

## ② Theoretical Q's about the exponential family of distribution

MODULE 1

- use  $i$  index to specify that each observation might have different parameters and weights, but suppress here for simplicity (since likelihood not focus).

$$f(y) = \exp \left\{ \frac{y\theta - b(\theta)}{\phi} \cdot w + c(y, \phi, w) \right\}$$

### 1) Binomial distribution

#### • Bernoulli process:

- experiment made up of independent trials
- each trial: success or failure  
     $\downarrow A$
- $P(A)$  is the same in each trial

If we perform  $n$  trials the  $Y = \#$  successes is binomial  $(n, p)$ .

$$\begin{aligned} \bullet f(y) &= \binom{n}{y} p^y (1-p)^{n-y} \\ &= \exp \left( \ln \binom{n}{y} + y \ln p + (n-y) \ln(1-p) \right) \\ &= \exp \left( y [\ln p - \ln(1-p)] + n \ln(1-p) + \ln \binom{n}{y} \right) \end{aligned}$$

$$= \exp \left\{ y \cdot \underbrace{\ln\left(\frac{p}{1-p}\right)}_{\theta} + \underbrace{n \ln(1-p)}_{-b(\theta)} + \underbrace{\ln\binom{n}{y}}_{c(y, \phi, w)} \right\}$$

$$* \theta = \ln\left(\frac{p}{1-p}\right)$$

then  
for  
 $b(\theta)$

$$e^{\theta} = \frac{p}{1-p} \Leftrightarrow (1-p)e^{\theta} = p$$

$$e^{\theta} - pe^{\theta} = p$$

$$e^{\theta} = p + pe^{\theta} = p(1+e^{\theta})$$

$$\underline{p = \frac{e^{\theta}}{1+e^{\theta}}}$$

$$\text{so we have } \underline{(1-p)} \Rightarrow 1-p = 1 - \frac{e^{\theta}}{1+e^{\theta}} = \frac{1+e^{\theta} - e^{\theta}}{1+e^{\theta}} = \underline{\frac{1}{1+e^{\theta}}}$$

$$\text{and } \underline{b(\theta)} = -n \ln(1-p) = -n \ln\left(\frac{1}{1+e^{\theta}}\right) =$$

$$-n \underbrace{\ln 1}_0 + n \ln(1+e^{\theta}) = \underline{n \ln(1+e^{\theta})}$$

$$\text{then } \underline{\phi = 1} \text{ and } \underline{w = 1} \text{ and } \underline{c(y, \phi, w) = \ln\binom{n}{y}}$$

Textbook: Bernoulli ( $n=1$ ) has

$$\theta = \ln \frac{p}{1-p} \quad b(\theta) = \ln(1+e^{\theta}) \quad \text{and } \phi = 1$$



## 2) Poisson distribution

a) Poisson process: we observe events occurring within a time interval or area in space.

- the number of events occurring in one interval or area is independent of the number of events in disjoint intervals/areas
- the probability of an event inside an interval or area is proportional to the length of the interval or size of the area.
- the probability that more than one event occur in a small interval or area is negligible.  
 $\Rightarrow$  we have a Poisson process.

Then:  $Y =$  "the number of events within an interval or area" follows a Poisson distribution.

$$f(y) = \frac{\mu^y}{y!} e^{-\mu}, \text{ where } E(Y) = \text{Var}(Y) = \mu$$

$$f(y) = \frac{\mu^y}{y!} e^{-\mu} = \exp(-\mu + y \ln \mu - \ln y!)$$

$$= \exp\left( \underbrace{y \cdot \ln \mu}_{\theta} - \underbrace{\mu}_{b(\theta)} + \underbrace{(-\ln y!)}_{c(y)} \right)$$

so exp. form with

$$\theta = \ln \mu \Leftrightarrow \mu = \exp \theta$$

$$b(\theta) = \exp \theta$$

$$c(y) = -\ln y!$$

$$E(Y) = \frac{db}{d\theta} = \frac{d}{d\theta} \exp \theta = \exp \theta = \mu$$

$$Var(Y) = \frac{d^2 b}{d\theta^2} = \frac{d}{d\theta} \exp \theta = \exp \theta = \mu$$

### 3) Normal distribution

Measurements of physical properties or scientific measurement (with error) is known to be normally distributed. This might be thought of as a version of the central limit theorem.

$$f(y) = \frac{1}{\sqrt{2\pi}\sigma} \cdot \exp\left\{-\frac{1}{2} \frac{(y-\mu)^2}{\sigma^2}\right\}$$

$$\text{and } E(Y) = \mu, \text{Var}(Y) = \sigma^2$$

$$f(y) = \exp\left\{\left(-\frac{1}{2}\sigma^{-2}\right)(y^2 - 2\mu y + \mu^2) - \ln(\sqrt{2\pi}\sigma)\right\}$$

$$= \exp\left\{\left(-\frac{1}{2}\sigma^{-2}\right)(-2\mu \cdot y + \mu^2) + \left(-\frac{1}{2}\sigma^{-2}\right)y^2 - \ln(\sqrt{2\pi}\sigma)\right\}$$

$$= \exp\left\{\underbrace{\frac{1}{\sigma^2}}_{\frac{1}{\phi}} \cdot \underbrace{\mu \cdot y}_{\theta \cdot y} - \underbrace{\frac{1}{\sigma^2}}_{\frac{1}{\phi}} \cdot \underbrace{\frac{1}{2}\mu^2}_{b(\theta)} + \underbrace{\left(-\frac{1}{2}\frac{1}{\sigma^2}y^2 - \ln(\sqrt{2\pi}\sigma)\right)}_{c(y, \phi)}\right\}$$

$$\theta = \mu, \phi = \sigma^2, w = 1, c(y, \phi)$$

$$\Rightarrow \text{exponential family} \quad b(\theta) = \frac{1}{2}\theta^2$$

$$E(Y) = \frac{d}{d\theta} b(\theta) = \frac{d}{d\theta} \left( \frac{1}{2}\theta^2 \right) = \theta = \underline{\underline{\mu}}$$

$$\frac{d^2 b(\theta)}{d\theta^2} = \frac{d}{d\theta} \theta = 1$$

$$Var(Y) = \frac{d^2 b(\theta)}{d\theta^2} \cdot \frac{\phi}{\omega} = 1 \cdot \frac{\sigma^2}{1} = \underline{\underline{\sigma^2}}$$

#### 4) Gamma distribution

The waiting time until the  $v$ th event in a Poisson process has a Gamma (or Erlang) distribution. When  $v=1$  the exponential distribution occurs - and is much used in survival analysis. The  $\chi^2_\delta$  distribution is a special case of the gamma distribution ( $\frac{v}{\mu}=2$  and  $v=\frac{\delta}{2}$  with  $\delta$  as param. in the  $\chi^2$  - but usually  $v$ , so confusing).

Use parameterization on page 643 of appendix B

$$f(y) = \frac{1}{\Gamma(v)} \left(\frac{v}{\mu}\right)^v y^{v-1} \exp\left(-\frac{v}{\mu}y\right), y > 0$$

$$f(y) = \exp\left\{-\frac{v}{\mu} \cdot y + v \ln\left(\frac{v}{\mu}\right) + (v-1) \cdot \ln y - \ln(\Gamma(v))\right\}$$

$$= \exp\left\{-\frac{\frac{1}{\mu} \cdot y}{\frac{1}{v}} + \frac{\ln v + \ln\left(\frac{1}{\mu}\right)}{\frac{1}{v}} + (v-1) \ln y - \ln(\Gamma(v))\right\}$$

$$= \exp\left\{\underbrace{-\frac{1}{\mu} \cdot y + \ln\left(-\frac{1}{\mu}\right)}_{\substack{\theta \\ \phi}} + \underbrace{v \cdot \ln v + (v-1) \ln y - \ln(\Gamma(v))}_{c(y, \phi)}\right\}$$

$$b(\theta) = -\ln\left(-\frac{1}{\mu}\right) = -\ln(-\theta)$$

$$\theta = -\frac{1}{\mu} \Leftrightarrow \mu = -\frac{1}{\theta}$$

$w=1$

8

$$E(Y) = b'(\theta) = \frac{d}{d\theta} (-\ln(-\theta)) = -\frac{1}{-\theta} (-1) = \underline{\underline{-\frac{1}{\theta} = \mu}}$$

$$\frac{d^2 b(\theta)}{d\theta^2} = \frac{d}{d\theta} \left(-\frac{1}{\theta}\right) = \frac{1}{\theta^2} = \mu^2$$

$$\text{Var}(Y) = \frac{d^2 b(\theta)}{d\theta^2} \cdot \frac{\psi}{w} = \frac{1}{\theta^2} \cdot \frac{1}{1} = \underline{\underline{\frac{1}{\theta^2} = \frac{\mu^2}{\nu}}}$$