

$$\Rightarrow f(y_1, \dots, y_m) = \frac{m^{\frac{1}{2}} e^{-\frac{1}{2} m y_1^2}}{(2\pi)^{\frac{1}{2}}} \cdot \frac{m^{\frac{1}{2}} e^{-\frac{1}{2} \left[\sum_{i=2}^m y_i^2 + \left(\sum_{i=2}^m y_i \right)^2 \right]}}{(2\pi)^{\frac{m-1}{2}}}$$

$\Rightarrow Y_1$ and Y_2, \dots, Y_m are independent and thereby

\bar{X} and S .

Proof of c

$$(m-1) S_m^2 = \sum_{i=1}^m (X_i - \bar{X})^2 = \sum_{i=1}^{m-1} (X_i - \bar{X})^2 + (X_m - \bar{X})^2$$

$$= \sum_{i=1}^{m-1} (X_i - \bar{X}_{m-1} + \bar{X}_{m-1} - \bar{X})^2 + (X_m - \bar{X}_{m-1} + \bar{X}_{m-1} - \bar{X})^2$$

$$= \sum_{i=1}^{m-1} (X_i - \bar{X}_{m-1})^2 + (m-1) (\bar{X}_{m-1} - \bar{X})^2 + \frac{(X_m - \bar{X}_{m-1})^2 + (\bar{X}_{m-1} - \bar{X})^2}{m (\bar{X}_{m-1} - \bar{X})^2}$$

$$+ 2 ((X_m - \bar{X}_{m-1})(\bar{X}_{m-1} - \bar{X}))$$

We have

$$\bar{X}_{m-1} = \frac{m\bar{X} - X_m}{m-1} \Rightarrow m\bar{X}_{m-1} - \bar{X}_{m-1} = m\bar{X} - X_m$$

$$\Rightarrow \bar{X}_{m-1} - \bar{X} = \frac{\bar{X}_{m-1} - X_m}{m}$$

$$\text{Therefore: } (m-1) S_m^2 = \sum_{i=1}^{m-1} (X_i - \bar{X}_{m-1})^2 + \underbrace{\frac{(X_m - \bar{X}_{m-1})^2}{m} + \frac{(X_m - \bar{X}_{m-1})^2}{m} - \frac{2(X_m - \bar{X}_{m-1})^2}{m}}_{\frac{m-1}{m} (X_m - \bar{X}_{m-1})^2}$$

$$= \sum_{i=1}^{m-1} (X_i - \bar{X}_{m-1})^2 + \frac{m-1}{m} (X_m - \bar{X}_{m-1})^2$$

$$= (m-2) S_{m-1}^2 + \frac{m-1}{m} (X_m - \bar{X}_{m-1})^2$$

$m=2$

$$S_2^2 = \frac{1}{2} (X_2 - \bar{X}_1)^2 = Y^2, \text{ where } Y = \frac{X_2 - X_1}{\sqrt{2}}$$

$$\Rightarrow \text{Var}(Y) = 1 \Rightarrow Y \sim N(0,1) \Rightarrow Y^2 \sim \chi^2(1).$$

$m=3$

$$2 S_3^2 = S_2^2 + \frac{2}{3} (X_3 - \bar{X}_2)^2 = S_2^2 + \frac{2}{3} \left(X_3 - \frac{X_1 + X_2}{2} \right)^2$$

$$Y = \frac{X_3 - \frac{X_1 + X_2}{2}}{\sqrt{\frac{3}{2}}} \Rightarrow \text{Var}[Y] = \frac{\frac{3}{2}}{\frac{3}{2}} = 1 \Rightarrow Y \sim N(0,1)$$

$$\text{and } Y^2 \sim \chi^2(1) \Rightarrow 2 S_3^2 = S_2^2 + Y^2 \sim \chi^2(1) + \chi^2(1) \sim \chi^2(2)$$

since S_2^2 and $X_3 - \bar{X}_2$ are independent.

5.3.2. Student's t distribution and Snedecor's F

Let $X_1, \dots, X_m \sim N(\mu, \sigma^2)$ and independent, $E[\bar{X}] = \mu$

$$\text{Var}[\bar{X}] = \frac{\sigma^2}{m} \text{ and } \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{m}}} \sim N(0,1)$$

σ^2 unknown: $S^2 = \frac{1}{m-1} \sum_{i=1}^m (X_i - \bar{X})^2$ is an estimator for σ^2

and $S = \sqrt{\frac{1}{m-1} \sum_{i=1}^m (X_i - \bar{X})^2}$ an estimator for σ

What is the distribution of $\frac{\bar{X} - \mu}{\frac{S}{\sqrt{m}}}$?

$$\frac{\bar{X} - \mu}{\frac{S}{\sqrt{m}}} = \frac{\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{m}}}}{\frac{S}{\sigma}} = \frac{\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{m}}} \sim N(0,1)}{\sqrt{\frac{S^2}{\sigma^2}} \sim \sqrt{\frac{\chi^2(m-1)}{m-1}}} \Rightarrow \text{independent.}$$

Let $X \sim N(0,1)$ and $Y \sim \chi^2(m-1) = \text{gamma}(\frac{m-1}{2}, 2)$

Then $f_{X,Y}(x,y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cdot \frac{1}{\Gamma(\frac{m-1}{2}) 2^{\frac{m-1}{2}}} y^{\frac{(m-1)}{2}-1} e^{-\frac{y}{2}}, \quad \begin{matrix} -\infty < x < \infty \\ 0 < y < \infty \end{matrix}$

Define $Z = \frac{X}{\sqrt{\frac{Y}{m-1}}}$ } \Rightarrow $X = \sqrt{\frac{W}{m-1}} Z$
 $W = Y$

$\Rightarrow J = \begin{vmatrix} \sqrt{\frac{W}{m-1}} & \frac{\partial X}{\partial W} \\ 0 & 1 \end{vmatrix} = \sqrt{\frac{W}{m-1}}$

$f_{Z,W}(z,w) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2 \frac{w}{m-1}} \cdot \frac{1}{\Gamma(\frac{m-1}{2}) 2^{\frac{m-1}{2}}} w^{\frac{(m-1)}{2}-1} e^{-\frac{w}{2}} \cdot \sqrt{\frac{w}{m-1}}, \quad \begin{matrix} -\infty < z < \infty \\ 0 < w < \infty \end{matrix}$

$\Rightarrow f_z(z) = \frac{1}{\sqrt{2\pi}} \cdot \frac{1}{\Gamma(\frac{m-1}{2})} \cdot \frac{1}{2^{\frac{m-1}{2}}} \cdot \frac{1}{\sqrt{m-1}} \int_0^\infty e^{-\frac{1}{2}(1+\frac{z^2}{m-1})w} \cdot w^{\frac{m-1}{2}} dw$

$\frac{1}{\sqrt{2\pi}} \cdot \frac{1}{\Gamma(\frac{m-1}{2})} \cdot \frac{1}{2^{\frac{m-1}{2}}} \cdot \frac{1}{\sqrt{m-1}} \int_0^\infty \frac{2^{\frac{m-1}{2}-1} t^{\frac{m-1}{2}-1} \cdot e^{-t}}{(1+\frac{z^2}{m-1})^{\frac{m-1}{2}-1} \cdot (1+\frac{z^2}{m-1})} \cdot \frac{2}{m-1} dt$

$= \frac{1}{\sqrt{\pi(m-1)}} \cdot \frac{\Gamma(\frac{m}{2})}{\Gamma(\frac{m-1}{2})} \cdot \frac{1}{(1+\frac{z^2}{m-1})^{\frac{m}{2}}}$

Definition 5.3.4

If T has a pdf given by $f_T(t) = \frac{1}{\sqrt{\pi p}} \cdot \frac{\Gamma(\frac{p+1}{2})}{\Gamma(\frac{p}{2})} \cdot \frac{1}{(1+\frac{t^2}{p})^{\frac{p+1}{2}}}$

$-\infty < t < \infty$, it is said to be t -distributed with p degrees of freedom.

Some properties.

$$T = \frac{X}{\sqrt{\frac{Y}{p}}}, \quad X \text{ and } Y \text{ are independent, } X \sim N(0, 1), Y \sim \chi^2(p)$$

$$p > 1 \\ E[T] = E[X] \cdot E\left[\frac{1}{\sqrt{\frac{Y}{p}}}\right] = 0$$

$$\text{Var}[T] = E[X^2] \cdot E\left[\frac{p}{Y}\right] = p \cdot E\left[\frac{1}{Y}\right] \quad \text{where } Y \sim T\left(\frac{p}{2}, 2\right)$$

$$X \sim \text{gamma}(d, \beta) \Rightarrow E[X^m] = \frac{\Gamma(d+m)\beta^m}{\Gamma(d)}, \quad m > -d$$

$$E[X^{-1}] = \frac{\Gamma(\frac{p}{2}-1) \cdot 2^{-1}}{\Gamma(\frac{p}{2})} = \frac{1}{2(\frac{p}{2}-1)} = \frac{1}{p-2} \Rightarrow \text{Var}[T] = \frac{p}{p-2}, \quad p > 2.$$

Definition 5.3.6

let $X_1, \dots, X_m \sim N(\mu_X, \sigma_X^2)$ and $Y_1, \dots, Y_m \sim N(\mu_Y, \sigma_Y^2)$

Assume X_1, \dots, X_m and Y_1, \dots, Y_m are independent random samples. Then $F = \frac{\frac{S_X^2}{\sigma_X^2}}{\frac{S_Y^2}{\sigma_Y^2}} \sim$ Snedecor's F with $m-1$ and

$m-1$ degrees of freedom.

$$F = \frac{\frac{X^2(m-1)}{m-1}}{\frac{X^2(m-1)}{m-1}} \Rightarrow \text{independent.}$$

Equivalently if $f_F(x) = \frac{\Gamma(\frac{p+q}{2})}{\Gamma(\frac{p}{2})\Gamma(\frac{q}{2})} \cdot \left(\frac{p}{q}\right)^{\frac{p}{2}} \cdot \frac{x^{\frac{p}{2}-1}}{(1+\frac{p}{q}x)^{\frac{p+q}{2}}}, \quad 0 < x < \infty$

then F is F (Fisher) distributed with p and q degrees of freedom.