

Some examples. One mother.

$$Y \sim \text{Po}(\lambda), \quad X|Y \sim \text{B}(Y, p)$$

$$E[X|Y] = pY \Rightarrow E[X] = E[E[X|Y]] = pE[Y] = p\lambda$$

$$\begin{aligned} \text{Var}[X|Y] &= Yp(1-p) \Rightarrow \text{Var}[X] = \lambda p(1-p) + p^2 \lambda \\ &= \lambda p(1-p + p) = \lambda p \end{aligned}$$

Randomly chosen mother

$$\left. \begin{array}{l} X|Y \sim \text{B}(Y, p) \\ Y|\Lambda \sim \text{Po}(\Lambda) \\ \Lambda \sim \text{exp}\left(\frac{1}{\beta}\right) \end{array} \right\} \begin{array}{l} X|\Lambda \sim \text{Po}(p\Lambda) \\ Y \sim \text{translated negative} \\ \text{binomial with } n=1 \\ p = \frac{1}{1+\beta} \end{array}$$

$$E[X] = E[E[X|Y]] = E[pY] = pE[Y] = pE[E(Y|\Lambda)] = pE[\Lambda] = p\beta$$

$$\begin{aligned} \text{Var}[X] &= E[\text{Var}(X|Y)] + \text{Var}[E[X|Y]] \quad \left| \begin{array}{l} \text{Negative Binomial with } n=1 \\ \text{Var}[Y] = \mu(1+\mu) \\ \left(\text{In general } \mu\left(1 + \frac{\mu}{n}\right) \right) \\ p^a = \frac{1}{1+p\beta} \end{array} \right. \\ &= E[Yp(1-p)] + \text{Var}[pY] \\ &= p(1-p)E[Y] + p^2 \text{Var}[Y] \\ &= p(1-p)E[E(Y|\Lambda)] + p^2[E[\text{Var}(Y|\Lambda)] + \text{Var}[E(Y|\Lambda)]] \\ &= p(1-p)E[\Lambda] + p^2[E[\Lambda] + \text{Var}[\Lambda]] \\ &= p\beta - p^2\beta + p^2(\beta + \beta^2) = p\beta(1 + p\beta) \end{aligned}$$

For covariance

$$E[XY] = E[E[XY|X]] = E[XE[Y|X]]$$

4.5 Covariance and Correlation

Definition 4.5.1

The covariance of X and Y is defined by

$$\text{Cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X] \cdot E[Y]$$

Definition 4.5.2

The correlation (correlation coefficient) of X and Y is the

number defined by
$$\rho_{X, Y} = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$$

Consider the function $h(t) = E[(X - \mu_X)t + (Y - \mu_Y)]^2$

$$= t^2 \sigma_X^2 + 2t \text{Cov}(X, Y) + \sigma_Y^2$$

$$h'(t) = 2t \sigma_X^2 + 2 \text{Cov}(X, Y) = 0 \Rightarrow t = -\frac{\text{Cov}(X, Y)}{\sigma_X^2}$$

$$h''(t) = \sigma_X^2 > 0 \Rightarrow \text{minimum}$$

Substituted for t gives

$$h(t) = \frac{\text{Cov}(X, Y)^2 \sigma_X^2}{\sigma_X^4} - 2 \frac{\text{Cov}(X, Y)^2}{\sigma_X^2} + \sigma_Y^2$$

$$= -\frac{\text{Cov}(X, Y)^2}{\sigma_X^2} + \sigma_Y^2 \geq 0$$

or
$$\text{Cov}(X, Y)^2 \leq \sigma_X^2 \sigma_Y^2 \Leftrightarrow -\sigma_X \sigma_Y \leq \text{Cov}(X, Y) \leq \sigma_X \sigma_Y$$

$$\Rightarrow |\rho_{X, Y}| \leq 1$$

Since $((X-\mu_X)t + (Y-\mu_Y))^2 \geq 0$ we have
 $E[((X-\mu_X)t + (Y-\mu_Y))^2] = 0$
 $h(t) = 0 \iff P[((X-\mu_X)t + (Y-\mu_Y))^2 = 0] = 1$. ~~to do~~

$$\Rightarrow P[(X-\mu_X)t + (Y-\mu_Y) = 0] = 1 \iff$$

$$P(Y = -tX + \mu_X t + \mu_Y) = 1$$

Theorem 4.5.7

X, Y random variables

a) $-1 \leq \rho_{X,Y} \leq 1$

b) $|\rho_{X,Y}| = 1 \iff P(Y = aX + b) = 1$

$\rho_{X,Y} > 0 \Rightarrow$ t negative and a positive.

Example 4.5.9

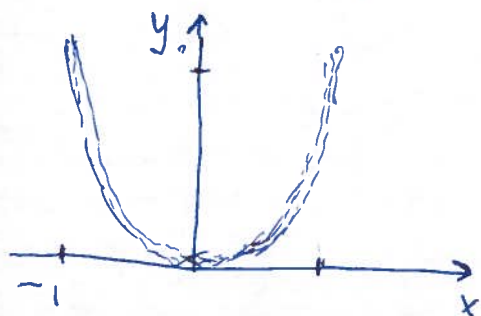
$X \sim$ uniform $[-1, 1]$ — independent

$Z \sim$ uniform $[0, \frac{1}{10}]$ /

$$Y = X^2 + Z$$

$$\text{Cov}(X, Y) = E[(X^2 + Z)X] - E[X^2 + Z] \cdot E[X]$$

$$= E[X^3] + E[Z] \cdot E[X] = E[X^3] = \frac{1}{2} \int_{-1}^1 x^3 dx = 0$$



Correlation does not measure non linear relationship.

4.7.

Inequalities.

Lemma 4.7.1 Let a, b be positive numbers and let p and q be any positive (> 1) ^{numbers} satisfying

$$\frac{1}{p} + \frac{1}{q} = 1 \quad \Rightarrow \quad \left(q = \frac{p}{p-1} \right)$$

Then $\frac{1}{p} a^p + \frac{1}{q} b^q \geq ab$ with equality if and only

$$\text{if } a^p = b^q$$

Proof. For given b let $g(a) = \frac{1}{p} a^p + \frac{1}{q} b^q - ab$

$$\frac{d}{da} g(a) = a^{p-1} - b = 0 \quad \Rightarrow \quad a = b^{\frac{1}{p-1}} \quad \text{or} \quad b = a^{p-1}$$

$$\frac{d^2}{da^2} g(a) > 0 \quad \Rightarrow \quad \text{minimum.}$$

$$g(b^{\frac{1}{p-1}}) = \frac{1}{p} b^{\frac{p}{p-1}} + \frac{1}{q} b^q - b^{\frac{1}{p-1}} \cdot b$$

$$= \frac{1}{p} b^q + \frac{1}{q} b^q - b^q = b^q \left(\frac{1}{p} + \frac{1}{q} \right) - b^q = 0$$

$\Rightarrow g(a) \geq 0$ and the ~~to~~ result follows.

$$g(a) = 0 \quad \Rightarrow \quad a^p = b^{\frac{p}{p-1}} = b^q.$$

Theorems 4.7.2

Hölder's inequality. X, Y random variables and $p > 0, q > 0$.

satisfies $\frac{1}{p} + \frac{1}{q} = 1$

Then $|E[XY]| \leq E|XY| \leq (E|X|^p)^{\frac{1}{p}} \cdot (E|Y|^q)^{\frac{1}{q}}$

Proof. $-|XY| \leq XY \leq |XY| \Rightarrow -E(|XY|) \leq E[XY] \leq E(|XY|)$

$$\Leftrightarrow |E(XY)| \leq E(|XY|)$$

$$\text{Let } a = \frac{|X|}{(E|X|^p)^{\frac{1}{p}}} \quad \text{and} \quad b = \frac{|Y|}{(E|Y|^q)^{\frac{1}{q}}}$$

From Lemma 4.7.1

$$E \left[\frac{1}{p} \frac{|X|^p}{(E|X|^p)} + \frac{1}{q} \frac{|Y|^q}{E|Y|^q} \right] \geq \frac{E(|XY|)}{(E|X|^p)^{\frac{1}{p}} (E|Y|^q)^{\frac{1}{q}}}$$

$$\Rightarrow E(|XY|) \leq (E|X|^p)^{\frac{1}{p}} (E|Y|^q)^{\frac{1}{q}}$$

In particular Theorem 4.7.3 Cauchy Schwarz

$$|E(XY)| \leq E(|XY|) \leq (E|X|^2)^{\frac{1}{2}} (E|Y|^2)^{\frac{1}{2}}$$

$$\text{or } (\text{Cov}(X, Y))^2 \leq \sigma_x^2 \sigma_y^2.$$