

TMA4295 Statistical inference

Exercise 4 - solution

Problem 4.15

$X \sim \text{Poisson}(\theta)$ and $Y \sim \text{Poisson}(\lambda)$
 $Z = X + Y \sim \text{Poisson}(\theta + \lambda)$

$$P(X = x|Z = z) = \frac{P(X = x, Y = y)}{P(Z = z)} = \frac{\frac{\theta^x e^{-\theta}}{x!} \frac{\lambda^{z-x} e^{-\lambda}}{(z-x)!}}{\frac{(\theta+\lambda)^z e^{-(\theta+\lambda)}}{z!}} = \binom{z}{x} \left(\frac{\theta}{\theta + \lambda}\right)^x \left(1 - \frac{\theta}{\theta + \lambda}\right)^{z-x}$$

which is a binomial with $p = \left(\frac{\theta}{\theta+\lambda}\right)$.

Problem 4.30

$Y|X = x \sim N(x, x^2)$ and $X \sim \text{Uniform}(0, 1)$.

a)

$$E(Y) = E(E(Y|X)) = E(X) = \frac{1}{2},$$

$$\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X)) = \frac{1}{12} + \frac{1}{3} = \frac{5}{12},$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{1}{3} - \left(\frac{1}{2}\right)^2 = \frac{1}{12}.$$

b) Consider the transformation $U = \frac{Y}{X}$ and $V = X$, then we have $x = v$ and $y = uv$ and the Jacobian is v . Since $f_{x,y}(x, y) = \frac{1}{2x\sqrt{\pi}} \exp\left(\frac{-(y-x)^2}{2x^2}\right) I_{(0,1)}(x)$ then

$$f_{U,V}(u, v) = \frac{1}{2v\sqrt{\pi}} e^{-\frac{(uv-v)^2}{2v^2}} I_{(0,1)}(v)v = g(u)h(v).$$

Problem 4.31

$Y|X = x \sim B(n, x)$ and $X \sim \text{Uniform}(0, 1)$.

a)

$$E(Y) = E(E(Y|X)) = E(nX) = \frac{n}{2},$$

$$\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X)) = E(nX(1-X)) + \text{Var}(nX) = \frac{n}{6} + \frac{n^2}{12}.$$

b)

$$f(x, y) = f(y|x)f(x) = \binom{n}{y} x^y (1-x)^{n-y} I_{(0,1)}.$$

c)

$$f_Y(y) = \int_0^1 f(x, y) dx = \int_0^1 \binom{n}{y} x^y (1-x)^{n-y} dx = \binom{n}{y} \frac{\Gamma(y+1)\Gamma(n-y+1)}{\Gamma(n+2)}.$$

Problem 4.32

$Y|\Lambda \sim \text{Poisson}(\Lambda)$ and $\Lambda \sim \text{Gamma}(\alpha, \beta)$.

a)

$$\begin{aligned} f_Y(y) &= \int_0^\infty f(y|\lambda)f(\lambda)d\lambda = \int_0^\infty \frac{\lambda^y e^{-\lambda}}{y!} \frac{1}{\Gamma(\alpha)\beta^\alpha} \lambda^{\alpha-1} e^{-\frac{\lambda}{\beta}} d\lambda \\ &= \frac{1}{y!\Gamma(\alpha)\beta^\alpha} \int_0^\infty \lambda^{y+\alpha-1} \exp\left(\frac{-\lambda}{1+\beta}\right) d\lambda = \frac{1}{y!\Gamma(\alpha)\beta^\alpha} \Gamma(y+\alpha) \left(\frac{\beta}{1+\beta}\right)^{y+\alpha}. \end{aligned}$$

If α is an integer

$$f_Y(y) = \binom{y+\alpha-1}{y} \left(\frac{\beta}{1+\beta}\right)^y \left(\frac{1}{1+\beta}\right)^\alpha$$

that is a negative binomial.

$$E(Y) = E(E(Y|\Lambda)) = E(\Lambda) = \alpha\beta,$$

$$\text{Var}(Y) = \text{Var}(E(Y|\Lambda)) + E(\text{Var}(y|\Lambda)) = E(\Lambda) + \text{Var}(\Lambda) = \alpha\beta + \alpha\beta^2 = \alpha\beta(\beta+1).$$

b) $Y|N \sim \text{Binomial}(N, p)$, $N|\Lambda \sim \text{Poisson}(\Lambda)$ and $\Lambda \sim \text{gamma}(\alpha, \beta)$ from example 4.4.2 we see that

$$P(Y = y|\lambda) = \sum_{n=y}^{\infty} P(Y = y|N = n, \lambda)P(N = n|\lambda) = \frac{(p\lambda)^y e^{-p\lambda}}{y!},$$

so we have $Y|\Lambda \sim \text{Poisson}(p\Lambda)$, calculations similar to those in a) prove that Y is negative binomial distributed (if α is a positive integer).

Problem 4.35

$X|P \sim \text{Binomial}(n, P)$ and $P \sim \text{Beta}(\alpha, \beta)$

a)

$$\begin{aligned} \text{Var}(X) &= E(\text{Var}(X|P)) + \text{Var}(E(X|P)) = E(nP(1-P)) + \text{Var}(nP) \\ &= n(E(P) - E(P^2)) + n^2 \text{Var}(P) = nE(P) - n\text{Var}(P) - n(E(P)^2) + n^2 \text{Var}(P) \\ &= nE(P)(1 - E(P)) + n(n-1)\text{Var}(P). \end{aligned}$$

b)

$$\text{Var}(Y) = \alpha\beta + \alpha\beta^2 = \mu + \frac{\mu^2}{\alpha}.$$

Problem 4.36

$X_i|P_i \sim \text{Bernoulli}(P_i)$ $i = 1, \dots, k$

$P_i \sim \text{Beta}(\alpha, \beta)$.

We first compute

$$E(X_i) = E(E(X_i|P_i)) = E(P_i) = \frac{\alpha}{\alpha + \beta},$$

$$\text{Var}(X_i) = E(\text{Var}(X_i|P_i)) + \text{Var}(E(X_i|P_i)) = \frac{\alpha\beta}{(\alpha + \beta)^2}.$$

a) Since $Y = \sum_{i=1}^k X_i$

$$E(X) = \sum_{i=1}^k E(X_i) = \frac{n\alpha}{\alpha + \beta}.$$

b)

$$\text{Var}(X) = \sum_{i=1}^k \text{Var}(X_i) = \frac{n\alpha\beta}{(\alpha + \beta)^2}.$$

To find the distribution of Y let's consider the mgf

$$M_Y(t) = E(e^{\sum_i x_i t}) = E(e^{x_1 t})^n = E(E(e^{x_1 t}|p_1))^n.$$

Now

$$\begin{aligned} E(e^{x_1 t}|p_1) &= \sum_{x_1=0}^1 e^{x_1 t} (p_1^{x_1} (1-p)^{1-x_1}) = 1 - p + pe^t \\ \Rightarrow M_Y(t) &= E(1 - p + pe^t)^n = \left(1 - \frac{\alpha}{\alpha + \beta} + e^t \frac{\alpha}{\alpha + \beta}\right)^n. \end{aligned}$$

Which is the mgf of a *binomial*($n, \frac{\alpha}{\alpha + \beta}$).

c) $X_i|P_i \sim \text{Binomial}(n_i, P_i)$ and $P_i \sim \text{Beta}(\alpha, \beta)$ for $i = 1, \dots, k$.

$$E(X_i) = E(E(X_i|P_i)) = E(n_i P_i) = \frac{n_i \alpha}{\alpha + \beta},$$

$$\text{Var}(X_i) = E(\text{Var}(X_i|P_i)) + \text{Var}(E(X_i|P_i)) = \frac{n_i \alpha \beta (\alpha + \beta + n_i)}{(\alpha + \beta)^2 (\alpha + \beta + 1)},$$

$$\Rightarrow E(Y) = \sum_{i=1}^k E(X_i) = \frac{\alpha}{\alpha + \beta} \sum_{i=1}^k n_i$$

$$\text{Var}(Y) = \sum_{i=1}^k \text{Var}(X_i).$$

Problem 4.58

a)

$$\begin{aligned} \text{Cov}(X, Y) &= \int \int (x - \mu_x)(y - \mu_y) f(x, y) dx dy \\ &= \int (x - \mu_x) \int (y - \mu_y) f(y|x) dy f(x) dx \\ &= \int (x - \mu_x) (E(Y|X) - \mu_y) f(x) dx \\ &= \text{Cov}(X, E(Y|X)). \end{aligned}$$

b)

$$\begin{aligned} \text{Cov}(X, (Y - E(Y|X))) &= E((X - \mu_x)(Y - E(Y|X))) \\ &= \int \int (x - \mu_x)(y - E(Y|x)) f(y|x) dy f(x) dx \\ &= \int (x - \mu_x) \int (y - E(Y|x)) f(y|x) dy f(x) dx \\ &= \int (x - \mu_x) (E(Y|x) - E(Y|x)) f(x) dx = 0. \end{aligned}$$

c)

$$\begin{aligned} \text{Var}(Y - E(Y|X)) &= \int \int (y - E(Y|x))^2 f(x, y) dx dy \\ &= \int \int (y - E(Y|x))^2 f(y|x) dy f(x) dx \\ &= \int \text{Var}(Y|x) f(x) dx = E(\text{Var}(Y|X)). \end{aligned}$$