

# M4: Poisson and gamma regression

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Poisson process  $\rightarrow$  number of events in time and/or space follows  
a Poisson distribution  $Y_i$   $i=1, \dots, n$

$$Y_i \sim \text{Poisson}(\lambda_i)$$

$$f(y_i) = \frac{\lambda_i^{y_i}}{y_i!} e^{-\lambda_i} \quad y_i = 0, 1, 2, \dots$$

$$E(Y_i) = \lambda_i, \text{Var}(Y_i) = \lambda_i \leftarrow \text{NB: } E(Y_i) = \text{Var}(Y_i)$$

Univ. exponential family:  $f(y_i; \theta_i, \phi) = \exp\left(\frac{y_i \cdot \theta_i - b(\theta_i)}{\phi} w_i + c(y_i, \phi, w_i)\right)$

$$\theta_i = \ln \lambda_i, \phi = 1, w_i = 1$$

canonical link:  $\eta_i = \theta_i$

Poisson regression  $\theta_i = \eta_i = \ln(\lambda_i)$

1) Random component:  $Y_i \sim \text{Poisson}(\lambda_i)$   
with  $\mu_i = E(Y_i) = \lambda_i$  and  $\text{Var}(Y_i) = \lambda_i$

2) Systematic component:  $\eta_i = x_i^T \beta$

3) Link:

$$\eta_i = \ln(\lambda_i)$$

$$\lambda_i = \exp(\eta_i)$$

log link

exponential response function

canonical link

## Interpreting regression coefficients

$$x_i^T \beta = \beta_0 + \beta_1 x_{i1} + \dots$$

$$\lambda_i = \exp(\eta_i) = \exp(\beta_0) \cdot \exp(\beta_1 x_{i1}) \dots \exp(\beta_k x_{ik})$$

What happens if  $x_{i1}$  increases to  $x_{i1} + 1$ :

$$\lambda_i = \exp(\beta_0) \cdot \exp(\beta_1) \cdot \exp(\beta_1 x_{i1}) \dots$$

} multiplicative effect of covariate

$\downarrow$   
 $E(Y_i)$  will change with a factor  $\exp(\beta_1)$

Ex: crebs & skeletons:

1)  $\eta_i = \beta_0 : \lambda_i = \exp(\beta_0)$

and  $\hat{\lambda}_i = \exp(\hat{\beta}_0) = \text{average number of } S_0 \text{ in data}$   
 $= 2.91$

2)  $\eta_i = \beta_0 + \beta_1 x_{i1}$        $\hat{\lambda}_i$  "increase by  $e^{\hat{\beta}_1} = 1.17$

$\uparrow$   
width

$\uparrow$   
multiply by

if  $W$  increase by one unit

3)  $\bar{x}_1 = \text{average width} = 26.3$

$$\hat{\lambda} = \exp(\hat{\beta}_0 + \hat{\beta}_1 \cdot \bar{x}_1) \approx 2.74$$

## Maximum likelihood estimation

$$L(\beta) = \prod_{i=1}^n \frac{\lambda_i^{y_i}}{y_i!} e^{-\lambda_i} \quad \begin{cases} \eta_i = \ln \lambda_i \\ \lambda_i = \exp(\eta_i) \end{cases}$$

$$\begin{aligned} l(\beta) &= \sum_{i=1}^n (y_i \ln \lambda_i - \lambda_i - \ln(y_i!)) \\ &= \sum_{i=1}^n [y_i \eta_i - e^{\eta_i} - \ln(y_i!)] = \sum_{i=1}^n l_i(\beta) \end{aligned}$$

$$s(\beta) = \sum_{i=1}^n s_i(\beta) = \sum_{i=1}^n \underbrace{\frac{\partial l_i}{\partial \eta_i}}_{1 \times 1} \underbrace{\frac{\partial \eta_i}{\partial \beta}}_{p \times 1} \quad \frac{\partial (x_i^T \beta)}{\partial \beta} = \begin{matrix} x_i \\ \uparrow \\ p \times 1 \end{matrix}$$

$$\frac{\partial l_i}{\partial \eta_i} = \frac{\partial}{\partial \eta_i} (y_i \eta_i - e^{\eta_i} - \ln(y_i!)) = y_i - e^{\eta_i}$$

$$s(\beta) = \sum_{i=1}^n (y_i - e^{\eta_i}) x_i = \sum_{i=1}^n \overbrace{(y_i - \lambda_i)}^{s_i(\beta)} x_i$$

$$\text{Observe } E(s(\beta)) = \sum_{i=1}^n (E(Y_i) - \lambda_i) \cdot x_i = 0$$

$$F(\beta) = \text{Cov}(s(\beta)) = \sum_{i=1}^n \text{Cov}(s_i(\beta)) = \sum_{i=1}^n F_i(\beta)$$

$$F_i(\beta) = \text{Cov}([Y_i - \lambda_i] x_i) \quad \text{Cov}(z) = E((z - \mu_z)(z - \mu_z)^T)$$

$$= E(((Y_i - \lambda_i) x_i - 0)((Y_i - \lambda_i) x_i - 0)^T)$$

$$= E((Y_i - \lambda_i)^2) x_i x_i^T = \text{Var}(Y_i) \cdot x_i x_i^T = \lambda_i x_i x_i^T$$

$$\underline{\underline{F(\beta) = \sum_{i=1}^n \lambda_i x_i x_i^T}} \quad (p \times p)$$

Hypothesis testing  $\left\{ \begin{array}{l} \text{Wald } C\beta = d \\ \text{LRT} \end{array} \right\}$  "same" as for binary regression

The score test

A: large model :  $p_A$  param.  $\rightarrow$  ML  $\hat{\beta}$     Ex:  $S \sim W + C$      $p_A = 5$   
 B: small model :  $p_B$  param.  $\rightarrow$  ML  $\tilde{\beta}$      $S \sim W$      $p_B = 2$

$H_0$ : large and small model is equally good  
 "the  $\beta$ 's in  $\beta_A$  but not in  $\beta_B$  are all zero"

$H_1$ : not so

$$U = \left( s(\tilde{\beta}) - \overset{0}{s(\hat{\beta})} \right)^T \underbrace{F^{-1}(\tilde{\beta})}_{\substack{[(p_A - p_B) \times 1]^T \\ (p_A - p_B) \times (p_A - p_B)}} \underbrace{\left( s(\tilde{\beta}) - \overset{0}{s(\hat{\beta})} \right)}_{(p_A - p_B) \times 1}$$

↑ test statistics

If  $U$  is large  $\Rightarrow$  " $H_0$  is not correct"  $\rightarrow$  reject  $H_0$   
 $\Rightarrow$  the B model is not as good as the A model

When  $n$  is large  $U \sim \chi^2_{p_A - p_B}$

See R-code on module page for details on calculating  $U$ .