet)  $z_0(t)$ .

Man ønsker å estimere påliteligheten av denne komponenttypen ved hjelp av akselerert levetidstesting. Dette gjøres ved å utsette komponenter for stress  $s$ ,  $0 \leq s < \infty$ , og måle levetiden (eventuelt sensurerte levetider). Normalstress svarer til  $s = 0$ .

To modeller betraktes:

**Modell 1: Proporsjonal hasardmodell.** Under stress  $s$  har komponenten hasardfunksjon

$$z_s^{PH}(t) = z_0(t)g(s)$$

for en funksjon  $g(s)$  med  $g(0)=1$ .

**Modell 2: Akselerert levetidsmodell.** Under stress  $s$  har komponenten overlevelsfunksjon

$$R_s^{AL}(t) = R_0(\phi(s)t)$$

for en funksjon  $\phi(s)$  med  $\phi(0) = 1$ .

(a) Forklar kort hva som er hensikten med akselerert levetidstesting. Hva er ideen bak de to modellene? Hva uttrykker de to funksjonene  $g(s)$  og  $\phi(s)$ ?

(b) La  $R_s^{PH}(\cdot)$  være overlevelsesfunksjonen for en komponent under stress  $s$  i Modell 1. Vis at

$$R_s^{PH}(t) = R_0(t)^{g(s)}$$

La videre  $z_s^{AL}(\cdot)$  være hasardfunksjonen for en komponent under stress  $s$  i Modell 2. Uttrykk  $z_s^{AL}(\cdot)$  ved funksjonene  $z_0(\cdot)$  og  $\phi(\cdot)$ .

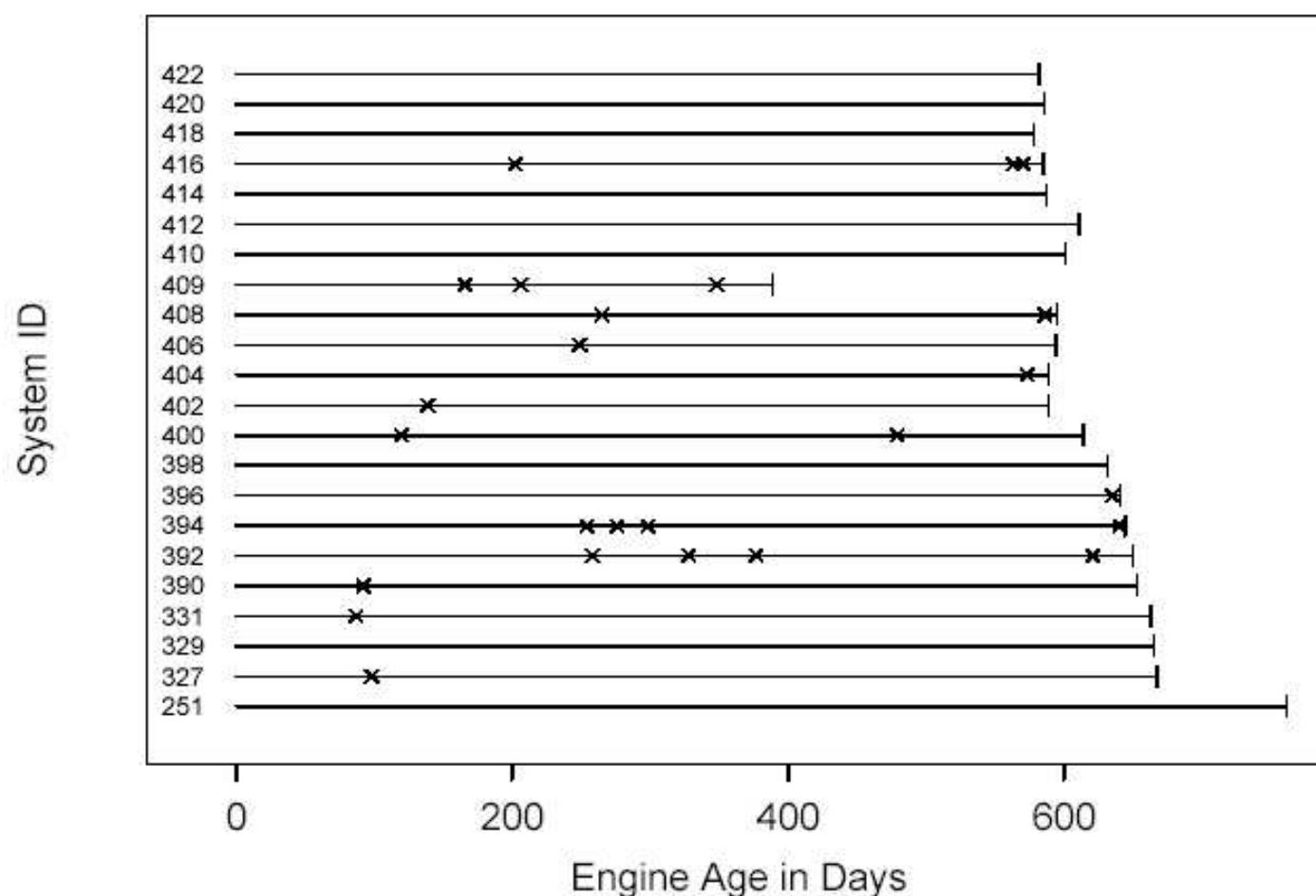
(c) Anta at komponentens levetid under normalstress er Weibull( $\alpha, \theta$ ), definert ved

$$R_0(t) = e^{-(t/\theta)^\alpha}$$

Vis at levetiden under stress  $s > 0$  også er Weibull-fordelt under begge modellene. Hva blir parametrene i de tilsvarende Weibull-fordelingene?

I hvilken forstand kan vi si at Modell 1 og Modell 2 er ekvivalente ved Weibull-fordelte levetider?

## Valve Seat Replacement Times Event Plot (Nelson and Doganaksoy 1989)

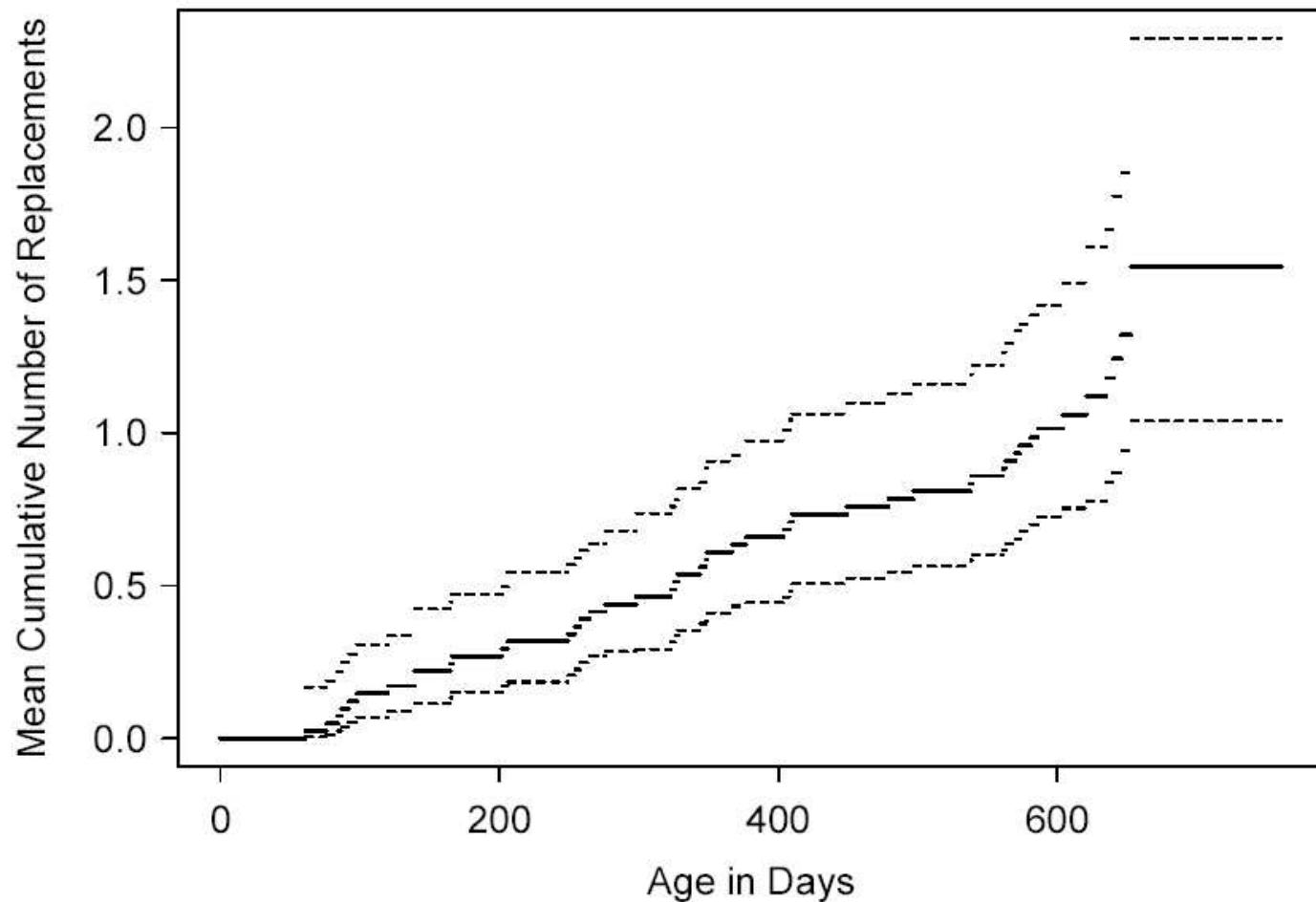


## **Valve Seat Replacement Times (Nelson and Doganaksoy 1989)**

Data collected from valve seats from a fleet of 41 diesel engines (days of operation)

- Each engine has 16 valves
- Does the replacement rate increase with age?
- How many replacement valves will be needed in the future?
- Can valve life in these systems be modeled as a renewal process?

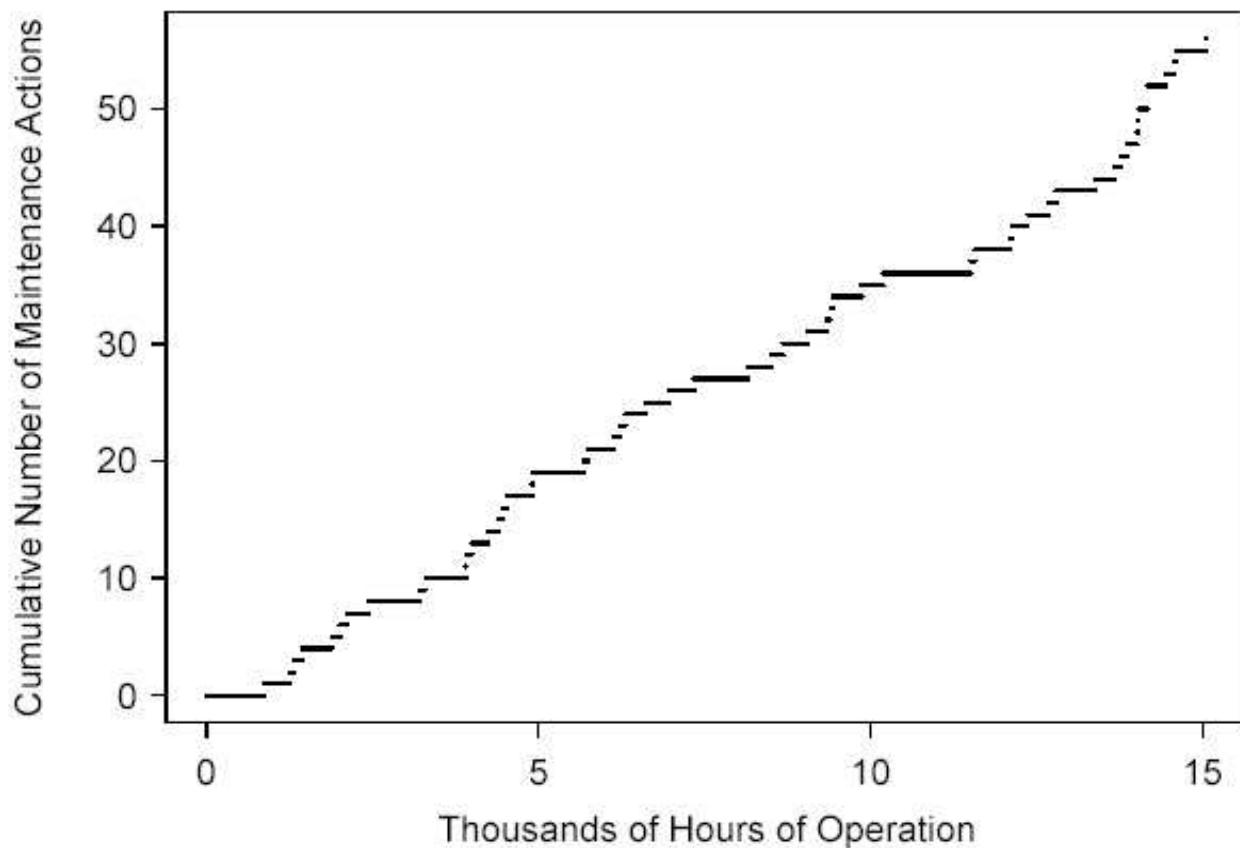
## Estimate of Number of Valve Seat $\mu(t)$



## **Times of Unscheduled Maintenance Actions for a USS Grampus Diesel Engine**

- Unscheduled maintenance actions caused by failure or imminent failure.
- Unscheduled maintenance actions are inconvenient and expensive.
- Data available for 16,000 operating hours.
- Data from Lee (1980).
- Is the system deteriorating (i.e., are failures occurring more rapidly as the system ages)?
- Can the occurrence of unscheduled maintenance actions be modeled by an HPP?

**Cumulative Number of Unscheduled Maintenance Actions Versus Operating Hours for a USS Grampus Diesel Engine**  
Lee (1980)



## The Likelihood for the NHPP - Single Unit

- With **interval** recurrence data.

Suppose that the unit has been observed for a period  $(0, t_a]$  and the data are the number of recurrences  $d_1, \dots, d_m$  in the nonoverlapping intervals  $(t_0, t_1], (t_1, t_2], \dots, (t_{m-1}, t_m]$  (with  $t_0 = 0, t_m = t_a$ ).

$$\begin{aligned} L(\boldsymbol{\theta}) &= \Pr [N(t_0, t_1) = d_1, \dots, N(t_{m-1}, t_m) = d_m] \\ &= \prod_{j=1}^m \Pr [N(t_{j-1}, t_j) = d_j] \\ &= \prod_{j=1}^m \frac{[\mu(t_{j-1}, t_j; \boldsymbol{\theta})]^{d_j}}{d_j!} \exp [-\mu(t_{j-1}, t_j; \boldsymbol{\theta})] \\ &= \prod_{j=1}^m \frac{[\mu(t_{j-1}, t_j; \boldsymbol{\theta})]^{d_j}}{d_j!} \times \exp [-\mu(t_0, t_a; \boldsymbol{\theta})] \end{aligned}$$

## The Likelihood for the NHPP (Continued)

- If the number of intervals  $m$  increases and there are **exact** recurrences at  $t_1 \leq \dots \leq t_r$  (here  $r = \sum_{j=1}^m d_j$ ,  $t_0 \leq t_1$ ,  $t_r \leq t_a$ ), then using a limiting argument it follows that the likelihood in terms of the density approximation is

$$L(\theta) = \prod_{j=1}^r \nu(t_j; \theta) \times \exp [-\mu(0, t_a; \theta)]$$

- For simplicity, above we assumed that the intervals are contiguous. Obvious changes to the formula above give the likelihood when there are gaps among the intervals.
- In both cases (the interval data or exact recurrences data) the same methods used in Chapters 7, 8 can be used to obtain the ML estimate  $\hat{\theta}$  and confidence regions for  $\theta$  or functions of  $\theta$ .

## CONDITIONAL ROCOF BY MINIMAL REPAIR (NHPP) AND PERFECT REPAIR (RENEWAL PROCESS)

