

## TMA4295 Statistical inference

### Exercise 2 - solution

#### Problem 1

$$X \sim \Gamma(\alpha, \beta)$$

a) To find the distribution of  $cX$  we can use theorem 2.1.5.

$$y = cx \Rightarrow x = \frac{y}{c} \Rightarrow \frac{dx}{dy} = \frac{1}{c}$$

$$\Rightarrow f_Y(y) = \frac{1}{\Gamma(\alpha)\beta^\alpha} \frac{y^{\alpha-1}}{c^{\alpha-1}} e^{-\frac{y}{c\beta}} \frac{1}{c} = \frac{1}{\Gamma(\alpha)(c\beta)^\alpha} y^{\alpha-1} e^{-\frac{y}{c\beta}} \sim \Gamma(\alpha, c\beta)$$

$$E(Y) = cE(X) = c\alpha\beta$$

$$Var(Y) = c^2 Var(X) = c^2\alpha\beta^2.$$

b)  $\alpha = \frac{n}{2}$  and  $c = \frac{2}{\beta}$

$$\Rightarrow f_Y(y) = \frac{1}{\Gamma(\frac{n}{2})} \frac{1}{2^{n/2}} y^{n/2-1} e^{-\frac{y}{2}} \sim \chi_p^2$$

since Y it also a  $\Gamma(\frac{n}{2}, 2)$  then from part (a) we have that  $bY \sim \Gamma(\frac{n}{2}, 2b)$

c)  $Z_1, Z_2, \dots, Z_n \sim N(\mu, \sigma^2)$  and  $S^2$  variance estimator,

Recall that

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

and from the from the previous point we also know that  $\frac{(n-1)S^2}{\sigma^2} \sim \Gamma(\frac{n-1}{2}, 2b)$

$$\Rightarrow S^2 \sim \Gamma\left(\frac{n-1}{2}, \frac{2\sigma^2}{n-1}\right)$$

$$Var(S^2) = \frac{n-1}{2} \frac{4\sigma^4}{(n-1)^2} = \frac{2\sigma^4}{n-1}$$

d)  $X \sim \Gamma(\alpha, \beta)$  and  $k > -\alpha$

$$E(X^k) = \int_0^\infty \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha+k-1} e^{-\frac{x}{\beta}} dx = \frac{\beta^k}{\Gamma(\alpha)} \int_0^\infty \frac{x^{\alpha+k-1}}{\beta^{\alpha+k}} e^{-\frac{x}{\beta}} dx = \frac{\beta^k}{\Gamma(\alpha)} \Gamma(\alpha+k).$$

To compute  $E(S)$  we can use the previous part with  $k = 1/2$  since  $S^2$  has a gamma distribution. Hence

$$E(S) = \frac{\Gamma(\frac{n-1}{2} + \frac{1}{2}) \left(\frac{2\sigma^2}{n-1}\right)^{1/2}}{\Gamma(\frac{n-1}{2})} = \frac{\Gamma(\frac{n}{2}) 2^{1/2} \sigma}{\Gamma(\frac{n-1}{2})(n-1)^{1/2}}$$

e) We can then choose as unbiased estimator

$$\hat{S} = \frac{\Gamma(\frac{n-1}{2})(n-1)^{1/2}}{\Gamma(\frac{n}{2}) 2^{1/2}} S.$$

$$Var(\hat{S}) = \left(\frac{\Gamma(\frac{n-1}{2})(n-1)^{1/2}}{\Gamma(\frac{n}{2}) 2^{1/2}}\right)^2 Var(S) = \left(\frac{\Gamma(\frac{n-1}{2})(n-1)^{1/2}}{\Gamma(\frac{n}{2}) 2^{1/2}}\right)^2 (E(S^2) - E(S)^2) =$$

$$= \left(\frac{\Gamma(\frac{n-1}{2})(n-1)^{1/2}}{\Gamma(\frac{n}{2}) 2^{1/2}}\right)^2 (\sigma^2 - (\sigma/c)^2) = \sigma^2 \left(\frac{\Gamma(\frac{n-1}{2})^2 (n-1)}{\Gamma(\frac{n}{2})^2} - 1\right)$$

**Problem 3.20**

X random variable with the pdf  $f(x) = \frac{2}{\sqrt{2\pi}}e^{-x^2/2}$

a) Mean:

$$E(x) = \int_0^{\infty} \frac{2}{\sqrt{2\pi}}e^{-x^2/2}x dx = -\frac{2}{\sqrt{2\pi}}e^{-x^2/2}\Big|_0^{\infty} = \frac{2}{\sqrt{2\pi}}.$$

Variance: since  $Var(X) = E(x^2) - E(x)^2$  we need to compute  $E(x^2)$ .

$$\begin{aligned} E(x^2) &= \int_0^{\infty} \frac{2}{\sqrt{2\pi}}e^{-x^2/2}x^2 dx = -\frac{2}{\sqrt{2\pi}}e^{-x^2/2}x\Big|_0^{\infty} + \int_0^{\infty} \frac{2}{\sqrt{2\pi}}e^{-x^2/2}dx = 1 \\ &\Rightarrow Var(x) = 1 - \frac{2}{\pi}. \end{aligned}$$

b) We notice using the transformation  $y = x^2$  and so  $x = \sqrt{y}$  that:

$$f_Y(y) = \frac{2}{\sqrt{2\pi}} \frac{1}{2\sqrt{y}}e^{-y/2} = \frac{1}{\sqrt{\pi}} \frac{1}{\sqrt{2}}y^{-\frac{1}{2}}e^{-y/2} = \frac{1}{\Gamma(\frac{1}{2})2^{\frac{1}{2}}}y^{\frac{1}{2}-1}e^{(-y/2)}$$

that is gamma distributed with  $\alpha = \frac{1}{2}$  and  $\beta = 2$ .

**Problem 3.23**

The Pareto distribution has pdf :

$$f(x) = \frac{\beta\alpha^\beta}{x^{\beta+1}} \quad \alpha < x < \beta, \quad \alpha > 0, \quad \beta > 0.$$

a) Verify that  $f(x)$  is a pdf:

$$\int_{\alpha}^{\infty} \frac{\beta\alpha^\beta}{x^{\beta+1}} dx = \beta\alpha^\beta \left[ -\frac{1}{\beta}x^{-\beta} \right]_{\alpha}^{\infty} = 1.$$

b) Mean and variance:

$$\begin{aligned} E(x) &= \int_{\alpha}^{\infty} \frac{\beta\alpha^\beta}{x^{\beta+1}} x dx = \frac{\beta\alpha^\beta}{(\beta-1)\alpha^{\beta-1}} = \frac{\beta\alpha}{\beta-1} \\ E(x^2) &= \int_{\alpha}^{\infty} \frac{\beta\alpha^\beta}{x^{\beta+1}} x^2 dx = \frac{\beta\alpha^\beta}{(\beta-2)\alpha^{\beta-2}} = \frac{\beta\alpha^2}{\beta-2} \\ &\Rightarrow Var(x) = \frac{\beta\alpha^2}{\beta-2} - \left( \frac{\beta\alpha}{\beta-1} \right)^2. \end{aligned}$$

c)  $E(x^2)$  does not exist for  $\beta < 2 \Rightarrow$  the variance does not exist.

### Problem 3.30

a) Let's rewrite the pdf of a binomial random variable

$$f(x) = \binom{n}{x} p^x (1-p)^{n-x} = \binom{n}{x} (1-p)^n e^{x \log(\frac{p}{1-p})}$$

that is an exponential family with

$$c(p) = (1-p)^n, \quad t(x) = x, \quad \omega(p) = \log\left(\frac{p}{1-p}\right), \quad h(x) = \binom{n}{x}$$

$$\begin{aligned} \omega'(p) &= \frac{1}{p(1-p)} & \omega''(p) &= \frac{2p-1}{p^2(1-p)^2} \\ \frac{d}{dp} \log c(p) &= -\frac{n}{1-p} & \frac{d^2}{dp^2} \log c(p) &= -\frac{n}{(1-p)^2} \end{aligned}$$

Theorem 3.4.2 implies that

$$\begin{aligned} \text{Var} \left( \frac{1}{p(1-p)} X \right) &= \frac{n}{(1-p)^2} - E \left( \frac{X(2p-1)}{p^2(1-p)^2} \right) \\ \Rightarrow \frac{1}{p^2(1-p)^2} \text{Var}(X) &= \frac{n}{(1-p)^2} - \frac{np(2p-1)}{p^2(1-p)^2} \\ \Rightarrow \text{Var}(x) &= np^2 - 2np^2 + np = np(1-p). \end{aligned}$$