

Enough with two walls. The two at the end can be removed

$$\begin{array}{ccc} \text{MIMIM} & \text{IMMIM} & \text{MMIMI} \\ | \quad 3 & 2 \quad 2 & 1 \quad 1 \end{array}$$

Two walls ($m-1$) and two numbers can be arranged in $(3-1+2)! = ((m-1+n)!)$ ways.

To remove double counting we get.

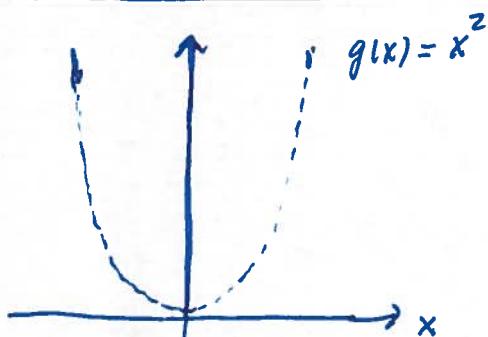
$$\frac{4!}{2! 2!} \quad \text{or} \quad \frac{(m-1+n)!}{(m-1)! n!} = \binom{m-1+n}{n}$$

Chapter 2. Transformations and Expectations

$$Y = g(X) : g(x) : X \rightarrow Y$$

$$F_Y(y) = P(g(X) \leq y) = P(X \in \mathcal{X} : g(x) \leq y) \stackrel{\text{cont.}}{=} \int_{\{x : g(x) \leq y\}} f_X(x) dx$$

Example 2.1.7



X continuous with support
set $\mathcal{X} = (-\infty, \infty)$

$$Y = g(X) = X^2$$

$$F_Y(y) = P(X^2 \leq y) = P(-\sqrt{y} \leq x \leq \sqrt{y})$$

$$= P(X \leq \sqrt{y}) - P(X \leq -\sqrt{y}) = F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

$$\Rightarrow f_Y(y) = \frac{1}{2\sqrt{y}} f_X(\sqrt{y}) + \frac{1}{2\sqrt{y}} f_X(-\sqrt{y})$$

In particular let $f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$, $x \in (-\infty, \infty)$

i.e. $X \sim N(0, 1)$

$$Y = X^2 \Rightarrow f_Y(y) = \frac{1}{2\sqrt{y}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(1/y)^2}{2}} + \frac{1}{2\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-\frac{(-1/y)^2}{2}}$$

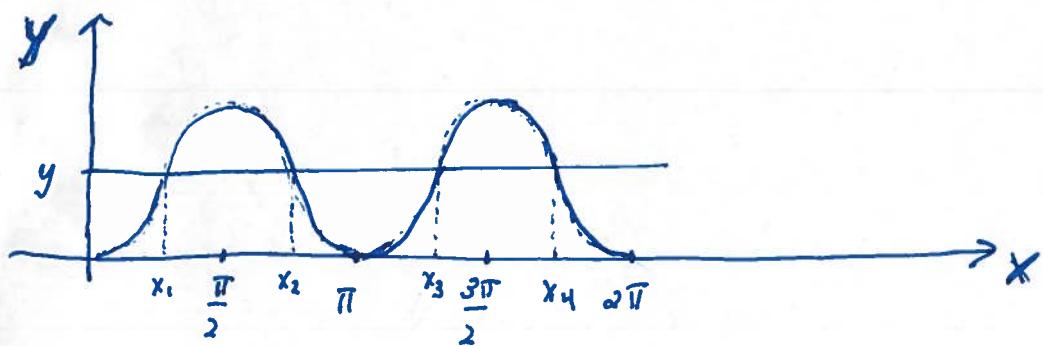
$$= \frac{1}{2\sqrt{y}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{y}{2}}, \quad 0 < y < \infty \quad \text{i.e. } Y \sim \chi^2(1)$$

Example 2.1.2

X cont, uniform on $(0, 2\pi)$

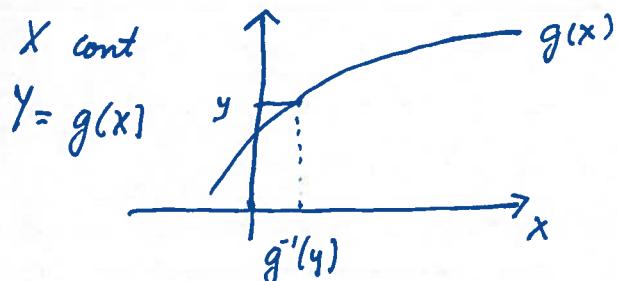
i.e. $f_X(x) = \begin{cases} \frac{1}{2\pi} & 0 < x < 2\pi \\ 0 & \text{otherwise} \end{cases}$, $F_X(x) = \frac{x}{2\pi}, \quad 0 < x < 2\pi$

$$Y = \sin^2 X$$

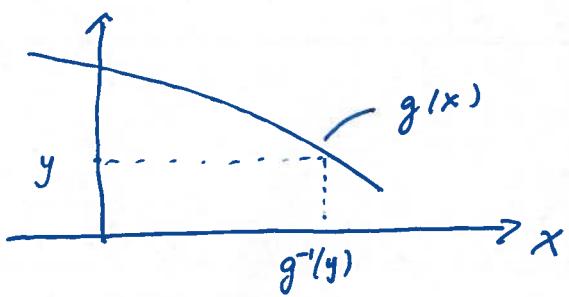


$$P(Y \leq y) = P(X \leq x_1) + P(x_2 \leq X \leq x_3) + P(X = x_4) = F_X(x_1) + (F_X(x_3) - F_X(x_2)) + 1 - F_X(x_4) = \frac{1}{2\pi} \{x_1 + x_3 - x_2 - x_4 + 2\pi\}$$

Monotone transformation. Theorem 2.1.5



$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(g(x) \leq y) \\ &= P(X \leq g^{-1}(y)) = F_X(g^{-1}(y)) \end{aligned}$$



$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(g(x) \leq y) \\ &= 1 - P(X \geq g^{-1}(y)) = 1 - F_X(g^{-1}(y)) \end{aligned}$$

We get $f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|$

A special transformation 2.1.10.

X cont. $F_X(x) = P(X \leq x)$. $F_X(x)$ is monotone

Let $y = F_X(x) = P(Y \leq y) = P(F_X(x) \leq y) = P(X \leq F_X^{-1}(y))$
 $= F_X(F_X^{-1}(y)) = y$ i.e. Y is uniform on $[0, 1]$.

Let Y be uniform on $[0, 1]$. Let $X = F_X^{-1}(Y)$

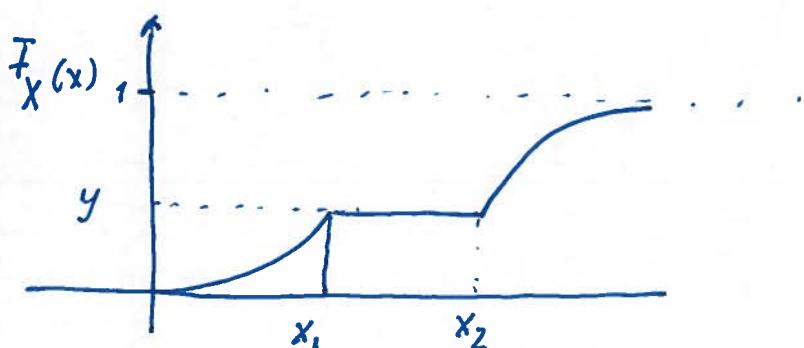
$$\Rightarrow P(X \leq x) = P(F_X^{-1}(Y) \leq x) = P(Y \leq F_X(x)) = F_X(x)$$

Example $X \sim \exp(\lambda) \Rightarrow F_X(x) = 1 - e^{-\lambda x}, x > 0$

Y cont. unif $[0, 1]$. $y = F_X(x) \Rightarrow 1 - e^{-\lambda x} = y$

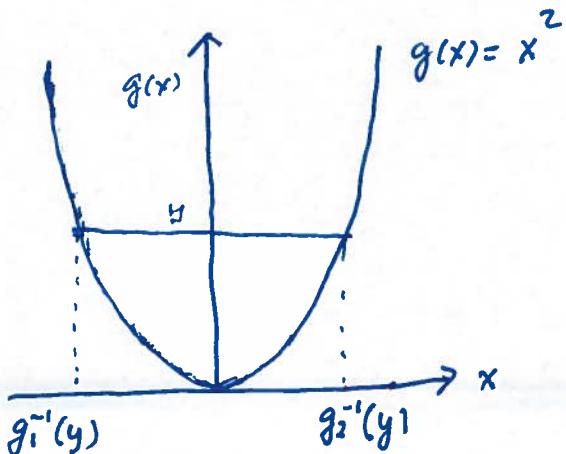
$$\Rightarrow e^{-\lambda x} = 1 - y \Rightarrow x = -\frac{1}{\lambda} \ln(1-y)$$

What if $F_X(x)$ is constant on some interval



$$F_X^{-1}(y) = \inf \{x : F_X(x) \geq y\} \text{ s.t. } x,$$

Example 2.1.7 reconsidered



$$Y = g(x) = X^2$$

Define: $A_0 = \{0\}$

$$A_1 = (-\infty, 0), \quad g_1(x) = x^2, \quad g_1^{-1}(y) = -\sqrt{y}$$

$$A_2 = (0, \infty), \quad g_2(x) = x^2, \quad g_2^{-1}(y) = \sqrt{y}$$

$$Y = \{y : y = g_i(x)\} = (0, \infty) \quad \text{for } i = 1, 2$$

$$\begin{aligned} F_Y(y) &= P(X^2 \leq y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) \\ &= P(X \leq \sqrt{y}) - P(X \leq -\sqrt{y}) = F_X(g_2^{-1}(y)) - F_X(g_1^{-1}(y)) \end{aligned}$$

$$\Rightarrow f_Y(y) = f_X(g_2^{-1}(y)) \frac{d}{dy} g_2^{-1}(y) - f_X(g_1^{-1}(y)) \frac{d}{dy} g_1^{-1}(y)$$

$$= \sum_{i=1}^2 f_X(g_i^{-1}(y)) \left| \frac{d}{dy} g_i^{-1}(y) \right|$$

2.2. - Expectation

$$E[X] = \begin{cases} -\int_{-\infty}^{\infty} x f_X(x) dx & \text{if } \begin{cases} \int_{-\infty}^{\infty} |x| f_X(x) dx < \infty \\ \sum |x| P(X=x) < \infty \end{cases} \\ \sum x P(X=x) \end{cases}$$

Example 2.24

$$f_X(x) = \frac{1}{\pi(1+x^2)}, \quad -\infty < x < \infty$$

$$\begin{aligned} E[X] &= \int_{-\infty}^{\infty} \frac{x}{\pi(1+x^2)} dx = \int_{-\infty}^0 \frac{x}{\pi(1+x^2)} dx + \int_0^{\infty} \frac{x}{\pi(1+x^2)} dx \\ &= \int_{-\infty}^0 \frac{-y}{\pi(1+y^2)} dy + \int_0^{\infty} \frac{x}{\pi(1+x^2)} dx \\ &= - \int_0^{\infty} \frac{y}{\pi(1+y^2)} dy + \left[\frac{\log(1+y^2)}{2\pi} \right]_0^{\infty} \\ &= -\infty + \infty \end{aligned}$$

$$E[|X|] = \int_{-\infty}^{\infty} \frac{|x|}{\pi} \cdot \frac{1}{1+x^2} dx = 2 \int_0^{\infty} \frac{x}{\pi} \frac{1}{1+x^2} dx = \infty \Rightarrow$$

$E[X]$ does not exist.

Important results

$$E[g(x)] = \begin{cases} -\int g(x) f_X(x) dx \\ \sum_{x \in X} g(x) P(X=x) \end{cases} \quad \text{if they exist}$$

$$E\left[\sum_{i=1}^n g(X_i)\right] = \sum_{i=1}^n E[g(X_i)]$$