



English

SOLUTION SKETCH FOR
TMA4285 TIME SERIES MODELS
11. December 2008
Time: 09:00–13:00

Problem 1

- a) The ACF $\rho(k) = \rho_k$ for the given AR(1) process is given as $\rho_k = (-\phi)^k$ for $k = 0, 1, 2, \dots$. From the figures one may therefore read off the values of ϕ : 1) $\phi = -0.9$, 2) $\phi = -0.4$, 3) $\phi = 0.8$, 4) $\phi = 0.5$
- b) The plot of Y_t versus Y_{t-1} displays a significant negative correlation, while the plot of Y_t versus Y_{t-2} indicates a weak positive correlation (draw a vertical and horizontal line through the origin and study each of the quadrants). No such trend can be detected in the plot of Y_t versus Y_{t-3} . This indicates that the data is generated by an MA(2) model.
- c) The recorded time series displays both increasing variability (variance), trend and seasonality. Therefore, in order to stabilize the variability, a transformation of the data is suggested. This could be done by the Cox-Box transformation, a log transformation is a typical choice. After an appropriate transformation, the trend is removed by differencing. Here one differencing would seem sufficient. Finally, the seasonal component may be removed by s -differencing, where s denotes the identified seasonal period. The residual process after these operations can then be studied for a possible ARMA model fit.

Problem 2

- a) For X_t to be an ARMA(2,1) process it has to be stationary, and the AR and MA polynomials cannot have common roots. Stationarity is guaranteed by $\phi(z) \neq 0$ for $|z| = 1$ ($z \in \mathbf{C} =$ the complex numbers). That is, the roots of the AR polynomial $\phi(z) = 1 - z + \phi^2 z^2$ cannot lie on the unit circle. For $\phi = 0$, obviously $\phi(z) = 0$ for $z = 1$.

Hence, a necessary condition for X_t to be stationary is that $\phi \neq 0$. For $\phi \neq 0$, i.e. $\phi^2 > 0$, the roots of $\phi(z)$ are obtained as,

$$z^\pm = \frac{1 \mp \sqrt{1 - 4\phi^2}}{2\phi^2} = \frac{2}{1 \pm \sqrt{1 - 4\phi^2}}.$$

Can $|z^\pm| = 1$ for some value of $\phi \neq 0$? It is seen that z^\pm are real for $0 < \phi^2 \leq 1/4$, and that $|z^\pm| > 1$. For $\phi^2 > 1/4$, there will be two complex conjugate roots $z^\pm = 2/(1 \pm i\sqrt{4\phi^2 - 1})$ ($i = \sqrt{-1}$). It follows that $|z^\pm| = 1$ if $4 = 1 + (4\phi^2 - 1) = 4\phi^2$, which is satisfied for $\phi = \pm 1$. Conclusion: X_t is stationary for $\phi \notin \{0, \pm 1\}$

The requirement of no common roots need only be checked when $0 < \phi^2 \leq 1/4$. The root of the MA polynomial is $z = -1/\theta$. Hence, there will be no common roots when

$$\theta \neq \frac{-1 \pm \sqrt{1 - 4\phi^2}}{2}, \text{ for } 0 < \phi^2 \leq 1/4,$$

else θ is arbitrary.

- b) For X_t to be causal, $|z^\pm| > 1$. According to a) we only need to investigate the parameter range $\phi^2 > 1/4$. In this range, it follows that $|z^\pm| > 1$ if and only if $\phi^2 < 1$. Hence, X_t is causal if and only if $0 < \phi^2 < 1$.

X_t is invertible if and only if $1 + \theta z \neq 0$ for $|z| \leq 1$, that is, for $|1/\theta| > 1$. Hence, X_t is invertible if and only if $|\theta| < 1$.

- c) For the AR(2) process at hand there will be three cases to consider:

1. For $0 < \phi^2 < 1/4$, and $h \in \mathbf{Z}$,

$$\gamma(h) = c_1 \left(\frac{1 + \sqrt{1 - 4\phi^2}}{2} \right)^{|h|} + c_2 \left(\frac{1 - \sqrt{1 - 4\phi^2}}{2} \right)^{|h|},$$

where c_1 and c_2 are two real constants.

2. For $\phi^2 = 1/4$, and $h \in \mathbf{Z}$,

$$\gamma(h) = (c_1 + c_2|h|)2^{-|h|}$$

where c_1 and c_2 are two real constants.

3. For $1/4 < \phi^2 < 1$, and $h \in \mathbf{Z}$,

$$\gamma(h) = c \left(\frac{1 + i\sqrt{4\phi^2 - 1}}{2} \right)^{|h|} + \bar{c} \left(\frac{1 - i\sqrt{4\phi^2 - 1}}{2} \right)^{|h|},$$

where $c = c_1$ is a complex number in general, and $c_2 = \bar{c}$, which is the complex conjugate of c . This is necessary to make $\gamma(h)$ real. Using polar representation of a complex number, we may write $c = (a/2)e^{ib}$ for suitable real numbers a and b . Also

$$\frac{1 + i\sqrt{4\phi^2 - 1}}{2} = |\phi| e^{i\theta},$$

where $\tan \theta = \sqrt{4\phi^2 - 1}$. It is then obtained that

$$\gamma(h) = a|\phi|^{|h|} \cos(\theta|h| + b),$$

- d) Using the relation $X_t = X_{t-1} - \frac{1}{2}X_{t-2} + Z_t$, and calculating $E[X_t X_t]$ and $E[X_t X_{t+1}]$, the following two equations are obtained,

$$\gamma(1) = \frac{1}{4}\gamma(0) + \sigma^2$$

and

$$\gamma(2) = \frac{1}{2}\gamma(1) - \gamma(0) + 2\sigma^2 = -\frac{7}{8}\gamma(0) + \frac{5}{2}\sigma^2$$

From the previous point we know that

$$\gamma(h) = c\left(\frac{1+i}{2}\right)^{|h|} + \bar{c}\left(\frac{1-i}{2}\right)^{|h|},$$

which leads to the equations

$$c\frac{1+i}{2} + \bar{c}\frac{1-i}{2} = \frac{1}{4}(c + \bar{c}) + \sigma^2,$$

and

$$ci - \bar{c}i = -\frac{7}{8}(c + \bar{c}) + \frac{5}{2}\sigma^2.$$

Solving the equations gives the solution

$$c = \frac{4}{13}(7 + 2i)\sigma^2 = \frac{4}{13}\sqrt{53}e^{i0.278},$$

where $0.278 = \tan^{-1}(2/7)$. Noting that $(1+i)/2 = (1/\sqrt{2})e^{i\pi/4}$, it follows that

$$\gamma(h) = \frac{8}{13}\sqrt{53}\sigma^2\left(\frac{1}{2}\right)^{|h|/2} \cos\left(\frac{\pi}{4}|h| + 0.278\right),$$

The ACF then becomes

$$\rho(h) = \frac{\cos\left(\frac{\pi}{4}|h| + 0.278\right)}{2^{|h|/2} \cos(0.278)}, \quad h \in \mathbf{Z} \quad (1)$$

Problem 3

- a) Since ε_t is IID, it follows that ε_t is independent of ε_s for every $s < t$. This implies that ε_t is independent of $\varepsilon_{t-1}^2 \cdots \varepsilon_{t-j}^2$ for every $j = 1, 2, \dots$. Except for constants, $H(t)$ consists of a sum of such terms, and therefore ε_t and H_t are independent random variables for each t . Then ε_t and $\sqrt{H_t}$ are also independent random variables for each t . It is then obtained that,

$$\mathbb{E}[X_t] = \mathbb{E}[\varepsilon_t] \mathbb{E}[\sqrt{H_t}] = 0,$$

and

$$\begin{aligned} \mathbb{E}[X_t^2] &= \mathbb{E}[\varepsilon_t^2 H_t] = \mathbb{E}[\varepsilon_t^2] \mathbb{E}[H_t] = \mathbb{E}[H_t] \\ &= \alpha_0 \left(1 + \sum_{j=1}^{\infty} \alpha_1^j \mathbb{E}[\varepsilon_{t-1}^2] \cdots \mathbb{E}[\varepsilon_{t-j}^2]\right) = \alpha_0 \left(1 + \sum_{j=1}^{\infty} \alpha_1^j\right) = \frac{\alpha_0}{1 - \alpha_1}. \end{aligned}$$

Since ε_t is IID, it is seen that X_t is strictly stationary. Hence, moments, when they exist, are independent of t . From the equation $X_t^2 = \varepsilon_t^2(\alpha_0 + \alpha_1 X_{t-1}^2)$, it follows that

$$X_t^4 = \varepsilon_t^4 (\alpha_0^2 + 2\alpha_0\alpha_1 X_{t-1}^2 + \alpha_1^2 X_{t-1}^4),$$

Hence, if $m_4 = \mathbb{E}[X_t^4]$ exists, it must satisfy the equation ($\mathbb{E}[\varepsilon_t^4] = 3$),

$$m_4 = 3 \left(\alpha_0^2 + 2\alpha_0\alpha_1 \frac{\alpha_0}{1 - \alpha_1} + \alpha_1^2 m_4\right),$$

This leads to the equation

$$m_4 = \frac{3\alpha_0^2(1 + \alpha_1)}{(1 - \alpha_1)(1 - 3\alpha_1^2)}.$$

This equation can only be satisfied if $0 < \alpha_1^2 < 1/3$, which becomes the condition for finite m_4 .

- b) By a similar argument as above, it is seen that η_t is a strictly stationary process. By the assumption that $\mathbb{E}[X_t^4] < \infty$, it follows that η_t has finite second order moments. It is therefore a (weakly) stationary process.

$$\mathbb{E}[\eta_t] = \mathbb{E}[X_t^2] - \mathbb{E}[H_t] = 0.$$

and for $h \geq 1$,

$$\mathbb{E}[\eta_{t+h} \eta_t] = \mathbb{E}[(\varepsilon_{t+h}^2 - 1)(\varepsilon_t^2 - 1) H_{t+h} H_t] = \mathbb{E}[(\varepsilon_{t+h}^2 - 1)] \mathbb{E}[(\varepsilon_t^2 - 1) H_{t+h} H_t] = 0.$$

since ε_{t+h} is independent of ε_s for every $s < t+h$, and therefore $\varepsilon_{t+h}^2 - 1$ is independent of $(\varepsilon_t^2 - 1) H_{t+h} H_t$. It follows that η_t is white noise with variance

$$\begin{aligned} \sigma_0^2 &= \mathbb{E}[\eta_t^2] = \mathbb{E}[(\varepsilon_t^2 - 1)^2 H_t^2] = \mathbb{E}[(\varepsilon_t^2 - 1)^2] \mathbb{E}[H_t^2] = \\ &= \mathbb{E}[\varepsilon_t^4 - 2\varepsilon_t^2 + 1] \frac{\mathbb{E}[X_t^4]}{\mathbb{E}[\varepsilon_t^4]} = \frac{2}{3} m_4 = \frac{2\alpha_0^2(1 + \alpha_1)}{(1 - \alpha_1)(1 - 3\alpha_1^2)} \end{aligned}$$

c) From the definition of η_t , it follows that

$$X_t^2 = H_t + \eta_t = \alpha_0 + \alpha_1 X_{t-1}^2 + \eta_t. \quad (2)$$

Introducing the process $Y_t = X_t^2 - \alpha_0/(1 - \alpha_1)$, it is obtained that,

$$Y_t = \alpha_1 Y_{t-1} + \eta_t. \quad (3)$$

Since $0 < \alpha_1 < 1/\sqrt{3}$, it is seen that $\phi(z) = 1 - \alpha_1 z \neq 0$ for $|z| \leq 1$. It follows that Y_t becomes a causal AR(1) process, and therefore also X_t^2 (in the non-zero mean form).